MODELLING FOR ENGINEERING & HUMAN BEHAVIOUR 2019



Instituto Universitario de Matemática Multidisciplinar Polytechnic City of Innovation

Edited by R. Company, J.C. Cortés, L. Jódar and E. López-Navarro







July 10th - 12th 2019

CIUDAD POLITÉCNICA DE LA INNOVACIÓN

Modelling for Engineering & Human Behaviour 2019

València, 10 - 12 July 2019

This book includes the extended abstracts of papers presented at XXIst Edition of the Mathematical Modelling Conference Series at the Institute for Multidisciplinary Mathematics "Mathematical Modelling in Engineering & Human Behaviour". I.S.B.N.: 978-84-09-16428-8 Version: 18/11/19 Report any problems with this document to ellona1@upvnet.upv.es.

Edited by: R. Company, J. C. Cortés, L. Jódar and E. López-Navarro. Credits: The cover has been designed using images from kjpargeter/freepik.



Instituto Universitario de Matemática Multidisciplinar This book has been supported by the European Union through the Operational Program of the [European Regional Development Fund (ERDF) / European Social Fund (ESF)] of the Valencian Community 2014-2020. [Record: GJIDI/2018/A/010].





Fons Europeu de Desenvolupament Regional

Una manera de fer Europa

Contents

A personality mathematical model of placebo with or without deception: an application of the Self-Regulation Therapy
The role of police deterrence in urban burglary prevention: a new mathematical approach
A Heuristic optimization approach to solve berth allocation problem14
Improving the efficiency of orbit determination processes
A new three-steps iterative method for solving nonlinear systems
Adaptive modal methods to integrate the neutron diffusion equation26
Numerical integral transform methods for random hyperbolic models
Nonstandard finite difference schemes for coupled delay differential models37
Semilocal convergence for new Chebyshev-type iterative methods
Mathematical modeling of Myocardial Infarction 46
Symmetry relations between dynamical planes
Econometric methodology applied to financial systems
New matrix series expansions for the matrix cosine approximation
Modeling the political corruption in Spain70
Exponential time differencing schemes for pricing American option under the Heston model
Chromium layer thickness forecast in hard chromium plating process using gradient boosted regression trees: a case study
Design and convergence of new iterative methods with memory for solving nonlinear problems
Study of the influence falling friction on the wheel/rail contact in railway dynamics88
Extension of the modal superposition method for general damping applied in railway dynamics
Predicting healthcare cost of diabetes using machine learning models

Sampling of pairwise comparisons in decision-making105
A multi-objective and multi-criteria approach for district metered area design: water operation and quality analysis 110
Updating the OSPF routing protocol for communication networks by optimal decision- making over the k-shortest path algorithm
Optimal placement of quality sensors in water distribution systems
Mapping musical notes to socio-political events
Comparison between DKGA optimization algorithm and Grammar Swarm surrogated model applied to CEC2005 optimization benchmark
The quantum brain model
Probabilistic solution of a randomized first order differential equation with discrete delay
A predictive method for bridge health monitoring under operational conditions 155
Comparison of a new maximum power point tracking based on neural network with conventional methodologies
Influence of different pathologies on the dynamic behaviour and against fatigue of railway steel bridges
Statistical-vibratory analysis of wind turbine multipliers under different working con- ditions
Analysis of finite dimensional linear control systems subject to uncertainties via prob- abilistic densities
Topographic representation of cancer data using Boolean Networks
Trying to stabilize the population and mean temperature of the World 185
Optimizing the demographic rates to control the dependency ratio in Spain193
An integer linear programming approach to check the embodied CO_2 emissions of the opaque part of a façade
Acoustics on the Poincaré Disk
Network computational model to estimate the effectiveness of the influenza vaccine <i>a</i> posteriori
The key role of Liouville-Gibbs equation for solving random differential equations: Some insights and applications

Numerical integral transform methods for random hyperbolic models

M.-C. Casabán^{b1}, R. Company^b and L. Jódar^b

(b) Instituto Universitario de Matemática Multidisciplinar, Universitat Politècnica de València.

1 Introduction

This work deals with the construction of analytic-numerical solutions, in the mean square sense [1], of the random heterogeneous telegraph type problem

$$u_{tt}(x,t) = (k(x) u_x(x,t))_x + a(x) u(x,t) + \psi(x,t), \quad x > 0, \ t > 0,$$
(1)

$$u(0,t) = g_1(t),$$

$$u(0,t) = g_2(t)$$
(2)
(3)

$$u_x(0,t) = g_2(t), (3)$$

$$u(x,0) = g(x), \qquad (4)$$

where a(x), k(x), $\psi(x,t)$, $g_1(t)$, $g_2(t)$ and g(x) are stochastic processes (s.p.'s) with a finite degree of randomness [1].

Efficient methods for solving numerically deterministic problems such as finite-difference methods become unsuitable for the random case because of the computation of the expectation and the variance of the approximation solution s.p. The drawbacks are essentially of computational complexity such as the handling of big random matrices which appear throughout the iterative levels of the discretization steps and the necessity to store the information of all the previous levels of the iteration process. Then, they motive the search of non iterative alternatives. In this sense, this paper provides an approximation solution s.p. of the problem (1)–(4) which combines the random Fourier sine transform, the Gauss-Laguerre quadrature rule and the Monte Carlo method.

2 Gauss-Laguerre solution of a random hyperbolic model

The construction of an approximated solution s.p. of the problem (1)-(4) will be in two-stages. Firstly, using the Fourier sine transform, an infinite integral form solution of the theoretical

¹email: macabar@imm.upv.es

solution is obtained. Then, using random Gauss-Laguerre quadrature formulae a random numerical solution is represented that is further computer by means of Monte Carlo simulations at appropriated root points of the Laguerre polynomials.

Let $V(x)(\xi) = \mathfrak{F}_s[u(x,\cdot)](\xi)$ be the Fourier sine transform of the unknown $u(x,\cdot)$:

$$V(x)(\xi) = \mathfrak{F}_s[u(x,\cdot)](\xi) = \int_0^{+\infty} u(x,t) \,\sin(\xi \, t) \,\mathrm{dt} \,, \quad \xi > 0 \,, \ x > 0 \,. \tag{5}$$

Let us denote

$$G_1(\xi) = \mathfrak{F}_s[u(0,\cdot)](\xi) = \mathfrak{F}_s[g_1(t)](\xi), \quad \xi > 0,$$
(6)

$$G_2(\xi) = \mathfrak{F}_s[u_x(0,\cdot)](\xi) = \mathfrak{F}_s[g_2(t)](\xi), \quad \xi > 0,$$
(7)

$$\Psi(x)(\xi) = \mathfrak{F}_s[\psi(x,\cdot)](\xi), \quad x > 0, \ \xi > 0.$$
(8)

Let us assume that the s.p.'s k(x), a(x), $\psi(x,t)$, $g_1(t)$, $g_2(t)$ and g(x) of problem (1)–(4) are mean four (m.f.) continuous with a finite degree of randomness. Let k(x) be a positive s.p. 4-differentiable and let $\psi(x,t)$, $g_1(t)$, $g_2(t)$ be m.f. absolutely integrable s.p.'s in t > 0. By applying random Fourier sine transform to problem (1)–(4) and using the properties of the random Fourier sine transform, [2], one gets, for $\xi > 0$ fixed

$$\frac{d^2}{dx^2}(V(x))(\xi) + \frac{k'(x)}{k(x)}\frac{d}{dx}(V(x))(\xi) + \frac{a(x) + \xi^2}{k(x)}V(x)(\xi) = \frac{\Psi(x)(\xi) - \xi g(x)}{k(x)}, \qquad (9)$$

together with

$$V(0)(\xi) = G_1(\xi) ,$$

$$\frac{d}{dx}(V(0))(\xi) = G_2(\xi) .$$
(10)

Solution of problem (9)–(10) is the first component of the solution of extended random linear differential system, $V(x)(\xi) = [1, 0] X(x)(\xi)$,

$$X'(x)(\xi) = L(x)(\xi) X(x)(\xi) + B(x)(\xi), \quad x > 0, X(0)(\xi) = Y_0(\xi),$$
(11)

where

$$L(x)(\xi) = \begin{bmatrix} 0 & 1\\ -\frac{\xi^2 + a(x)}{k(x)} & -\frac{k'(x)}{k(x)} \end{bmatrix}, B(x)(\xi) = \begin{bmatrix} 0\\ \frac{\Psi(x)(\xi) - \xi g(x)}{k(x)} \end{bmatrix}, \\ Y_0(\xi) = \begin{bmatrix} G_1(\xi)\\ G_2(\xi) \end{bmatrix}.$$
(12)

Assuming that 4-s.p.'s a(x), k(x) and k'(x) satisfy the moment condition

$$\mathbb{E}\left[\left|s(x)\right|^{r}\right] \le m h^{r} < +\infty, \quad \forall r \ge 0,$$
(13)

for every x > 0, it is guaranteed that the entries of the matrix s.p. $L(x)(\xi) \in L_4^{2\times 2}(\Omega)$, for $\xi > 0$ fixed, satisfy condition (13). Condition (13) guarantees that $L(x)(\xi)$ is 4-locally absolutely integrable. Furthermore, it is verifies that vector s.p.'s both $B(x)(\xi)$ and $Y_0(x)(\xi)$ lie in $L_4^{2\times 1}(\Omega)$ and they are absolutely integrables in $x \in [0, +\infty)$. By using random inverse Fourier sine transform to $V(x)(\xi)$ one gets

$$u(x,t) = \frac{2}{\pi} \int_0^\infty V(x)(\xi) \,\sin(\xi t) \,\mathrm{d}\xi = \frac{2}{\pi} \int_0^\infty X_1(x)(\xi) \,\sin(\xi t) \,\mathrm{d}\xi \,, \tag{14}$$

where $X_1(x)(\xi) = [1, 0] X(x)(\xi)$. Now, taking advantage of the Gauss-Laguerre quadrature formula of degree N, see page 890 of [3], for a s.p. $\mathcal{J}(\xi) \in L_2(\Omega)$ being m.f.-absolutely integrable respect to $\xi > 0$, we can consider the following numerical approximation for each event $\omega \in \Omega$

$$I_N^{\text{G-L}}[\mathcal{J}](\omega) = \sum_{j=1}^N \nu_j \ \mathcal{J}(\vartheta_j;\omega) , \qquad \nu_j = \frac{\vartheta_j}{\left[(N+1) L_{N+1}(\vartheta_j) \right]^2} , \tag{15}$$

where ϑ_j is the *j*-th root of the deterministic Laguerre polynomial, $L_N(\vartheta)$, of degree N and ν_j is the weight. This quadrature formula is going to be applied to the r.v. u(x,t) given by (14) taking

$$\mathcal{J}(\xi) = \mathcal{J}(x, t, \xi) = \frac{2}{\pi} X_1(x)(\xi) \sin(\xi t) e^{\xi}.$$

Given the degree N, let us denote by $u_N^{\text{G-L}}(x,t)$ the Gauss-Laguerre s.p. approximation of degree N of the exact solution s.p. u(x,t) of the random problem (1)–(4), evaluated at (x,t) and expressed as the r.v.

$$u_N^{\text{G-L}}(x,t) = \frac{2}{\pi} \sum_{j=1}^N \nu_j \,\sin(\vartheta_j \, t) \, e^{\vartheta_j} \, X_1(x)(\vartheta_j) \,. \tag{16}$$

The exact solution $X_1(x)(\vartheta_j)$ is going to be obtained using Monte Carlo simulation because it is not available. We denoted by $\mathbb{E}_{MC}^K[\bar{X}_1(x)(\vartheta_j)]$ and $\operatorname{Cov}_{MC}^K\left[\bar{X}_1(x)(\vartheta_j), \bar{X}_1(x)(\vartheta_\ell)\right]$ the expectation and the covariance, respectively, of K number of realizations used in the Monte Carlo (MC) simulation and $\bar{X}_1(x)(\vartheta_j)$ the deterministic numerical solution obtained after taking K realizations. Thus the final expressions for the approximations of the expectation and the variance of the solution s.p. take the form

$$\mathbb{E}[u_N^{\text{G-L}}(x,t)] \approx \mathbb{E}[u_{N,K}^{\text{G-L}}(x,t)] = \frac{2}{\pi} \sum_{j=1}^N \nu_j \,\sin(\vartheta_j \, t) \, e^{\vartheta_j} \, \mathbb{E}_{MC}^K[\bar{X}_1(x)(\vartheta_j)] \,, \tag{17}$$

$$\operatorname{Var}\left[u_{N}^{\text{G-L}}(x,t)\right] \approx \operatorname{Var}\left[u_{N,K}^{\text{G-L}}(x,t)\right] = \left(\frac{2}{\pi}\right)^{2} \sum_{j=1}^{N} \sum_{\ell=1}^{N} \nu_{j} \nu_{\ell} \sin(\vartheta_{j} t) \sin(\vartheta_{\ell} t) e^{\vartheta_{j} + \vartheta_{\ell}} \operatorname{Cov}_{MC}^{K}\left[\bar{X}_{1}(x)(\vartheta_{j}), \bar{X}_{1}(x)(\vartheta_{\ell})\right].$$
(18)

3 Numerical example

Consider the random heterogeneous telegraph type problem (1)-(4) with the following input data having a finite degree of randomness

$$k(x) = 1 + b \cos(\pi x), \ a(x) = e^{-ax}, \ \psi(x,t) = e^{-(x+t)} \\ g_1(t) = 0, \ g_2(t) = 0, \ g(x) = 0$$
(19)

where parameters a and b are assumed to be both independent r.v.'s, specifically, a has a uniform distribution giving values in [0, 1], that is, $a \sim Un(0, 1)$, and b > 0 has an exponential distribution of parameter 2 truncated on the interval [0.1, 0.2], that is, $b \sim Exp_{[0.1,0.2]}(2)$. Then it is verified that s.p.'s k(x), a(x) and the deterministic functions $\psi(x,t)$, $g_1(t)$, $g_2(t)$ and g(x)are 4-continuous and 4-absolutely integrable with respect to the time variable those depending on t. Furthermore, k(x) is positive and 4-differentiable.

To study the numerical convergence of the approximations of both the expectation and the standard deviation we have studied the behaviour of their root mean square deviations (RMSD) and the absolute deviations (AbsDev), that is

$$\begin{split} \text{RMSD} \left[\mathbb{E}[u_{N,K_{\ell}K_{\ell+1}}^{\text{G-L}}(x_{i},t)] \right] &= \sqrt{\frac{1}{(n+1)} \sum_{\ell=0}^{n} \left(\mathbb{E}[u_{N,K_{\ell+1}}^{\text{G-L}}(x_{i},t)] - \mathbb{E}[u_{N,K_{\ell}}^{\text{G-L}}(x_{i},t)] \right)^{2}}, \\ \text{RMSD} \left[\sqrt{\text{Var}[u_{N,K_{\ell}K_{\ell+1}}^{\text{G-L}}(x_{i},t)]} \right] &= \sqrt{\frac{1}{(n+1)} \sum_{\ell=0}^{n} \left(\sqrt{\text{Var}[u_{N,K_{\ell+1}}^{\text{G-L}}(x_{i},t)} - \sqrt{\text{Var}[u_{N,K_{\ell}}^{\text{G-L}}(x_{i},t)]} \right)^{2}}, \\ \text{AbsDev} \left(\mathbb{E}[u_{N_{\ell}N_{\ell+1},K}^{\text{G-L}}(x,t)] \right) &= \left| \mathbb{E}[u_{N_{\ell+1},K}^{\text{G-L}}(x,t)] - \mathbb{E}[u_{N_{\ell}K_{\ell}}^{\text{G-L}}(x,t)] \right|, \\ \text{AbsDev} \left(\sqrt{\text{Var}[u_{N_{\ell}N_{\ell+1},K}^{\text{G-L}}(x,t)]} \right) &= \left| \sqrt{\text{Var}[u_{N_{\ell+1},K}^{\text{G-L}}(x,t)]} - \sqrt{\text{Var}[u_{N_{\ell}K_{\ell}}^{\text{G-L}}(x,t)]} \right|, \end{split}$$

in two stages. Firstly, varying the number K of realizations in the Monte Carlo method but considering fixed N in the Gauss-Laguerre quadrature rule and secondly, varying N but considering the number of realizations K fixed. Table 1 and Figure 1 illustrated the numerical convergence of our approximations.

Figures

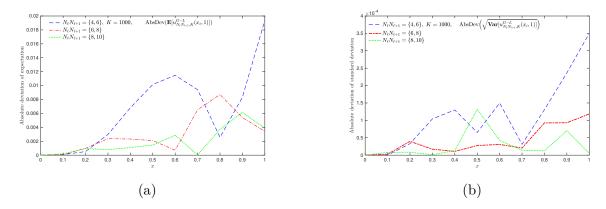


Figure 1: (a): Comparative graphics of the absolute deviations for successive approximations to the expectation $\mathbb{E}[u_{N_{\ell}N_{\ell+1},K}^{\text{G-L}}(x_i, 1)]$. (b): Comparative graphics of the absolute deviations for successive approximations to the standard deviation $\sqrt{\text{Var}[u_{N_{\ell}N_{\ell+1},K}^{\text{G-L}}(x_i, 1)]}$. Both graphics correspond to the time t = 1 on the spatial interval $0 \le x \le 1$, K = 1000 realizations and the degrees $N = \{4, 6, 8, 10\}$ for the Laguerre polynomials.

Tables

$\underline{K_{\ell}K_{\ell+1}}$	$\text{RMSD}\left[\mathbb{E}[u_{6,K_{\ell}K_{\ell+1}}^{\text{G-L}}(x_i,1)]\right]$	$\text{RMSD}\left[\sqrt{\text{Var}[u_{6,K_{\ell}K_{\ell+1}}^{\text{G-L}}(x_i,1)]}\right]$
$K_0 K_1$	2.81922e - 05	2.91484e - 05
K_1K_2	1.96565e - 05	1.18180e - 05
K_2K_3	1.11618e - 05	2.49163e - 06
K_3K_4	1.07918e - 05	4.96128e - 06
K_4K_5	3.59937e - 06	2.83452e - 06

Table 1: Values of the RMSDs for the approximations of the expectation, RMSD $\left[\mathbb{E}_{N,K_{\ell}K_{\ell+1}}^{\text{G-L}}(x_i,t)\right]$, and the standard deviation, RMSD $\left[\sqrt{\operatorname{Var}_{N,K_{\ell}K_{\ell+1}}^{\text{G-L}}(x_i,t)}\right]$, at t = 1 on the spatial domain $0 \le x \le 1$, N = 6 the degree of the Laguerre polynomial and the realizations $K_0 = 2500$, $K_1 = 5000$, $K_2 = 10^4$, $K_3 = 2 \times 10^4$, $K_4 = 4 \times 10^4$ and $K_5 = 5 \times 10^4$.

Acknowledgements

This work has been partially supported by the Spanish Ministerio de Economía y Competitividad grant MTM2017-89664-P.

References

- Soong, T. T., Random Differential Equations in Science and Engineering, New York, Academic Press, 1973.
- [2] Casabán, M.C., Cortés, J.C. and Jódar, L., Solving linear and quadratic random matrix differential equations: A mean square approach. *Appl. Math. Model.*, 40: 9362–9377, 2016.
- [3] Abramowitz, M. and Stegun, I.A., Handbook of Mathematical Functions with Formulas, Graphs, and Mathematical Tables, New York, Dover Publications, Inc., 1972.