

META-ANALYSIS OF GENETIC PARAMETERS FOR ECONOMIC TRAITS IN RABBIT USING A RANDOM-EFFECTS MODEL

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Abstract: The genetic improvement of rabbits helps increase their productivity and, consequently, increase the supply of animal protein for human consumption. The aim of this study was to perform a meta-analysis of genetic parameters (heritability and genetic correlation) for litter size at birth, litter weight at birth, litter size at weaning, litter weight at weaning and slaughter weight in rabbits. The final dataset contained 147 estimates of heritability and 32 estimates of genetic correlation across 34 articles published between 1992 and 2022. A random-effects model was used and the heterogeneity of estimates was assessed using *Q* and *I* 2 statistics. Heritability estimates were of low magnitude for all traits, ranging from 0.09 to 0.18. The lowest heritability estimate was observed for litter size at weaning and the highest for slaughter weight. Most genetic correlations between traits were positive and moderate, ranging from 0.44 to 0.60. Significant heterogeneity among studies justified the use of random-effects models. The meta-analysis study provided reliable genetic parameter estimates and these results can support the development of rabbit breeding programmes.

Key Words: genetic correlation, heritability, heterogeneity, genetic selection, rabbit.

INTRODUCTION

Rabbit (*Oryctolagus cuniculus*) production has evolved over the years due to its potential for meat, fur, laboratory animal, pharmaceutical industry, biological production and other purposes of commercial interest (Dige *et al.*[, 2012\)](#page-6-0). In 2019, global rabbit meat production reached 883 936 tonnes, becoming a source of quality animal protein in developing countries (Kumar *et al.*[, 2023\).](#page-7-0) These animals have high fecundity and prolificacy, being an economic alternative in urban and rural areas ([Montes-Vergara](#page-7-1) *et al.*, 2021). Genetic improvement of rabbits contributes to increased herd productivity and supply of animal protein for human consumption. When selecting these animals, it is essential to know the genetic parameters of traits of economic importance in the production system. Traits such as litter size, litter weight and body weight are some of the aspects that have been evaluated in rabbit breeding programmes ([Sakthivel](#page-7-2) *et al.*, 2017; Farouk *et al.*[, 2022\)](#page-7-3). Among these traits, litter size at birth has the greatest economic value ([Nguyen](#page-7-4) *et al.*, 2017) in prolific species. The number of young rabbits achieved in commercial lines depends on the number kits born alive and their postnatal survival. Rabbits with larger birth weight have higher body weight at their first mating, as well as during their reproductive life (Szendrő *et al.*, 2019). The decision on selection criteria should consider the knowledge of the genetic architecture that these traits present in the population and how they are associated.

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In this sense, meta-analysis employs statistical methods to combine and summarise the results from multiple independent scientific studies (Sutton *et al.*[, 2000\)](#page-7-5). In the context of rabbit breeding, meta-analysis affords synthesised knowledge of the genetic parameters for economically important traits. In animal breeding programmes, the application of meta-analysis study helps obtain useful estimates of the parameters that could support genetic evaluations when reliable estimates for traits in rabbit are not available [\(Oliveira](#page-7-6) *et al.*, 2017). Previous studies have provided genetic parameter estimates for several crucial economic traits in rabbits, including litter size at birth, litter weight at birth, litter size at weaning and litter weight at weaning [\(Sakthivel](#page-7-2) *et al.*, 2017; [Ezzeroug](#page-7-7) *et al.*, 2019; [Farouk](#page-7-3) *et al.*[, 2022\).](#page-7-3) However, these estimates are derived from populations with different sample sizes, breed and statistical models. In addition, there is a high variability in the estimates obtained in different studies, mainly due to genetic differences within and between breeds [\(Akanno](#page-6-1) *et al.*, 2013). Therefore, a meta-analysis considering the variability between studies appears as a practical and efficient solution (Sutton *et al.*[, 2000\).](#page-7-5)

Meta-analysis based on random-effects models is a statistical tool used to provide estimates closer to the true unknown parameters ([Oliveira](#page-7-6) *et al.*, 2017). According to [Borenstein](#page-6-2) *et al.* (2009), the results of meta-analysis using random-effects models can be considered more reliable than those obtained from individual studies and may be applied to a large reference population. However, meta-analysis for genetic parameters of economic traits in rabbits has not yet been reported in the literature, being a field of research to be explored in the genetic improvement of rabbits.

The aim of this study was to perform a meta-analysis of genetic parameters (heritability and genetic correlation) for traits of economic importance in rabbits using random-effects models.

MATERIALS AND METHODS

Literature search and Traits

Initially, a literature review was performed to identify all references that reported estimates of genetic parameters (heritability and genetic correlation) for economic traits in rabbit populations. The traits analysed were: litter size at birth (LSB), litter weight at birth (LWB), litter size at weaning (LSW), litter weight at weaning (LWW) and slaughter weight (SW). The following search strategies and selection criteria were considered for the review of this study: (1) search for articles published from 1992 to 2022 on access platforms: Web of Science, PubMed, Google Scholar and Scopus; (2) search terms applied to extract potentially relevant articles, such as "genetic parameters", "heritability", "genetic correlation", "rabbits", "growth traits", "litter size at birth", "litter weight at birth", "litter size at weaning", "litter weight at weaning" and "slaughter weight". Alternative or combined words were also allowed and (3) only scientific articles published with informative descriptions for the estimates were considered. Studies of all formats and languages were admitted according to the search terms mentioned above.

Data recording and exploratory analysis

A database was constructed containing genetic parameters information for all traits evaluated. In addition, the following were recorded: year of publication, journal name, number of phenotypic records, phenotypic mean, standard deviation, coefficient of variation and model statistical details. Genetic parameter estimates were derived from different methods: Henderson's method, restricted maximum likelihood method or Bayesian inference in a mixed animal model. For some published genetic parameter estimates, the standard error (SE) was not reported, and in these cases the approximate standard error was calculated using the pooled variance method, as described by [Sutton](#page-7-5) *et al.* [\(2000\):](#page-7-5)

$$
SE_{ij} = \sqrt{\left(\sum_{k=1}^{k} s_{ik}^2 n_{ik}^2 / \sum_{k=1}^{k} n_{ik})/n'_{ij}\right)}
$$
(1)

where $S\!E_j$ is the predicted SE for the published parameter estimate for the *i*th trait in the *j*th study that has not reported the SE, s_{ik}^{\prime} is the published SE for the parameter estimate for the *i*th trait in the *k*th study that has reported the SE, n_k is the number of records used to predict the published parameter estimate for the *i*th trait in the *k*th study that has reported the SE, and n'_j is the number of used records to predict the published parameter estimate for the *i*th trait in the *j*th study that has not reported the SE.

Most meta-analysis studies in animal breeding do not use the published genetic correlation estimates because they do not follow a normal distribution ([Oliveira](#page-7-6) *et al.*, 2017). Thus, the genetic correlation estimates published were first transformed to an approximate normal scale by using the Fisher's Z transformation, as described below [\(Borenstein](#page-6-2) *et al.* [2009\):](#page-6-2)

$$
Z_{ij,k} = 0.5[\ln(1 + r_{g_{ij,k}}) - \ln(1 - r_{g_{ij,k}})]
$$
\n(2)

where and z_{ijk} and $r_{g_{ijk}}$ are, respectively, the transformed and published genetic correlation estimates among traits *i* and j in the k^{th} study.

As noted by [Hossein-Zahed \(2021\)](#page-7-8), the results of the meta-analysis, such as the estimated parameter and its confidence interval, would then be converted back to correlations for presentation using the following equation:

$$
r_{g_{ij,k}}^* = (e^{2Z_{ij,k}^*} - 1) / (e^{2Z_{ij,k}^*} + 1)
$$
\n(3)

where $r_{g_{ik}}^*$ is the re-transformed genetic correlation estimate and z_{ik}^* is the output from the meta-analysis randomeffects model.

Data quality control

A box plot weighted by the number of records was used to identify possible outliers, which were constructed for each trait assessed. To ensure the reliability of the meta-analysis and avoid biased estimates, a minimum number of scientific articles was calculated for each trait, based on the relative standard error (RSE) (Zarkovich, 1979). A maximum RSE limit of 25% was assumed, as recommended by [Oliveira](#page-7-6) *et al.* (2017), with the higher RSE indicating a greater impact of uncontrolled variation sources on the estimates.

The RSE calculation is obtained as follows (Zarkovich, 1979):

$$
RSE_i = \frac{\left(\frac{S_i}{\sqrt{n_i}}\right)}{\overline{x_i}} \times 100\tag{4}
$$

where *RSE_i* is the relative standard error, *s_i* is the standard deviation (SD) estimated from the published parameter estimates for the *i*th trait, n_j is the number of studies that have reported parameter estimates for the *i*th trait, and \overline{x}_j is *i* the average of parameter estimates for the *i*th trait.

The total number of records for each trait was calculated as the sum of the number of records in each study found. Means and standard deviations were calculated for all traits by using the sample sizes as weights. In addition, the coefficient of variation (CV) was calculated for each trait:

$$
CV_i(\%) = \frac{S_i}{\overline{X}_i} \times 100
$$
\n⁽⁵⁾

where $\mathcal{S}_{_{\!I}}$ is the SD for the *i*th trait, and $\overline{X}_{_{\!I}}$ is the trait mean. *i*

Meta-analysis

Weighted parameters mean estimates were obtained by fitting a random-effect model for all traits studied. Estimates were assumed to be independent and normally distributed through the Box-Pierce and Shapiro-Wilk tests, respectively, using R software ([R Core Team, 2021](#page-7-9)).

The meta-analysis was performed considering the following random-effects model for each trait:

$$
\theta_j = \bar{\theta} + u_j + e_j \tag{6}
$$

where $\hat{\theta}_j$ is the published parameter estimate in the j^{th} study, $\bar{\theta}$ is the weighted population parameter mean, u_j is the among-study component of the deviation from the mean, and *ej* is the within-study component due to sampling error in the parameter estimate in the *j*th study. The *u_j* and e_j were assumed as $u_j \sim N(0,\tau^2)$ and $e_j \sim N(0,\sigma_o^2)$, respectively, where τ^2 is the variance representing the amount of heterogeneity among studies and σ_e^2 is the within-study variance.

The *metaphor* package (Viechtbauer, 2010) available in R software ([R Core Team, 2021\)](#page-7-9) was used to perform the meta-analysis. Forest plots were constructed to indicate the effect size of each study, containing the mean estimates of heritability and genetic correlation with the 95% confidence intervals.

Heterogeneity and publication bias

To quantify the degree of heterogeneity (τ²) between studies and describe the percentage of total variation that is due to heterogeneity rather than chance, the *I*² index was used ([Higgins](#page-7-10) *et al.*, 2003), described as:

$$
I^2 = \frac{Q - df}{Q} \times 100\tag{7}
$$

where *Q* is the statistic [\(Cochran, 1954\)](#page-6-3) given by:

$$
Q = \sum_{j=1}^{J} w_j \, (\widehat{\theta}_j - \bar{\theta})^2 \tag{8}
$$

where *wj* is the parameter estimate weight (assumed as the inverse of published sampling variance for the parameter, (1/*s_j*) in the *j*th study; $\hat{\theta}$, is the parameter estimate published in the *j*th study and $\overline{\theta}$ is the populational parameter weighted mean estimate, both were defined above in the random-effects model. The *df* is the degrees of freedom (*J* – 1, where *J* is the number of used studies) of a Chi-squared distribution assumed for the expected *Q* value on the assumption that $\tau^2 = 0$.

Variations between studies were assessed using the *Q* statistic with a significance level set at 0.05, as it has relatively low power with a small number of studies ([Huedo-Medina](#page-7-11) *et al.*, 2006). The I^2 index was also used to measure the degree of heterogeneity.

In addition, the 95% confidence intervals were considered, and the lower and upper limits will be calculated by:

$$
LL_{\overline{\theta}} = \overline{\theta} - 1.96 \times SE_{\overline{\theta}}, \tag{9}
$$

$$
UL_{\overline{\theta}} = \overline{\theta} + 1.96 \times SE_{\overline{\theta}} \tag{10}
$$

where $S\!E_{\bar{\theta}}$ is the predicted standard error for the estimated parameter $\bar{\theta}$, i.e.:

$$
SE_{\overline{\theta}} = \sqrt{1 / \sum_{j=1}^{j} w_j}
$$
\n(11)

where all terms were previously described.

Egger's linear regression asymmetry was used to examine the presence of publication bias and a *P*-value of 0.05 was set. The trim-and-fill method ([Duval and Tweedie, 2000\)](#page-6-4) was applied when bias was detected (*P*<0.05) to find the number of missing studies. In addition, Funnel plots were used to present the asymmetry. When heterogeneity (*Q* statistic, *P*<0.05) was detected for the analysed parameter estimates, testing for publication bias is inappropriate since it may lead to positive missing results ([Hossein-Zadeh, 2021\)](#page-7-8).

RESULTS

Descriptive statistics

Weighted descriptive statistics for all traits analysed in this study are shown in Table 1. Outliers were found and excluded from the posterior analysis. After quality control, the final dataset contained 147 estimates of heritability

Traits	Abbreviation	Unit	ΝA	ΝR	Mean	SD	CV(%)
Litter size at birth	LSB		26	245780	8.68	1.24	14.27
Litter weight at birth	LWB		14	15995	402.43	197.37	49.04
Litter size at weaning	LSW		20	167978	6.58	1.33	20.27
Litter weight at weaning	LWW		23	104852	2920.53	657.02	22.49
Slaughter weight (individual)	SW			33469	2104.35	329.82	15.67

Table 1: Number of articles (NA), number of phenotypic records (NR), mean, standard deviation (SD) and coefficient of variation (CV%) for each trait evaluated in rabbits.

and 32 estimates of genetic correlation from 34 peer-reviewed scientific articles published between 1992 and 2022 (Supplementary File S1). Among all the traits evaluated, LWB was found to be highly variable with a CV of 49.04% (Table 1). The weighted coefficients of variation for LSB and SW traits were generally low, 14.27 and 15.67%, respectively.

Heritability

The number of contributing estimates (N), the weighted heritability estimates (h²), the relative standard error (RSE) and the heterogeneity of the estimates (based on Q and I^2 statistics) obtained from the random-effects meta-analysis are shown in Table 2.

Heritability estimates were of low magnitude for all traits, ranging from 0.09 to 0.18 (Table 2). The lowest heritability estimate was observed for LSW and the highest for SW. All heritability estimates had low standard errors and were statistically significant (*P*<0.05). In addition, their 95% confidence intervals were narrow and the RSE values were low (<25%), as shown in Table 2. RSE values for heritability estimates ranged from 12.41% (LSB) to 21.68% (LWW). The test for heterogeneity of weighted heritability estimates performed by the *Q* statistic (Table 2) showed significant heterogeneities (*P<*0.05). Overall, the *I²* index showed high values (Table 2) for most traits, LWW (96.37%), LSB (90.36%), SW (78.95%) and LSW (88.56%), indicating substantial heterogeneity among studies, except for LWB (47.70%). This suggests the importance of accounting for this variability in the random-effects model used to estimate weighted means in order to obtain reliable estimates.

Results of heritability estimates for evaluated traits obtained from published papers are shown in the Forest plot available in Supplementary [Figures 1-](#page-8-0)[5](#page-10-0), respectively, for LSB, LWB, LSW, LWW and SW.

Genetic correlation

The mean estimates of genetic correlation for analysed traits are shown in Table 3. Considering the combinations between the five traits studied, a large number of genetic correlations were not available or are limited in the literature. In addition, a few of genetic correlations had to be eliminated from the analysis because their RSE values were greater than 25% (Table 3).

Table 2: Number of estimates (N), heritability (h²), standard error (SE), 95% confidence interval (95% Cl), significance of the random-effect model (*P*-value), *Q* statistic (*Q*) and their significance (*P*-value), relative standard error (RSE %), and heterogeneity index (l^2 %) estimated through meta-analysis based on random-effects models for each trait evaluated in rabbits.

1 LSB: litter size at birth; LWB: litter weight at birth (g); LSW: litter size at weaning; LWW: litter weight at weaning (g); SW: slaughter weight (g).

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LSB: litter size at birth; LWB: litter weight at birth (g); LSW: litter size at weaning; LWW: litter weight at weaning (g).

All genetic correlations for the studied traits were positive and moderate (*P*<0.01), except for the low association between LSB and LWW (0.23). LSB showed genetic correlation with LSW of 0.60 and with LWB of 0.49, while the estimate found between LSW and LWW was 0.56. Most estimates of genetic correlations (Table 3) showed high heterogeneity and significance among studies (*P*-value of *Q* statistic < 0.01 and *I* 2 >98.22%), which justifies the use of a random-effects model. The genetic correlation between LSB and LWW showed no significant heterogeneity (Table 3). Results from the statistical test to evaluate publication bias and the trim-and-fill method to correct funnel plot asymmetry in mean genetic correlation between LSB and LWW are shown in [Supplementary File S2](#page-16-0). The result of Egger's test did not indicate significant publication bias (*P*>0.05) for genetic correlation between LSB and LWW.

Results of genetic correlation estimates among evaluated traits obtained from published papers are shown in the Forest plot available in [Supplementary Figures 6-10,](#page-10-1) respectively, between LSB and LSW, LSB and LWW, LSW and LWW, LSB and LWB, and LWB and LWW.

DISCUSSION

The literature presents a great variability in the magnitude of genetic parameter estimates for the traits evaluated, emphasising the need to use a random-effect model in the meta-analysis study. To support rabbit breeding programmes, it is crucial that meta-analysis provides solid and reliable estimates of genetic parameters ([Oliveira](#page-7-6) *et al.*, 2017). Thus, the random-effect model was able to consider the sources of variation between and within studies, derived from different populations, breeds, designs and statistical methodologies. The low RSE values suggest acceptable variation in the genetic parameters (heritability and genetic correlation) across the studies, allowing the estimation of combined effects. In addition, there was high heterogeneity among the published studies, as indicated by the *Q* and *I*² statistics, reinforcing the importance of adopting the random-effect model.

The high number of studies found indicates that there is greater emphasis on the traits collected at birth and weaning in rabbits, in relation to the other. The broad variation in the traits mean shows that there can be considerable variation in the management conditions under which these populations are reared. Individual weight at slaughter had few studies, despite its economic importance in rabbit meat production, perhaps due to the difficulty of measurement or associated cost.

The low weighted mean of heritability estimates found for all evaluated traits were consistent with the literature ([Sakthivel](#page-7-2) *et al.*, 2017; Farouk *et al.*[, 2022\).](#page-7-3) These traits are largely influenced by non-additive genetic and environmental effects, including management practices. Nevertheless, there is a portion of additive genetic variance acting on the expression of all traits, suggesting potential improvement through genetic selection. The rabbit population will respond better to direct selection for slaughter weight. Alternatively, selecting for litter size at birth, size weight at birth and size weight at weaning would show similar results in response to selection.

Litter size is one of the most important traits in rabbit breeding and is associated with the high prolificacy of the species. In addition, the survival rate of litters during the lactation period, recorded through the number of animals and weight at weaning, should be highlighted. Rabbit breeding programmes focused on meat production have been established through specialised lines by genetic selection (Moura *et al.*[, 2001\)](#page-7-12). Different maternal lines are often selected for reproductive traits such as number of live births, litter size at weaning and weaning weight ([Nagy](#page-7-13) *et al.*,

[2011\).](#page-7-13) The direction of selection in paternal lines is focused on improvements in growth and carcass traits [\(El-Attrouny](#page-6-5) [and Habashy, 2020\).](#page-6-5) Both lines (maternal and paternal) are at the top of the pyramid in breeding programmes as specialised lines and represent the core populations (Moura *et al.*[, 2001\).](#page-7-12) The development and maintenance of these lines is a crucial activity for the success of the programmes. To this end, it is necessary to establish direct or indirect selection strategies for litter size traits at birth or weaning ([Ragab and Baselga, 2011\).](#page-7-14)

In this context, it is essential to know the genetic associations between the traits used as selection criteria within the lines. According to meta-analysis results, genetic correlation between litter size at birth is positive and moderate with litter weight at birth (0.49) and litter size at weaning (0.60). Based on these results, it can be suggested that selection for litter size at birth would be effective to improve the other traits, and to a lesser extent on litter weight at weaning (0.23). Similarly, it is suggested that selection for litter weight at weaning should achieve moderate genetic gains in the traits of litter weight at birth (0.44) and litter size at weaning (0.56). Thus, a rabbit breeding programme can use a selection index that combines these traits related to litter size and weight at birth and at weaning to optimise the improvement obtained in prolificacy and reproduction traits, as well as in growth rate (Farouk *et al.*[, 2022\)](#page-7-3). According to Moura *et al.* [\(2001\),](#page-7-12) the development of a multi-purpose line could be an interesting alternative through simultaneous selection for prolificacy and growth performance traits in situations where it is not possible to select and maintain specialised sire and dam lines for a subsequent crossbreeding programme. Further research is needed to investigate the effects of selection for the traits evaluated in this study on other economic traits included in the breeding objective.

CONCLUSION

The meta-analysis study provided reliable estimates of heritability and genetic correlations for economic traits in rabbits. There is genetic variability concerning the expression of traits litter size at birth, litter weight at birth, litter size at weaning, litter weight at weaning and slaughter weight. Therefore, improvement in these traits can be achieved by genetic selection. In addition, all traits are genetically associated, suggesting that indirect selection will be efficient way to increase prolificacy and body weight in rabbit production.

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Supplementary Figure 1: The forest plot of individual studies and the overall outcome for heritability estimates of LSB in rabbit. The horizontal bars represents the 95% confidence intervals for the study.

Supplementary Figure 2: The forest plot of individual studies and the overall outcome for heritability estimates of LBW in rabbit. The horizontal bars represents the 95% confidence intervals for the study.

Supplementary Figure 3: The forest plot of individual studies and the overall outcome for heritability estimates of LSW in rabbit. The horizontal bars represents the 95% confidence intervals for the study.

Supplementary Figure 4: The forest plot of individual studies and the overall outcome for heritability estimates of LWW in rabbit. The horizontal bars represents the 95% confidence intervals for the study.

Supplementary Figure 5: The forest plot of individual studies and the overall outcome for heritability estimates of SW in rabbit. The horizontal bars represents the 95% confidence intervals for the study.

Supplementary Figure 6: The forest plot of individual studies and the overall outcome for genetic correlation estimates between LSB and LSW in rabbit. The horizontal bars represents the 95% confidence intervals for the study.

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Supplementary Figure 7: The forest plot of individual studies and the overall outcome for genetic correlation estimates between LSB and LWW in rabbit. The horizontal bars represents the 95% confidence intervals for the study.

Supplementary Figure 8: The forest plot of individual studies and the overall outcome for genetic correlation estimates between LSW and LWW in rabbit. The horizontal bars represents the 95% confidence intervals for the study.

Supplementary Figure 9: The forest plot of individual studies and the overall outcome for genetic correlation estimate between LSB and LBW in rabbit. The horizontal bars represents the 95% confidence intervals for the study.

Supplementary Figure 10: The forest plot of individual studies and the overall outcome for genetic correlation estimate between LBW and LWW in rabbit. The horizontal bars represents the 95% confidence intervals for the study.

SUPPLEMENTARY FILE 1

Table S1: Details of published studies used in meta-analysis.

Number	Author (year)	Model	Method	Sample size	Breed		
Litter size at birth							
1	Shehab El-Din (2022)	Animal	REML	441	Multiracial		
\overline{c}	Rastogi et al. (2000)	Animal	REML	1118	Multiracial		
3	Abdel-Kafy et al. (2012)	Animal	REML	642	Baladi Black		
4	Abou Khadiga et al. (2012)	Animal	REML	1400	Multiracial		
5	Behiry et al. (2021)	Animal	REML	330	Multiracial		
6	Farouk et al. (2022)	Animal	REML	625	New Zealand		
7	Youssef et al. (2008)	Animal	REML	2834	Multiracial		
8	Iraqi (2008)	Animal	REML	364	New Zealand		
9	Sorensen et al. (2001)	Animal	REML	808	White Danish		
10	Ragab and Baselga (2011)	Animal	REML	47132	Multiracial		
11	Al-Saef et al. (2008)	Animal	REML	3496	Multiracial		
12	García and Baselga (2002a)	Animal	REML	12651	Multiracial		
13	Ayyat et al. (1995)	Animal	REML	519	New Zealand		
14	Moustafa et al. (2014)	Animal	REML	3144	Multiracial		
15	Nagy <i>et al.</i> (2014)	Animal	REML	11582	Pannon White		
16	García and Baselga (2002b)	Animal	REML	9842	Multiracial		
17	Nagy et al. (2011)	Animal	REML	15900	Multiracial		
18	Odubote and Somade (1992)	Sire	REML	260	Multiracial		
19	Ezzeroug et al. (2019)	Animal	REML	3242	Multiracial		
20	Montes-Vergara et al. (2021)	Sire	REML	210	New Zealand White		
21	El-Deghadi (2019)	Animal	REML	765	New Zealand White		
22	El-Attrouny and Habashy (2020)	Animal	REML	3214	New Zealand White		
23	Rabie et al., (2019)	Animal	REML	896	Multiracial		
24	Nguyen et al. (2017)	Animal	REML	5830	Pannon Large		
25	Piles et al. (2006)	Animal	REML	35776	Synthetic Lines		
26	Ziadi et al. (2013)	Animal	BAYESIAN	3563	Synthetic Lines		
Litter weight at birth							
1	Shehab El-Din (2022)	Animal	REML	441	Multiracial		
\overline{c}	Abdel-Kafy et al. (2012)	Animal	REML	642	Baladi Black		
3	Abou Khadiga et al. (2012)	Animal	REML	1400	Multiracial		
4	Behiry et al. (2021)	Animal	REML	330	Multiracial		
5	Farouk et al. (2022)	Animal	REML	625	New Zealand White		
6	Youssef et al. (2008)	Animal	REML	2833	Multiracial		
$\overline{7}$	Iragi (2008)	Animal	REML	364	New Zealand White		
8	Al-Saef et al. (2008)	Animal	REML	3496	Multiracial		
9	Ayyat <i>et al.</i> (1995)	Animal	REML	519	New Zealand White		
10	Odubote and Somade (1992)	Animal	REML	260	Multiracial		
11	Montes-Vergara et al. (2021)	Animal	REML	210	New Zealand White		
12	El-Deghadi (2019)	Animal	REML	765	New Zealand White		
13	El-Attrouny and Habashy (2020)	Animal	REML	3214	New Zealand White		
14	Rabie et al., (2019)	Animal	REML	896	Multiracial		
	Litter size at weaning						
1	Sorensen et al. (2001)	Animal	REML	1021	White Danish		
\overline{c}	Shehab El-Din (2022)	Animal	REML	423	Multiracial		

META-ANALYSIS' REFERENCE

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SUPPLEMENTARY FILE 2

Table S2: Egger's test, number of missing studies, mean, and 95% confidence interval (95% CI) estimated through meta-analysis.

For traits, see Table 1. Missing: number of missing studies.

Figure S2: Funnel plot of Fisher's *Z* for the genetic correlation between LSB and LWW in rabbit.