





MASTER INTERUNIVERSITARIO EN MEJORA GENÉTICA ANIMAL Y BIOTECNOLOGÍA DE LA REPRODUCCIÓN

EVALUATION OF THE POTENTIAL USE OF MULTIPLE POPULATIONS IN GENOMIC EVALUATION

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ABSTRACT

Marker-assisted selection has been introduced into the breeding programs of livestock populations during the last few years. Particularly, the cattle industry was the first to incorporate the Genomic Selection (GS) into its breeding programs. GS has been proved to be a very useful tool for increasing the efficiency of the breeding programs of dairy cattle but until now it has not given the expected results in beef cattle. The structure of the beef cattle industry consists of many small populations. The absence of a large reference population and the low predictive ability of one population over another prevent the obtainment of accurate predictions. Combining information of animals from different populations in order to construct a large, reference population has been proposed in order to increase the accuracy of the predictions for all populations.

In this study, 4 purebred populations were simulated. Two of these populations diverged under selection, each for a different trait, and the other two under random mating. From these purebred populations 6 admixed populations were created combining them in pairs and 1 admixed population that included data from all purebreds. The predictive ability of all purebred and admixed populations was evaluated under two scenarios of marker density, two scenarios of number of genes that control the traits and 3 scenarios of number of generations of divergence.

The results of this study show that training in purebred populations gives good accuracies when validating in the same population, although slightly lower, and declining with the generations, when the population is under selection for the trait of interest. Training in admixed populations gives higher accuracies for the breeds that participate in the training set than for the ones that do not, but these accuracies are significantly lower than those obtained from training in purebred population. Finally, when the training set is made of all 4 breeds, the accuracies are slightly better for all breeds that those obtained with the 2-breed admixed population.

When the marker density is higher the accuracies are better for all purebred and admixed populations but the admixed populations are more favored than the purebreds. The number of genes that control the traits seems to affect the predictions of the populations under selection but not the others.

RESUMEN

En los últimos años, la selección asistida por marcadores se ha introducido en los programas de mejora genética animal. En particular, el sector de vacuno de leche ha sido el primero en introducir la selección genómica (GS) en sus esquemas de mejora. En ellos, se ha probado que la selección genómica es una herramienta de gran interés. Por el contrario, su potencial utilización en vacuno de carne todavía no es evidente, debido a la presencia de varias poblaciones, la ausencia de una población de referencia grande y la baja capacidad predictiva de una población en otra. Como alternativa, se ha propuesto la creación de una metapoblación a partir de la información disponible en varias poblaciones, que permita incrementar la precisión en todas ellas.

En este trabajo, se han simulado 4 poblaciones puras partiendo de un único tronco evolutivo. Dos de estas poblaciones se separaron mediante selección, cada una para un carácter diferente, y las otras dos exclusivamente mediante deriva genética. A partir de estas 4 poblaciones, se generaron 6 poblaciones compuestas o metapoblaciones de dos poblaciones combinándolas por parejas y una metapoblación compuesta por las 4 subpoblaciones. La capacidad predictiva de todas las poblaciones puras y compuestas se evaluó bajo dos escenarios de densidad de marcadores y de número de genes y 3 escenarios de número de generaciones de divergencia.

Los resultados del trabajo mostraron que cuando se utilizó la información de las poblaciones puras se obtuvieron buenos resultados en la propias poblaciones, aunque decrecieron lentamente a medida que se incrementaron las generaciones de selección sobre el carácter de interés. Las predicciones obtenidas en las metapoblaciones fueron más altas sobre las poblaciones que participan en su formación, y mucho menores para las demás poblaciones. Finalmente, las predicciones a partir de la metapoblación compuesta de las cuatro poblaciones fueron ligeramente mejores que a partir de las compuestas de dos poblaciones.

Cuando la densidad de marcadores se incrementó, los resultados fueron mejores tanto para las poblaciones puras como para las compuestas, pero el incremento fue superior para las metapoblaciones. Por otra parte, el número de genes implicados en las poblaciones afectó solamente los resultados de las poblaciones seleccionadas.

INTRODUCTION

Anessential task in the development of breeding programs in livestock populations is the definition of objectives and criteria of selection. The objectives of selection are composed by traits that define the general goal for the breeding program, with the aim to obtain better individuals for production. The criteria of selection are the measurable traits that permit to identifythe best individuals order to be used as reproductive individuals to generate the following generation. The simplest criteria of selection are defined by the information provided by only one trait and the methods of selection differ one from the other according to the sources of information used.

Mass selection

The first and most simple method was individual or mass selection (Falconer and Mackay, 1996). This method uses as criterion of selection exclusively the phenotypic performance of the candidates to selection. Thus, the best individuals are chosen as reproductive individuals assuming that a better phenotype reveals a better breeding value. The efficiency of this method depends largely on the heritability of the trait considered. Heritability is defined as the ratio of the additive variance and the phenotypic variance of a trait. It explains the proportion of the total phenotypic variability caused by the additive action of genes. Heritability's most important aspect comes from its predictive function, since it represents the degree of reliability of the phenotypic value as an indicator of the breeding value. Individual selection is quite efficient when the heritability is high but gives poor results when it is low.

To overcome the obstacle of the low accuracy of the predicted breeding values caused by the low heritability of some traits, the mean performance of repeated records of the candidates can be used, when available, as criterion of selection. The justification for the use of this method is the assumption that all the phenotypic records of an individual are controlled by the same genetic and permanent environmental effects and that the variation observed is only produced by temporary environmental effects or residuals. The last ones are averaged when the mean performance of various records is used. Thus, when the residual variation is large (or heritability is low) the increase of accuracy with the number of records very important. However, when the residual variance is low (or heritability is large), it provides barely any additional information. The main disadvantage of this method is that it extends the generation interval that results in the reduction of the annual genetic gain.

The annual genetic gain is a measurement of the expected progress of a breeding program (Falconer and Mackay, 1996) and it is defined as:

$$\Delta G = \frac{i\rho\sigma_a}{L}$$

where ΔG is the expected genetic gain, *i* is the intensity of selection, ρ is the accuracy of the prediction, σ_a is the additive standard deviation and *L* is the generation interval. Intensity of selection (*i*) is a standardized measure of the difference between the average performance of the selected individuals with respect to the whole population. The sense of the standardization is to have a dimensionless parameter that can be compared between breeding programs. Accuracy of prediction (ρ) is the correlation coefficient between the predicted breeding value and the true breeding value. The additive standard deviation (σ_a) is the square root of the additive variance of the trait considered. Generation interval (*L*) is defined as the average age of both parents when progeny are born.

Index of selection

Later on, new statistical developments allowed the use of other sources of phenotypic information. Data from the relatives of the candidates were incorporated to the breeding programs with the purpose of increasing the accuracy of the predicted breeding values. The increase in the accuracy is a result of the fact that related individuals share genes and in many cases also the same environment (ex. piglets from the same litter or piglets from different litters that share a common sire). This resemblance between the related individuals can contribute with additional information over the genetic merit of the candidates. The types of relatives more frequently used are full-sibs or half-sibs, parents and progeny because the closer the genetic relatedness between the individuals, more genes they share and more new information is contributed to the prediction of the breeding value. For selection purposes, an index is built using all phenotypic information available from the candidate and its relatives (Lynch and Walsh,1998), weighted according to their degree of genetic relatedness. When the true variance components are used, this index is the best lineal prediction of an individual breeding value and its properties include the maximization of the correlation between the true

breeding value and the index (Lynch and Walsh, 1998). An important limitation to the use of the selection index is the need for the data to be adjusted for systematic effects before its application. These systematic effects are generally unknown and their estimation can result to be a difficult task, especially when no prior data exist. Further, it requires the inverse of the covariance matrix of the phenotypic observations that may be impossible to calculate for large data sets.

Nonetheless, the availability of data of the relatives made it possible to apply methods of selection for traits that are not expressed by the candidates to selection. Some examples of these cases are the prediction of the breeding value of males for milk production, the evaluation of individuals for meat quality traits and the selection for longevity. In all cases, the weighted average performance of the relatives is the criterion of selection.

Multiple traits

In animal production, the economic gain may depend on various productive and reproductive traits. Thus, the objective and the criteria of selection can be composed by several traits and the total genetic merit of an individual is not only the additive value for one trait but the sum of the additive values of several traits adjusted by economic weights. The economic weight of a trait represents the increase in the economic gain due to the genetic improvement of a single unit of the trait when the mean of the rest of the traits does not change (Weller, 1994).

The selection index methodology allows the use of information of several traits to estimate the total genetic merit of the candidates, from its own information and from its relatives. In fact, the index is calculated as the sum of phenotypic records of various traits weighted by regression coefficients for each individual. The goal is to obtain those regression coefficients that allow the index to be the minimum quadratic estimator of the total genetic merit. The accuracy of the index is affected by the economic weights, the phenotypic and genetic parameters that are not known and have to be estimated previously.

Other alternatives for multiple trait selection are tandem selection (Falconer and MacKay, 1996), where one generation is devoted for selection for one trait and the following one for another trait, and selection by independent levels, where minimum limits of performance are established for every trait and the candidates have to overpass

them in order to be selected. The advantage of both methods is that they are easier to apply since not all animals' records are required. The efficiency of the individual index is always better because it uses all the information available, while the efficiency of the other two methods depends on the genetic characteristics of the traits under selection.

Best Linear Unbiased Predictor (BLUP)

As an extension of the selection index theory, Henderson (1949) developed the methodology named Best Lineal Unbiased Prediction (BLUP). This methodology permits the estimation of systematic effects and the prediction of breeding values simultaneously. With known variances components, the properties of BLUP include the maximization of the correlation between the true and predicted breeding value. Under BLUP, the predictors are lineal functions of observations and the estimations are unbiased ($E(a|\hat{a}) = \hat{a}$). Later on, the same author also presented the mixed-model equations that give the estimations of systematic effects and the predictions of the breeding values without the need to calculate the inverse of the observation matrix (Henderson, 1950). The mixed-model equations are:

$$\begin{bmatrix} X^{\mathbf{X}}R^{-1}X & X^{\mathbf{X}}R^{-1}Z \\ Z^{\mathbf{X}}R^{-1}X & Z^{\mathbf{X}}R^{-1}Z + G^{-1} \end{bmatrix} \begin{bmatrix} \widehat{b} \\ \widehat{a} \end{bmatrix} = \begin{bmatrix} X^{\mathbf{X}}R^{-1}y \\ Z^{\mathbf{X}}R^{-1}y \end{bmatrix}$$

Where

y is a vector of n observations,

 $\hat{\boldsymbol{b}}$ is a vector of estimates of p levels of systematic effects,

 \hat{a} is a vector of predictions of q breeding values,

X is a design matrix $n \ge p$ which relates records to systematic effects, and X is it's transposed matrix,

Z is a design matrix $n \ge q$ which relates records to breeding values, and **Z** is it's transposed matrix,

 R^{-1} is the inverse of the residual variance-covariance matrix,

 G^{-1} is the inverse of the additive genetic variance-covariance matrix.

Since \mathbf{R}^{I} is an identity matrix because residual effects are independent one from the other, it can be factorized from both sides of the equation to give the equivalent mixed-model equations:

$$\begin{bmatrix} X^{\mathsf{X}} & X^{\mathsf{X}} \\ Z^{\mathsf{X}} & Z^{\mathsf{X}} \end{bmatrix} \begin{bmatrix} \widehat{b} \\ \widehat{a} \end{bmatrix} = \begin{bmatrix} X^{\mathsf{Y}} \\ Z^{\mathsf{Y}} \end{bmatrix}$$

With $\alpha = \sigma_e^2 / \sigma_\alpha^2$ and A^{-1} being the inverse of numerator relationship matrix which indicates the additive genetic relationship among individuals.

Computing the numerator relationship matrix (A) and inverting it can be very difficult for the large data sets that are normally used for the genetic evaluations. Fortunately, Henderson (1976) proposed a recursive method to calculate the A matrix and its inverse just from the pedigree, when inbreeding is ignored. To take account of the effect of inbreeding, Henderson's (1976) approach needs the diagonal elements of the relationship matrix to calculate the A^{-1} matrix. Later on, Quaas (1976) generalized the procedure and allowed the fast computation of inbreeding coefficients or the construction of the A^{-1} matrix, without setting up the relationship matrix first.

The use of the mixed-model equations, for the prediction of breeding values and the estimation of systematic effects, has been extended to several models, such as:

Sire Model

The sire model was applied in the dairy cattle industry for the evaluation of the sires for milk production. In the Sire model, only the genetic effects of the sires are included in the model and the progeny records are used for the evaluation. However, the genetic merit of the dams is not accounted for as it is assumed that all mates are of similar genetic merit. The advantage of this model is that the number of equations is reduced compared to the animal model.

Animal Model

This model includes as random effects the additive values of all individuals in the population, regardless the number of systematic effects that have to be considered, and the residuals. It is the model more widely applied for several traits and animal species. It assumes that the additive variance and the residual variance are independent.

Repeatability Model

This model is applied when multiple records from an individual are available. It includes as random effects the additive effects and the permanent effects. It assumes that the genetic correlation and the environmental correlation between the pairs of records for each individual equals to 1. It predicts not only breeding values but also permanent environmental effects.

Reduced animal model

The reduced animal model (Quaas and Pollak, 1980) was developed to reduce the computational burden of the full animal model. This procedure allows setting up equations only for the parents and the breeding values of the progeny are obtained by back-solving from the parental breeding values.

Animal model with groups

The animal model with groups of unknown parents (Westell and Van Vleck, 1988) avoids assuming the same average breeding value for all the founders. Thus, they must be grouped according to the year of birth and genetic origin, and a genetic mean of breeding values is estimated for each group.

Model with common environment effects

The resemblance between relatives is produced by sharing genes and also by sharing common environment. This model includes as random effects and predicts the common environmental effects apart from the breeding values. It is used mostly in species with large families such as pigs and chickens.

Multivariate Animal Model

The multivariate Animal Model (Henderson and Quass, 1976) expands the mixed model methodology to several traits. As a consequence, the additive and residual (co-) variance matrices between traits are involved in the mixed model equations.

Animal Model with Maternal Effects

Some traits are affected by the individual's genes but also by environmental effects that are controlled by the genes of the mother. In the animal model with maternal effects (Quaas and Pollak, 1981), the random effects included in the model are the additive value of the individual (direct effect), the additive value of the mother to produce the

suitable environment (maternal genetic effect) and permanent environmental effects that have an influence on the mother (maternal permanent effect). It is assumed that there is a genetic covariance between direct and maternal effects and there is no correlation between the other effects.

Non Additive Animal Model

The non-additive Animal Model (Henderson, 1985) expands the mixed model equations to include other genetic components such as dominance or epistatic interactions. The procedure developed by Hoeschele and Van Raden (1991) also permits to invert the dominance relationship matrix of large data sets for pedigree files when there is no inbreeding in the population

Marker Assisted Selection

Since the 90's, molecular information is available due to the advances of techniques of molecular biology. This new source of information gives the opportunity to enhance the response to selection by incorporating it to traditional breeding programs, especially for traits that present difficulties in their improvement by traditional selection. Such traits are those with extremely low heritability (ex. Reproductive traits) and traits whose phenotypes are difficult to obtain (ex. Disease resistance or meat quality).

Following Dekkers (2004), the genetic markers used for selection can be classified into 3 different types, direct markers (or genes), linked markers in population-wide linkage disequilibrium with QTLs (LD markers), and linked markers in population-wide linkage equilibrium with the QTLs (LE markers). Note that a QTL is denoted as a segregating gene affecting the trait of interest. A favorable aspect of direct markers is that are easier to incorporate to an existing breeding program and their potential for increasing the response to selection is greater than the other types of markers. Nevertheless, the difficulty of detecting these genes makes more appealing the use of LD or LE markers. (Dekkers, 2004)

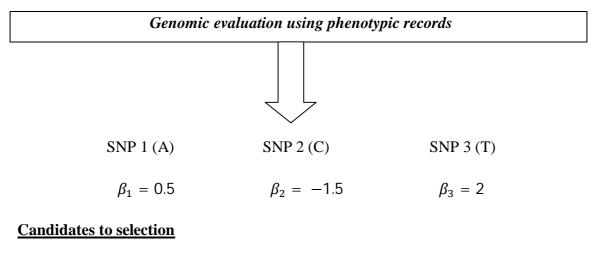
Genomic Selection

Meuwissen et al. (2001) proposed the use of a dense marker map for the prediction of total breeding values and, since then, the procedure is denoted as *Genomic Selection*. The development of commercial genotyping chips made available thousands of single

nucleotide polymorphisms (SNP) markers. It is expected that some of the SNP markers are located near the quantitative trait loci (QTLs) of traits of interest and due to linkage disequilibrium (LD) between them, they should be inherited jointly. In this way, all of QTLs affecting a trait may be in LD with one or more markers. As a consequence, if there are enough markers to cover the whole length of the genome, the additive effects of the QTL can be captured by the markers without the necessity of locating them. Genomic selection procedures consist in estimating the effects associated to the markers in a reference population where genotypes and phenotypes are available for all individuals. These estimates are then used to predict the genetic merit of young individuals with no records available except their genotype. Figure 1 illustrates this genomic selection strategy:

Reference population								
	SNP 1 (A/T)	SNP 2 (C/G)	SNP 3 (T/G)					
Animal 1	AA (2)	CG (1)	GG (0)					
Animal 2	AT (1)	GG (0)	TT (2)					
Animal 3	TT (0)	CC (2)	TG (1)					
Animal n	AT (1)	CG (1)	TG (1)					

Figure 1.Example of genetic selection.



	SNP 1	SNP 2	SNP 3	GBV^*
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Candidate 1	(AA) 2*0.5	+	(GG) 0*(-1.5)	+	(TG) $1*2 = +3$
Candidate 2	(TT) 0*0.5	+	(CC) 2*(-1.5)	+	(GG) 0*2 = -3
Candidate 3	(AT) 1*0.5	+	(CG) 1*(-1.5)	+	(TT) 2*2 = +4.5

*Genomic Breeding Value

From the example above, n individuals with phenotypic records are genotyped for 3 SNP markers. The genotypes are codified as 0 and 2 for the two types of homozygous and as 1 for the heterozygous markers. After the genomic evaluation, the estimated effects associated to each marker are obtained. Then, the genotypes of 3 candidates, which are codified in the same way, are multiplied by the estimates of each marker and summed in order to obtain the genomic breeding values. The candidate 3 would be selected for reproduction as the best candidate to selection.

In this example, the marker-effects to be estimated are only 3 and the phenotypic data available for this purpose are much more. Therefore, their estimation presents no difficulty. In reality, thousands of SNP markers are available and their effects need to be estimated from a limited number of phenotypic records. This leads to a situation known as "the large p small n problem". Traditional methods based on marker regression are unable to deal with this problem unless they introduce some shrinkage on the estimation of the effects. Several methods have been developed to address this problem and to perform genomic selection in practice:

Least squares

This method treats the marker effects as fixed. To deal with the large number of parameters to estimate from few data, a pre-selection of SNPs has to be applied before the analyses. The SNPs with a larger effect are included in the estimation assuming that the rest of the markers do not have any effect. In the original article of Meuwissen et al. (2001), this method gave accuracies of the predicted breeding values of 0.318. The low performance is probably due to the few degrees of freedom available for the estimation of every marker-effect, even after the pre-selection was applied.

Genomic BLUP

G-BLUP is a method similar to the traditional BLUP described by Henderson (1975) that uses a genomic relationship matrix instead of the pedigree relationship matrix. The genomic relationship matrix is built from molecular information in a way that, individuals that share identical by state genotypes for a larger number of markers are more similar and therefore, have larger values in the corresponding cell of the matrix. Misztal et al. (2009) proposed a modification of this method using the pedigree relationship matrix combined with the genomic relationship matrix with weighting parameters between them. The Genomic BLUP methodology does not suffer the large p small n problem since the amount of unknown effects is usually the same as in traditional BLUP (González-Recio et al., 2008).

Bayes A

Originally proposed by Meuwissen et al. (2001), this method assumes a normal distribution on the marker effects with zero mean and variance σ_i^2 associated to each marker. A scaled inversed Chi-squared distribution is assumed for this variance with 4.012 degrees of freedom and scale parameter 0.002. These values were used because they fitted the simulation study of the authors and gave accuracies of 0.8. Since then, they have been applied in many cases without a clear justification.

Bayes B

This method assumes a normal distribution on the marker-effects and variance σ_i^2 associated to each marker just as Bayes A. Bayes B differs from Bayes A when to the assumptions made for the distribution of the variance. A mixture of distributions on the variance is assumed, where the variance is zero with probability π and distributed as in Bayes A with probability 1- π . Although the election of π is arbitrary with no justification, Bayes B provides highly accurate predictions. Meuwissen et al. (2001) obtained accuracies of 0.85 with Bayes B, the highest among the ones obtained using Bayes A, BLUP or least-squares.

Bayes $C\pi \& D\pi$

To address some drawbacks of Bayes A and Bayes B, such as the prior probability of π and the hyper-parameters of the prior distribution of the variance, Habier et al. (2011) described the Bayes $C\pi$ and Bayes $D\pi$ methods. Bayes $C\pi$ method assumes a common variance to all markers with probability 1- π and variance zero with probability π .

Additionally, the proportion π of markers is treated as unknown and is estimated from the data. Bayes $D\pi$ imposes a prior on the scale parameter of the inverse chi-square distribution, which is the prior distribution of the variance of marker-effects. Both methods proved similar to the original methods regarding the accuracies.

Bayesian LASSO

De los Campos et al. (2009) and Usai et al. (2009) proposed the Bayesian LASSO method (Park and Casella, 2008) for genomic selection, where a double exponential prior distribution is assumed for the marker-effects with parameter λ . This method performs a larger shrinkage on the marker-effects than other methods in a way that a large number of markers are estimated with a very small effect, and only a few markers are allowed to have larger effects. The degree of shrinkage is determined by the parameter λ , which has to be estimated previously to the analyses. Park and Casella, (2008) proposed the use of Empirical Bayes by Marginal Maximum Likelihood using an appropriate hyperprior for the estimation of λ . Legarra et al. (2010) proposed a modification of this method (BL2Var) which considers two different variances for the distribution of marker-effects and the residuals. Moreover, there is no need to preestimate the parameter λ as it is estimated from the data simultaneously with the marker effects. Up until now Bayesian LASSO has been widely applied for genomic evaluations as it provides accurate predictions for low density genotyping (Usai et al., 2009) and for traits that are regulated by many genes with a small effect (Cleveland et al., 2010).

Elastic Net and SNP pre-selection

Croiseau et al. (2011) proposed the implementation of the elastic net algorithm for genomic selection. This is a combination of Genomic BLUP and Bayesian LASSO weighted by a parameter α which takes values from 0 to 1. When α =0, a BLUP model is defined whereas α =1 a LASSO model is chosen. Additionally, a pre-selection of markers can be applied prior to the analyses. The purpose of this method is to provide a more flexible tool to deal with the large p small n problem.

Non-parametric methods

The need to improve the accuracies of the predictions of breeding values, led to the introduction to the field of genetics of new methodologies from the machine learning

scope. These algorithms are more attractive for application to multiple and more complex situations found in biological systems as they are able to accommodate additive, dominant effects or even epistatic effects. It is believed that these methods can approach the genetic architecture of a trait more than the linear models. Some methods that have been proposed are: Reproducing Kernel Hilbert Spaces Regression (RKHS) (Gianola et al., 2006) which resulted in accuracies similar or even higher than the ones obtained by the Bayesian methods (Gonzalez-Recio et al., 2009), Random Forest (RF), that gives the possibility of capturing interactions between genes and between genes and environment (Sun, 2010) and Neural Networks (NN), which prove to be useful for predicting complex traits as it can capture non-linear relations (Gianola et al., 2011).

Genomic selection for multiple populations

Until now, Genomic selection has been implemented mainly in the dairy cattle industry with encouraging results. The existence of a large enough reference population, due to the world-wide supremacy of the Holstein-Friesian breed, permits us to achieve highly accurate predictions (Hayes et al., 2009). The beef cattle industry though, does not follow the structure of dairy cattle industry and the construction of a large enough reference population presents serious difficulties, due to the existence of many and small populations. Therefore, the evaluation within-breed gives poor results due to the small size of the training sets. Moreover, the estimations obtained from one breed cannot be applied to other breeds as they give very low accuracies (Harris et al., 2008).

To avoid this inconvenience, De Roos et al. (2009) proposed pooling animals from different breeds to obtain a large training set. In his study, two populations that diverged randomly for several generations were simulated. His results showed that adding individuals from the second population to the training set (composed only by the first population), had some effect on the reliability of the genomic breeding values in the first population and it was most beneficial when the heritability of the trait was low. Furthermore, when the two populations had diverged for only few generations and the marker density was high, the information from the second populations was most valuable.

In another study, Hayes et al. (2009) used 3 reference populations, composed by pure Holstein sires, pure Jersey and a population combined between them. They found that training in the admixed population gave similar accuracies for the Holstein as when training in the pure Holstein population but higher for Jersey. These results show that in situations of small population size, including animals from a different breed may be interesting to increase the accuracy of the predictions.

Further, Kizilkaya et al., (2010) used a multiple population composed by 8 breeds to predict the breeding values for one of the breeds. When the trait of interest had a heritability of 0.5 and the trait was controlled by 50 QTLs, they found accuracies of 0.38. However, when it was controlled by 500 QTLs, they only got an accuracy of 0.22. Their results implies that the predictive ability eroded as the number of QTL increases, and thus, that the genetic architecture of the trait plays an important role on the efficiency of genomic selection for multiple populations.

Finally, Toosi et al. (2010) developed a simulation study that created 4 populations that diverged randomly for 53 generations to set up several types of training sets (admixed and crossbreed). Those training set were used to predict one of the 4 populations. The accuracies obtained were 0.79 when training in purebred, 0.71 when training in admixed that included the validation population and 0.43 when it was not included. These accuracies increased with marker density for all training sets, though with more intensity for the admixed populations. The results of this study proved that marker density also determines the accuracy of genomic selection in multiple populations.

As a general conclusion of these studies, it is known that the predictive ability of the admixed training populations depends on the genetic architecture of the trait and the marker density. It is important to note that the previous simulation studies make the populations diverge exclusively by genetic drift. However, the causes of genetic differentiation of livestock populations include also natural and artificial selection jointly with the genetic drift. Therefore, this study pretends to evaluate the consequences of the number of generations and the cause of reproductive isolation of the populations on the efficiency of genomic selection for multiple populations. In addition, the marker density of genotyping and the number of genes that control the traits will be considered as variables to generate the scenarios of simulation.

OBJECTIVE

The objective of this study is to evaluate the accuracy of the predicted breeding values of purebred populations based on estimates of marker effects in purebred and admixed populations. Moreover, this study pretends to evaluate the influence of the following situations on the accuracy of the predictions and in the persistence of linkage disequilibrium:

- 1) Causes of reproductive isolation of the populations
- 2) Number of generations of divergence.
- 3) Number of genes involved in the genetic determinism of the trait.
- 4) Marker density of the genotyping.

As a side result, this study also will provide software to simulate populations under different scenarios of genetic evolution.

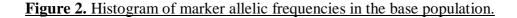
MATERIAL AND METHODS

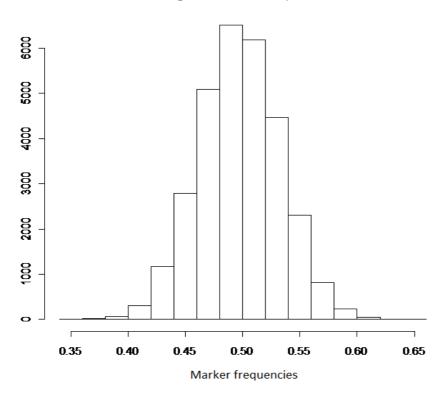
This chapter is structured in three main sections. In first place, the development of the simulation study is described. Secondly, and given the output of the simulation, the Bayesian Lasso method for estimation of markers effects is presented. In third place, the procedures of validation of the genomic prediction and the calculation of the linkage disequilibriumin in single and multiple populations are shown.

Simulation

The simulation study was performed through the development of a Fortran 90 program, which is included in the APPENDIX. The software simulated a base population of 100 unrelated individuals, with sex ratio 1:1. Each individual had 30,000 biallelic markers evenly distributed along the genome. In order to obtain two different scenarios of marker density, a genome of 4 chromosomes and a genome of 20 chromosomes of 1 Morgan each were simulated. Two alleles were considered for each marker and they were coded as 1 and 2. To assign the type of allele to a marker, a random number between 0 and 1 was drawn from a uniform distribution for every allele. The type 1 allele was assigned if the number was lower than 0.5 and the type 2 allele was assigned if the number was greater than 0.5. In this way, the frequencies of the markers in the base population were all near 0.50, as it is presented in Figure 2.

To simulate the next generation, two random numbers, as before, were drawn from a uniform distribution. The first number was used to choose a male and the second one to choose a female. The first marker on every chromosome was used to determine which allele from the pair of chromosomes would be passed on to the next generation. A random number was drawn from a uniform distribution. If it was greater than 0.5 the second allele was passed on and if it was lower the first allele was passed on.



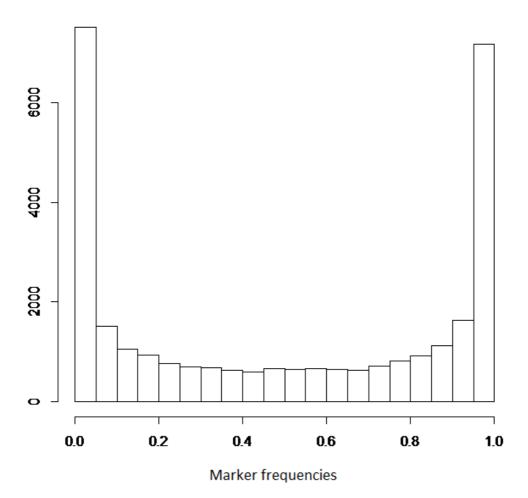


Histogram of marker frequencies

To simulate the recombination of the chromosomes, the number of recombination per chromosome was drawn from a Poisson distribution with parameter $\lambda=1$ (1) recombination per chromosome). If there was no recombination, the whole chromosome was passed on as it was. Further, the points of recombination were simulated by the extraction of as many random numbers from a uniform distribution as the number of recombinations. The markers after the point chosen were changed from the first allele to the second allele and vice versa. After the father's chromosome had been recombined and passed on to the son, the mutation was simulated by sampling the number of mutations from a Poisson distribution with parameter $\lambda=1$ and $\lambda=5$ for the 20 and 4 chromosomes scenario, respectively. A total of 1 and 5 mutations per chromosome were simulated in order to have the same mutation rate per locus under all scenarios of marker density. After some tuning, the mutation rate was set at 6.6×10^{-4} per locus with the objective of maintaining the variance after many generations of evolution. Then the points of mutation where chosen by sampling them from a uniform distribution and the allele of the marker selected was changed from type 1 to type 2 or from type 2 to type 1. This procedure was repeated for the chromosomes that the individual received from the

mother. The new generation was then used to create the next one in the same way. The population evolved for 1000 generations, each time using one generation to produce the next. There was no generation overlap and the size of the population was kept at 100 individuals. The allelic frequencies after 1000 generations of evolution are presented in Figure 3.

Figure 3, Histogram of marker allelic frequencies after 1000 generations



Histogram of marker frequencies after 1000 generations

At the end of the 1000-generation-evolution, two groups of 100 or 300 markers were selected to simulate the causative mutations of two different traits. To select them, we sampled them from a uniform distribution and they were attributed an effect sampled from a normal distribution with mean 0 and standard deviation 1 in order to have positive and negative effects, many markers with a small effect and a few of them with a large effect. Then, we calculated the frequencies of the causative mutations by

counting the number of type 2 allele in the population and dividing it by the total number of alleles. The additive variance that was created for each trait was calculated as

$$var(\alpha) = \sum_{1}^{m} 2p_i(1-p_i)\alpha_i^2$$

where p_i is the *i*th marker frequency, *m* is the total number of causative mutations and α_i is the substitution effect of the marker *i*. The heritability for the trait 1 was set 0.2 and for trait 2 was 0.4. The residual variance for each trait was calculated from the equation

$$h^2 = \frac{var(\alpha)}{var(\alpha) + var(r)}$$

At this point 4 subpopulations (A, B, C and D) of 100 individuals were separated from the base population. The subpopulation sA and B were put under a process of phenotypic selection for the traits 1 and 2, respectively. At the same time, subpopulations C and D did not experience any selection and evolved under genetic drift generated by random mating. This strategy allowed simulating 3 different scenarios of time of divergence for the 4 populations that evolved separately for 5, 50 and 200 generations.

To simulate the evolution of the population A, phenotypic values were assigned to all its individuals for the trait 1 as:

$$y_{lk} = \mu + \sum_{1}^{m} g_{ik} a_i + r_k ,$$

where y_{lk} is the phenotypic value of individual k for the trait 1, μ is the trait mean which was 100, m is the total number of causative mutations, g_{ik} is the genotype of the individual k at the *i*th locus, coded as -1, 0 and 1, a_i is the substitution effect of the *i*th locus and r_k is the residual of individual k sampled from a normal distribution with mean=0 and standard deviation the square root of the residual variance calculated earlier. Then, the phenotypic mean and standard deviation were calculated:

$$mean_p = \frac{\sum_{1}^{k} y}{k}, \qquad sd_p = \sqrt{\frac{\sum_{1}^{k} (y_k - mean_p)^2}{k-1}},$$

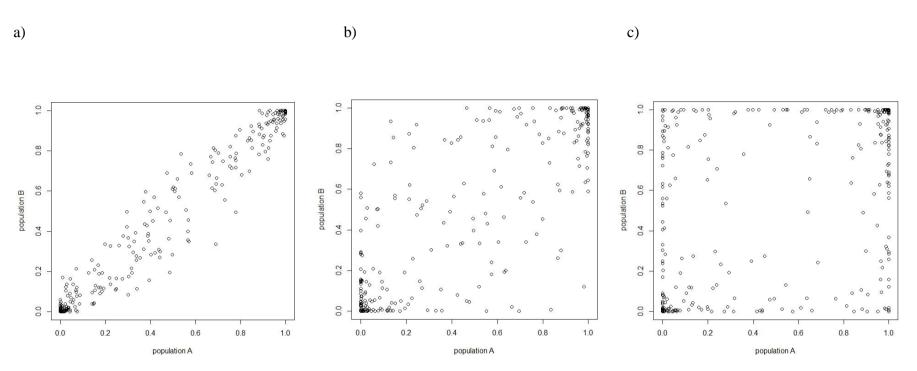
where k is the number of individuals and y_k is the phenotypic value of the individual k. To create the next generation, two individuals with phenotype greater than $(mean_p - mean_p)$ sd_p), were selected by sampling them with equal probability. This fact implies an intensity of selection of 0.30, approximately. The intensity of selection simulated was low in order to prevent the fixation of the alleles and the shrinkage of the additive variance after many generations of selection. The procedure of allele transmission from one generation to the next was the same as described for the base population. After the various generations of evolution, one more generation was simulated, where the size of the population was increased to 2,000 individuals, creating the breed A. Phenotypic values for the second trait were simulated at this point, the same way as for the first trait but with a trait mean μ =1,000. Finally, breeding values were simulated for every individual, for both traits as:

$$tbv = \sum_{l}^{m} g_{l}a_{l}$$
,

where g_i is the genotype of the i^{th} locus coded as -1, 0- and 1 and a_i is the substitution effect of the causative mutation.

Population B was simulated in the same way as population A except for being selected for the trait 2. Populations C and D were left under random mating. The simulation process for them was the same as the one used for the evolution of the base population. One more generation was added to them as well, to create the breeds B, C and D of 2,000 individuals each. Genotypic records, phenotypic records and breeding values for both traits were obtained at this point from all breeds.

Figure 4, Plots of frequencies of the causative mutations of trait 1 between population A and B after evolving for a) 5 generations, b) 50 generations and c) 200 generations.



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Furthermore, 7 admixed populations of 2,000 individuals were created using the 4 pure breeds. 6 of them were all possible combinations of two pure breeds (A+B, A+C, A+D, B+C, B+D and C+D) with 50% of each breed, while the last one was a combination of all pure breeds (A+B+C+D) with 25% of each breed. A first sampling from a uniform distribution gave us the individual, and a second sampling gave us the breed from which it would come from. The genotype and the phenotypes for both traits of the individual selected were copied to the admixed population.

The purebred and the admixed populations would be used as training sets for estimating marker effects while the 4 pure breeds would be used as validation sets to calculate the accuracy of the predictions. Figure 5 shows the procedure of obtaining the final populations and Table 1 shows the parameters used in the simulation study.

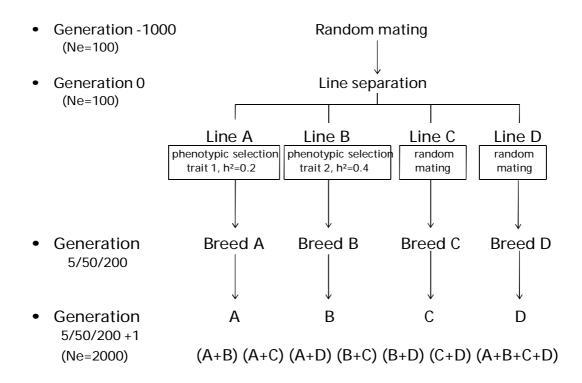


Figure 5.Schematic representation of the simulated population history.

Table 1. The parameters used in the simulation

Genome size	20 M, 4 M
Number chromosomes	20, 4
Number of total markers	30000
Number of causative mutations	100, 300
Marker density per cM	15, 75
Mutation rate per locus	6.6 x 10 ⁻⁴
Distribution of causative marker effects	Normal (0,1)
Generations of divergence	5, 50, 200
Population size	
Generation -1000 to 0	N = 100
Generation 0 to 5/50/200	N = 100
Generation +1	N = 2000
Heritability	0.2, 0.4

Estimation of marker effects

For each case of simulation, estimation of markers effects was performed with the 4 pure and the 7 admixed populations. The method of choice was the Bayesian Lasso (Park and Casella, 2008) with the approach developed by Legarra et al. (2010). Further, the software GS3 (Legarra et. al., 2012) was used. Among the plethora of procedures to estimate marker effects within the scope of genomic prediction, this procedure was chosen because it presents good prediction results when compared with other methods (Usai et al., 2009; Cleveland et al., 2010), and it is available in the commercial software GS3.

In this study, the model of analysis was:

$$y = \mathbf{1}\mu + \sum g_i a_i + e,$$

where y is the vector of phenotypic values of individuals in the training data, μ is a single unknown population mean, **1** is a vector of ones, \mathbf{g}_i is a column vector containing the genotypes (coded as 0, 1, or 2) of each individual at locus *i*, α_i is the random unknown allele substitution effect of marker *i*, and *e* is a random vector of unknown residuals with $e|\sigma_e^2 \sim \text{MVN}(0, \mathbf{I}\sigma_e^2)$. MVN stands for multi-variable normal distribution.

Legarra's et. al. (2010) modification of the Tibshirani's (1996) original Lasso considers two different variances, one for the conditional distribution of marker effects, and another for the residuals. This is more similar to the classical models used in quantitative genetics. The distribution of the marker effects is a conditional Laplace distribution and it can be written as:

$$\boldsymbol{g}|\lambda,\sigma_{\alpha}^{2}\sim\prod_{i}\frac{\lambda}{2\sigma_{\alpha}}\exp\frac{-\lambda|\alpha_{i}|}{\sigma_{\alpha}},\qquad \boldsymbol{e}|\sigma_{e}^{2}\sim\mathrm{MVN}(0,\mathbf{I}\sigma_{e}^{2}),$$

which is equivalent to the original form of Tibshirani's Lasso

$$\boldsymbol{g}|\lambda \sim \prod_{i} \frac{\lambda}{2} \exp(-\lambda |\alpha_{i}|)$$
, $\boldsymbol{e}|\sigma^{2}_{e} \sim MVN(0, I\sigma^{2}_{e})$,

because only the ratio λ/σ_{α} is used and it cannot be estimated separately. Also, λ determines the variance of the marker effects using

$$Var(\alpha) = \frac{2}{\lambda^2}$$

With this modification it's not needed for the λ to be estimated previously as it's estimated simultaneously with the markers effects.

In this study the prior distribution for the residual variance and the genetic variance was assumed to be uniform. The prior for λ was assumed uniform as well, bound between 0 and 1,000,000.

For estimation purposes, the markers selected as causative mutations for the two traits, were excluded from the marker panel, and only the neutral markers were used for the estimation. A Markov chain Monte-Carlo (MCMC) of length 100,000 cycles with a burn-in period of 20,000 cycles was conducted.

Validation of genomic prediction and calculation of linkage disequilibrium

Once estimates of marker effects were obtained from the training sets, the estimated breeding values of individual k (GEBV_k) in the validation data set (A, B, C and D purebred populations) were computed as

$$\mathit{GEBV}_k = \sum_{i=1}^m g_{ik} \, \hat{\mathsf{a}}_i$$
 ,

where g_{ik} is the genotype (-1, 0, and 1) of individual k at locus i, \hat{a}_i is the estimated effect of marker i, and m is the total number of markers. Accuracy was calculated as the

raw correlation between the estimated and the simulated breeding values of individuals in the 4 validation sets. This accuracy was used to compare the performance of the different scenarios and training sets simulated. All scenarios were replicated 10 times, and the results were averaged across replicates.

To evaluate the extent and the magnitude of linkage disequilibrium in the training populations and its impact on the accuracy of genomic predictions, linkage disequilibrium between the neutral markers and the causative mutations was calculated. A small Fortran 90 program was developed to calculate the observed gametic frequencies f_{11} , f_{22} , f_{12} and f_{21} of marker-causative mutation pairs and then the linkage disequilibrium was calculated as described by Falconer, using

$$D = (f_{11} * f_{22}) - (f_{12} * f_{21}).$$

Further, to evaluate the persistence of LD phase across validation and training populations, the marker-causative mutation pairs with D>0.1 and D<-0.1 in the validating populations were chosen and their D in all training population was compared. The comparison was calculated as the percentage of the pairs in strong LD in the validation populations that appeared also in strong LD in the training populations.

Additionally to the *D*, the correlation coefficient between pairs of loci was calculated. The correlation coefficient is a measure of LD, an alternative to *D*, expressed as:

$$r = \frac{D}{\sqrt{p(1-p)q(1-q)}}$$

and its advantage is that it is adjusted to the loci having different allele frequencies. Its permits to distinguish better the higher values of LD and serves as a comparable measure of LD between populations with different allele frequencies.

RESULTS AND DISCUSSION

In this chapter, the results of this study are presented and discussed. The objectives were to compare the predictive ability of several purebred and admixed populations under different scenarios of marker density, number of causative mutations affecting the traits, number of generations of divergence and type of reproductive isolation of the populations. Initially, the scenario with 15 markers per cM and 300 causative mutations for each trait is examined. In first place, the predictive ability of the purebred populations is evaluated. Secondly, the accuracies obtained from admixed populations composed of 2 purebred are compared and finally, the performance of the admixed population composed of all 4 purebred is presented. After that, these results are compared in terms of the accuracies of predicted breeding values with the ones from the two other scenarios. In one case, when there are 75 markers per cM, and, in the other, when only 100 causative mutations per trait are simulated.

For the first case, a population of 100 individuals with 30,000 biallelic markers each, evenly distributed along 20 chromosomes of 1M was simulated. From these markers, two groups of 300 markers were randomly chosen in order to simulate the causative mutations of two traits with heritabilities of 0.2 and 0.4. The population mated randomly for 1,000 generations. At generation 0, 4 new populations of 100 individuals were created from the base population. The first was put under selection for the trait 1, and the second for the trait 2, whereas the last two populations were left to evolve under random mating. After 5, 50 and 200 generations of evolution, the final purebred populations (A, B, C and D) were created by increasing their size to 2,000 individuals each. Moreover, several admixed populations were created combining phenotypic and genotypic records from the 4 purebred populations. There were created 6 admixed populations combining 2 purebred with 50% from each, and one admixed combining all 4 purebreds with 25% from each. The phenotypic and genotypic records of all purebred and admixed population were used for the genomic evaluation while the 4 purebred served as validation sets, as it can be seen in Figure 5 in the previous chapter.

Purebred populations

Table 2 shows the accuracies of the predicted breeding values for the 4 validation sets (A, B, C and D) for both traits, obtained from training in the same 4 populations. The upper half of the table shows the evaluations for the first trait, while the lower half for

the second trait. The validation sets are located in the first column and the training sets are found in the line above the results. The table is separated in 3 parts which show the results obtained for the same populations, under 3 scenarios of number of generations of divergence (5, 50 and 200 generations). The diagonal elements of every sub-table show the accuracies obtained when training and predicting in the same population, while the elements outside the diagonal show the predictive ability of the populations over the others.

The accuracies achieved overall, ranged between -0.023 (Trait 1, prediction in population B and validation in population A with 200 generations of divergence) and 0.822 (Trait 2, prediction and validation in population C) with the accuracies for the second trait being generally higher than the ones for the first trait. This difference in the accuracies between the two traits results from their different heritabilities. As expected, traits with higher heritability give better predictions (Falconer and Mackey, 1996).

When training and predicting in the same population, the accuracies obtained were around 0.7 for trait 1 and 0.8 for trait 2 and were maintained at the same level in all 3 scenarios of generations of evolution. Exceptions to this were populations A and B. Population A showed a decline to its predictive ability for trait 1 of 4.7% and 15.1% after 50 and 200 generations of evolution respectively. Likewise, the accuracy of the predictions for the trait 2 made using population B declined 2.2% and 10.4% for 50 and 200 generations. The loss of the accuracy observed is produced due to the fact that these populations were put under selection for the traits of interest (population A selected for trait 1 and population B selected for trait 2). The process of selection tends to change the allele frequencies of the causative mutations and force them near fixation or loss, as it is shown in Figure 6, for one specific replicate. As a result, there is a loss of genetic variance, as it is illustrated in Figure 7, which is essential for estimating marker effects and predicting breeding values (Falconer and MacKay, 1996). In addition, a slight decrease is also observed for populations A and B for the trait that is not used for selection, because of the lower effective size caused by the selection process.

Figure 6. The histogram of the gene frequencies of trait 1 in a) in the base population and b) after 200 generations of selection.

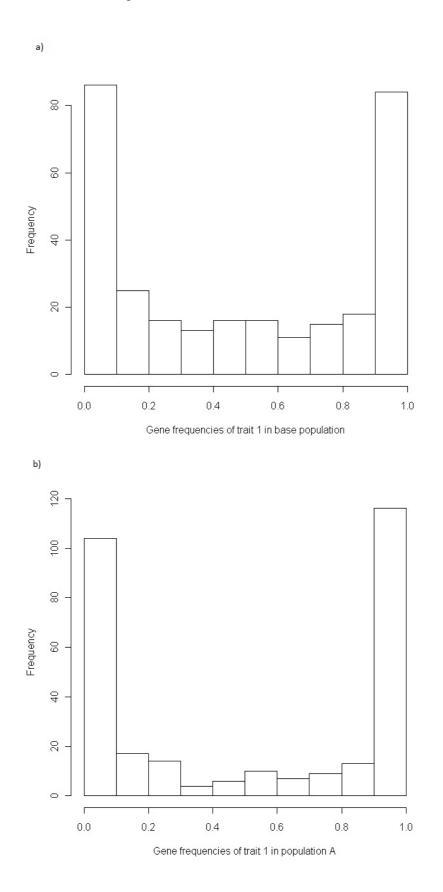
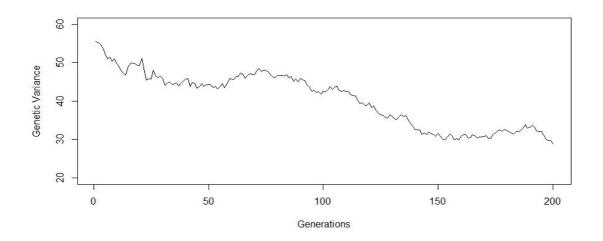


Figure 7. Graphical representation of the loss of genetic variance in 200 generations of selection.



The results show that the predictive ability of one population over the others is low to be used for selection purposes, as it is even lower than the one obtained by mass selection (0.44 for trait 1 and 0.63 for trait 2, respectively). When the populations are separated for 5 generations the accuracies ranged around 0.3 and 0.4 for train 1 and trait 2 respectively, which declined rapidly with the number of generations of separation, being 0.1 and 0.14 after 50 generations and very close to 0 after 200 generations. This occurs because the forces of selection and genetic drift that act, make that the alleles of the genes are fixed or lost in a different way in every population, as it is presented in Figure 4 in the Material and Methods section.

Besides the accuracies, the linkage disequilibrium (LD) between the causative mutations of both traits together and all the markers was calculated using

$$D = (f_{11} * f_{22}) - (f_{12} * f_{21})$$

as described by Falconer and Mackay (1996). The accuracy of the predictions depends largely on the extent of the LD between the markers and the genes. When more markers are in LD with a gene and when the LD is strong, then the accuracy improves. Likewise, when two populations share the same markers, genes and LD patterns then more accurate predictions are expected. Firstly, LD was calculated separately in all populations. Afterwards, the causative mutations-marker pairs that presented a LD with a D higher than 0.1, in the 4 validation sets (A, B, C and D), were selected. These pairs were then looked up in all populations to see how many of them presented also a D higher than 0.1 as in the validation sets and the percentage of common pairs in high LD was calculated.

Table 3 presents the percentage of common pairs in high LD between the validation sets and the training sets under the 3 scenarios of number of generations of divergence. Beneath the percentage of common pairs, the correlation between the LD in the validation set and the LD in the training set of the selected gene-marker pairs is also shown. Like in table 1, the validation sets are located in the first column and the training sets in the line above the results. The elements of the diagonal are all 100% (and 1 for the correlation) because the same population is used as training and validation set.

The gene-marker pairs in high LD that the training sets share with the validation sets reach up to 49%-56% when the populations are separated for 5 generations and drops to 11%-15% after 50 generations and to 2%-4% when they are separated for 200 generations. Likewise, the correlation of the LD between populations is higher than 0.9 in the case of 5 generations of divergence. But drops to half after 50 generations and reaches a value close to 0 after 200 generations. In all cases, the populations that were put under selection (A and B) present the lowest values of percentage and correlation. This occurs because the force of selection acts additionally to the force of genetic drift and creates patterns of LD specific of the population (Falconer and MacKay, 1996).

These results of the LD extent are consistent with the results of the predictive ability of each population. As it can be noticed, when the populations are separated for 5 generations they share around 50% of gene-marker pairs in high LD and the accuracies obtained are around 0.3 (for trait 1), half of the accuracy obtained from evaluating and predicting in the same population. Moreover, the predictive ability of one population over the other declines with the generations of divergence with the same rate as the percentage of gene-marker pairs in high LD that they share.

			5 Gene	rations			50 Gen	erations			200 Generations				
	Tr* Val*	Α	В	С	D	Α	В	С	D	Α	В	С	D		
(h ² =0.2)	A	0.702 (0.008)	0.304 (0.015)	0.313 (0.019)	0.261 (0.024)	0.669 (0.007)	0.076 (0.016)	0.107 (0.031)	0.076 (0.022)	0.596 (0.015)	-0.023 (0.019)	0.005 (0.026)	0.034 (0.021)		
it 1 (h	В	0.270 (0.022)	0.718 (0.011)	0.298 (0.023)	0.286 (0.023)	0.047 (0.026)	0.716 (0.011)	0.119 (0.024)	0.117 (0.022)	0.032 (0.019)	0.706 (0.008)	-0.017 (0.026)	-0.009 (0.021)		
Trait 1	С	0.305 (0.13)	0.276 (0.026)	0.713 (0.009)	0.299 (0.017)	0.087 (0.026)	0.105 (0.021)	0.689 (0.010)	0.103 (0.021)	0.025 (0.025)	0.016 (0.016)	0.703 (0.014)	0.010 (0.023)		
	D	0.298 (0.026)	0.356 (0.018)	0.283 (0.017)	0.705 (0.012)	0.074 (0.030)	0.093 (0.024)	0.091 (0.019)	0.709 (0.010)	0.030 (0.018)	-0.002 (0.023)	0.009 (0.019)	0.691 (0.017)		
	Tr Val	А	В	С	D	Α	В	С	D	Α	В	С	D		
(h ² =0.4)	A	0.808 (0.009)	0.403 (0.026)	0.413 (0.027)	0.379 (0.022)	0.812 (0.007)	0.072 (0.023)	0.145 (0.028)	0.138 (0.027)	0.789 (0.009)	-0.004 (0.025)	0.044 (0.020)	0.073 (0.026)		
Trait 2 (h	В	0.355 (0.014)	0.786 (0.005)	0.361 (0.018)	0.388 (0.013)	0.084 (0.011)	0.769 (0.011)	0.095 (0.017)	0.089 (0.018)	-0.005 (0.032)	0.704 (0.017)	0.012 (0.019)	-0.001 (0.018)		
Tra	С	0.406 (0.013)	0.428 (0.019)	0.822 (0.006)	0.429 (0.018)	0.126 (0.032)	0.074 (0.018)	0.816 (0.008)	0.128 (0.030)	-0.013 (0.027)	-0.009 (0.018)	0.813 (0.007)	-0.018 (0.018)		
	D	0.368 (0.016)	0.384 (0.033)	0.405 (0.020)	0.809 (0.007)	0.129 (0.021)	0.075 (0.025)	0.137 (0.017)	0.795 (0.008)	0.020 (0.021)	-0.004 (0.022)	0.031 (0.017)	0.806 (0.006)		

Table 2. Accuracy (standard error) of the predicted breeding values of the 4 purebred populations obtained from training in the same populations.

*Tr = Training sets, *Val = Validation sets

Table 3. Percentage (standard error) of gene-marker pairs in high LD in the validation sets found in the training sets, and correlation (standard error) between the LD of these pairs in the validation sets and the other purebred populations.

		50 Gene	erations		200 Generations							
Tr* Val*	Α	В	С	D	Α	В	С	D	Α	В	С	D
Α	100%	49% (0.849)	49.8% (0.923)	49% (1.071)	100%	11.5% (0.242)	11.9% (0.589)	12.3% (0.371)	100%	2.2% (0.234)	3.4% (0.175)	3.3% (0.207)
	1	0.917 (0.002)	0.926 (0.002)	0.923 (0.002)	1	0.348 (0.016)	0.375 (0.010)	0.381 (0.010)	1	0 (0.007)	0.020 (0.013)	0.025 (0.006)
В	48.8% (1.015)	100%	50.7% (1.032)	49.7% (1.070)	11.3% (0.473)	100%	12.4% (0.726)	13.1% (0.533)	2.6% (0.221)	100%	3.7% (0.312)	3.5% (0.241)
	0.917 (0.003)	1	0.923 (0.003)	0.925 (0.001)	0.353 (0.011)	1	0.376 (0.012)	0.386 (0.012)	-0.005 (0.007)	1	0.031 (0.012)	0.031 (0.007)
С	53.6% (0.746)	54.8% (0.567)	100%	55.1% (0.621)	13.4% (0.346)	14.2% (0.317)	100%	14.7% (0.676)	3% (0.192)	2.8% (0.329)	100%	3.9% (0.362)
	0.932 (0.002)	0.932 (0.002)	1	0.940 (0.001)	0.390 (0.008)	0.399 (0.009)	1	0.416 (0.007)	0.011 (0.010)	0.020 (0.006)	1	0.036 (0.008)
D	53.5% (0.806)	54.5% (0.700)	55.9% (0.854)	100%	13.6% (0.279)	14.9% (0.612)	14.3% (0.699)	100%	3% (0.136)	2.7% (0.199)	4% (0.294)	100%
	0.931 (0.002)	0.931 (0.001)	0.942 (0.002)	1	0.400 (0.010)	0.404 (0.010)	0.423 (0.009)	1	0.029 (0.008)	0.026 (0.004)	0.037 (0.007)	1

*Tr = Training sets, *Val = Validation sets

Admixed Populations x2

The use of admixed populations for genomic evaluation was first proposed by de Roos et al. (2009) with the purpose of overcoming the obstacles that genomic selection is facing when to its implementation in the beef cattle industry. These obstacles are the large number of populations and their reduced population size, due to which the withinbreed evaluation is expected to give very low accuracies. Moreover, the estimations made in one population do not produce accurate predictions for a different population (Harris et al., 2008). The use of phenotypic and genotypic data from different populations permits the construction of a reference population large enough as to increase the accuracy of the predictions that can be applied in all the populations involved. Moreover, the larger variability of this population may contribute further in the increase of the accuracy (Falconer and MacKay, 1996).

In this study, the admixed populations were simulated combining data of individuals from the 4 purebred populations. For this purpose, 1,000 individuals were chosen randomly from each one of the 4 purebred populations. Afterwards, these data were combined in pairs of two creating 6 admixed populations (A+B, A+C, A+D, B+C, B+D and C+D) of 2,000 individuals each with 50% from each purebred. These admixed datasets were used to predict the breeding values of the 4 purebred populations.

Table 4, structured as before, shows the accuracies of the predicted breeding values of the validation sets obtained from training in 4 admixed populations (A+B, A+C, B+D and C+D). The results from the A+D and B+C populations are not shown as they are similar to the A+C and B+D, respectively.

The highest accuracies for each trait were obtained for the populations that were included in the admixed training sets and when the populations were separated just for 5 generations. These values ranged around 0.35 and 0.5 for traits 1 and 2, respectively. As the number of generations increased, the accuracy of the predictions for the populations under selection declined. After 50 and 200 generations, Population A lost 11% (0.313 and 0.300) and 28% of accuracy (0.257 and 0.253) for trait 1 and population B lost 8% (0.433 and 0.441) and 27% (0.357 and 0.353) of accuracy for trait 2. Meanwhile, in the cases where selection had not acted the accuracies were maintained almost at the same levels.

The predictive ability of the admixed populations over the ones that did not participate in them was extremely low even when the populations are separated just for 5 generations. The values obtained were between 0.11 and 0.17 for the first trait and between 0.17 and 0.24 for the second trait and declined rapidly to 0 after 50 generations of divergence.

Regarding the linkage disequilibrium (Table 5), the admixed populations presented 65 to 70% of the gene-marker pairs in high LD found in the validations sets that were included in the mixture. The correlation of the LD of these pairs between the validation sets and the admixed populations was 0.98. This percentage declined to around 31% when the populations were separated for 50 generations, and to around 25% in the case of 200 generations. Nevertheless, the correlations between the LD of the pairs of markers did not decline in the same rate. They were maintained above 0.9 in the case of 50 generations and dropped at 0.85 after 200 generations. Although the percentage of common gene-marker pairs in high LD was reduced with the generations of divergence, the accuracies remained at the same levels showing a greater resemblance to the evolution of the correlations between the LD of the same pairs in the different populations throughout the generations.

When comparing the LD between the admixed and the purebred populations that were not included in the mixture, it can be noticed that the admixed populations present 49% to 56% of the pairs found in the purebred populations when they are separated for 5 generations and this percentage is reduced to 11-15% and 7-11% after 50 and 200 generations respectively. Similarly, the correlation between the LD that these pairs present in every population, starts from 0.94 and declines rapidly to 0.5 and finally to 0 with the generations. Though the admixed populations share the 50% of the pairs in high LD with the purebred populations that are not included in the mixture, and the correlation between the LD of these pairs is extremely high, the accuracies obtained are very low.

Table 4. Accuracy (standard error) of the predicted breeding values for the 4 purebred populations obtained from training in 4 admixed populations (A+B, A+C, B+D and C+D).

			5 Gene	rations			50 Gene	erations			200 Gen	erations	
	Tr* Val*	A+B	A+C	B+D	C+D	A+B	A+C	B+D	C+D	A+B	A+C	B+D	C+D
(h ² =0.2)	Α	0.360 (0.007)	0.335 (0.009)	0.107 (0.018)	0.128 (0.020)	0.313 (0.011)	0.300 (0.015)	0.015 (0.023)	0.074 (0.018)	0.257 (0.012)	0.253 (0.006)	-0.010 (0.018)	0.025 (0.013)
Trait 1 (h	В	0.370 (0.011)	0.113 (0.016)	0.352 (0.008)	0.131 (0.025)	0.331 (0.010)	0.008 (0.031)	0.350 (0.009)	0.072 (0.015)	0.332 (0.010)	0.001 (0.024)	0.335 (0.016)	-0.016 (0.021)
Tra	С	0.125 (0.023)	0.353 (0.009)	0.110 (0.020)	0.348 (0.012)	-0.014 (0.019)	0.323 (0.010)	0.010 (0.022)	0.344 (0.011)	0.015 (0.025)	0.344 (0.011)	0.032 (0.017)	0.352 (0.013)
	D	0.157 (0.028)	0.140 (0.021)	0.353 (0.009)	0.348 (0.009)	0.039 (0.018)	0 (0.022)	0.332 (0.012)	0.340 (0.011)	0.017 (0.023)	0.003 (0.020)	0.349 (0.011)	0.324 (0.019)
_	Tr Val	A+B	A+C	B+D	C+D	A+B	A+C	B+D	C+D	A+B	A+C	B+D	C+D
² =0.4)	Α	0.520 (0.015)	0.502 (0.013)	0.193 (0.022)	0.225 (0.028)	0.504 (0.010)	0.499 (0.016)	0.068 (0.022)	0.076 (0.020)	0.473 (0.011)	0.465 (0.012)	-0.002 (0.027)	-0.002 (0.017)
Trait 2 (h ² =0.4)	В	0.481 (0.010)	0.209 (0.020)	0.477 (0.011)	0.178 (0.019)	0.443 (0.017)	0.047 (0.019)	0.441 (0.014)	0.051 (0.020)	0.357 (0.017)	-0.007 (0.017)	0.353 (0.019)	0.034 (0.016)
Trai	С	0.228 (0.015)	0.518 (0.014)	0.245 (0.017)	0.526 (0.014)	0.048 (0.024)	0.513 (0.014)	0.025 (0.014)	0.520 (0.010)	-0.009 (0.021)	0.504 (0.008)	-0.022 (0.020)	0.514 (0.008)
	D	0.167 (0.019)	0.189 (0.020)	0.521 (0.007)	0.517 (0.011)	0.009 (0.021)	0.054 (0.028)	0.506 (0.008)	0.487 (0.007)	-0.009 (0.028)	0.022 (0.021)	0.498 (0.017)	0.494 (0.015)

*Tr = Training sets, *Val = Validation sets

Table 5. Percentage (standard error) of gene-marker pairs in high LD in the validation sets found in the training sets and correlation (standard error) between the LD of these pairs in the validation sets and the admixed populations.

5 Generations						50 Generations				200 Generations			
Tr* Val*	A+B	A+C	B+D	C+D	A+E	B A+C	B+D	C+D	A+B	A+C	B+D	C+D	
Α	65.6% (0.869) 0.982	65.6% (0.627) 0.984	49.8% (0.913) 0.937	49.5% (1.039) 0.939	31.89 (0.55) 0.90	3) (0.828)	11.7% (0.476) 0.436	11.2% (0.630) 0.447	27% (0.764) 0.842	24.9% (0.405) 0.863	8.2% (0.347) 0.035	6.7% (0.257) 0.033	
В	(0.001) 65.6% (0.905) 0.982	(0.001) 50.6% (0.966) 0.936	(0.002) 66.3% (0.731) 0.984	(0.002) 50.7% (1.027) 0.938	(0.00) 31.69 (0.36) 0.90	3)(0.002)%11.6%8)(0.690)	(0.009) 31.3% (0.954) 0.910	(0.011) 11.7% (0.648) 0.453	(0.004) 26.1% (0.561) 0.838	(0.004) 8% (0.353) 0.027	(0.008) 25.1% (0.796) 0.857	(0.010) 6.4% (0.346) 0.038	
С	(0.001) 55.6% (0.666) 0.949	(0.002) 68.8% (0.714) 0.985	(0.001) 55.8% (0.496) 0.952	(0.002) 70% (0.536) 0.987	(0.00) (0.00) (0.54) (0.54) (0.48)	2)(0.013)%31.7%0)(0.617)	(0.002) 14.8% (0.502) 0.492	(0.010) 32.3% (0.655) 0.919		(0.009) 26.1% (0.390) 0.844	(0.003) 9% (0.344) 0.040	(0.011) 25.3% (0.672) 0.866	
D	(0.001) 55.8% (0.781) 0.949	(0.001) 55.4% (0.807) 0.952	(0.001) 69.4% (0.484) 0.984	(0.001) 70.2% (0.684) 0.987	(0.010 15.89 (0.58) 0.499	0) (0.002) % 13.8% 3) (0.564)	(0.007) 33.4% (1.028) 0.913	(0.002) 31.2% (0.804) 0.918	(0.008) (0.008) (0.437) (0.437) 0.058	(0.003) 8.6% (0.311) 0.046	(0.009) 27.7% (0.538) 0.850	(0.002) 24.3% (0.496) 0.861	
	(0.001)	(0.001)	(0.001)	(0.001)	(0.00)	8) (0.009)	(0.001)	(0.002)	(0.007)	(0.006)	(0.002)	(0.003)	

*Tr = Training sets, *Val = Validation sets

Admixed Population x4

Besides the admixed populations composed by two purebred, one additional admixed population was created. In this mixture, all 4 purebred populations contributed with data from 500 individuals each, chosen randomly. Table 6 shows the accuracies of the estimated breeding values when the A+B+C+D population was used for training.

The highest values were obtained in the case of 5 generations of separation and ranged around 0.52 for all the validation sets for trait 1 and around 0.60 for trait 2. The values for trait 1 declined approximately 20% for all validation sets after 50 generations and an additional 25% after 200 generations. For the first trait, the highest reduction of accuracy (36%) was presented in population A, which was selected for this characteristic. The loss of accuracy for the second trait was about 30% after 50 generations. The accuracy for the population B presented an additional reduction of 20% after 200 generations due to the fact that this population was selected for this trait. Yet, the reduction of the accuracies for the rest of the populations was much smaller ranging from 1% (for populations C and D) to 10% (for population A).

The admixed population shares similar percentages of gene-marker pairs in high LD with all the purebred populations (Table 7) because all purebred contributed with the same proportion in the mixture. In the case of 5 generations of divergence the admixed population has around 62% of the pairs in high LD found in all 4 purebred populations. However, it should be noted this figure is slight lower (56.8 and 57.6) for populations A and B. This percentage decreases rapidly to 20% and 9% after 50 and 200 generations respectively. As before, it should be noted that populations A and B presented smaller values (16 and 8%). As in the admixed populations composed by two purebred, the correlation of the LD of these pairs in the admixed and the purebred populations does not decline at the same ratio as the percentages. It starts at 0.97 in the first case, 0.82 in the second and finally 0.69 after 200 generations.

		5 Generations	50 Generations	200 Generations
	Tr* Val*	A+B+C+D	A+B+C+D	A+B+C+D
=0.2	Α	0.510	0.386	0.245
Trait 1 (h ² =0.2)	В	(0.011) 0.517 (0.018)	(0.018) 0.444 (0.031)	(0.014) 0.326 (0.028)
Trai	С	0.532 (0.015)	0.416 (0.021)	0.294 (0.032)
	D	0.523 (0.015)	0.421 (0.033)	0.305 (0.033)
-	Tr Val	A+B+C+D	A+B+C+D	A+B+C+D
Trait 2 (h ² =0.4)	Α	0.598 (0.028)	0.428 (0.015)	0.383 (0.013)
it 2 (ŀ	В	0.535 (0.033)	0.358 (0.016)	0.286 (0.019)
Tra	С	0.612 (0.033)	0.427 (0.016)	0.408 (0.017)
	D	0.587 (0.034)	0.408 (0.009)	0.405 (0.013)

Table 6. Accuracy (standard error) of the predicted breeding values for the 4 purebred populations obtained from training in the admixed population (A+B+C+D).

Table 7. Percentage (standard error) of the gene-marker pairs in high LD in the
validation sets found in the training sets, and correlation (standard error) between the
LD of these pairs in the validation sets and the admixed population.

	5 Generations	50 Generations	200 Generations
Tr Val	A+B+C+D	A+B+C+D	A+B+C+D
A	56.8%	16.3%	7.9%
2	(0.875)	(0.598)	(0.172)
	0.971	0.814	0.690
	(0.001)	(0.003)	(0.005)
В	57.6%	16.8%	7.8%
D	(0.958)	(0.753)	(0.279)
	0.970	0.813	0.688
	(0.001)	(0.005)	(0.010)
С	62.7%	20%	9.1%
C	(0.588)	(0.434)	(0.360)
	0.977	0.824	0.662
	(0.001)	(0.004)	(0.009)
D	62.9%	20.2%	9.6%
2	(0.761)	(0.763)	(0.305)
	0.978	0.831	0.672
	(0.001)	(0.002)	(0.007)

From the results presented above, it can be seen that the purebred populations performed better under all scenarios, as previously reported by Kizilkaya et al. (2010) and Toosi et al. (2010). Though, the admixed populations presented some results that were not so clearly expected. The population A+B+C+D gave similar accuracies to the ones obtained from the other admixed x2 populations for the purebred populations involved and, in some cases, even higher. Yet, the LD extent that this population presents does not match with the pattern of the results of accuracy.

One example of this discrepancy is presented with more detail in Table 8. It shows the accuracies of the predicted breeding values for the first trait for population A obtained from training in this purebred population and two admixed populations, A+B and A+B+C+D, under the scenario of 50 generations of divergence. The percentage of genemarker pairs in high LD that each training set shares with the population A, together with the correlation between the LD values that these pairs present in the validation set and the training sets (between brackets) are shown beneath the correspondent accuracies. The population A+B maintains the 31.8% of the gene-marker pairs in high LD found in population A and yields an accuracy of 0.313. On the contrary, population A+B+C+D gave an accuracy of 0.386 while presenting just 16.3% of these pairs. Thus, the accuracies obtained from these two training sets are apparently not consistent with the results from the LD.

Table 8. Accuracy of predicted breeding values and percentage (correlation) of genemarker pairs in high LD for population A when training in populations A, A+B and A+B+C+D.

		Α	A+B	A+B+C+D
Δ	Accuracy	0.669	0.313	0.386
1	LD (corr.)	100% (1)	31.8% (0.905)	16.3% (0.814)

To investigate the causality of this phenomenon, the populations from the example above were taken and examined thoroughly. Firstly, the estimates of the marker effects obtained from these populations were represented graphically (figure 8). The magnitude of the estimated markers effects from population A+B+C+D ranged between -0.10 and +0.10 units and they were more similar to the estimates from the population A, which ranged from -0.03 to 0.03. On the contrary, the estimations from population A+B were

extremely high, ranging between -2 and +2. Similar results were observed in other replicates for the admixed x2 populations.

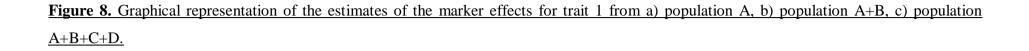
Afterwards, in order to examine the extent of LD in all these populations, the correlation coefficient (r) between pairs of loci was calculated. The average r of two consecutive loci, of loci that are located in the same chromosome, and of loci that are located in different chromosomes was calculated for all three populations and is presented in the Table 9. The average r of two consecutive loci in the population A is 0.253 and in population A+B is 0.247 because the change of the allelic frequencies caused by the mixture of the populations tends to break the LD between loci. When more populations are included in the mixture, the degradation of the LD is stronger as in the case of the A+B+C+D population in which the average r is 0.234. The average r declines with the distance between the loci and therefore in population A the average LD of loci within the same chromosome is 0.084 and between chromosomes is 0.045. In population A+B the average LD of loci within the same chromosome is 0.105, 25% more than in population A, and between chromosomes is 0.089, 97% more. The increase of the values of LD of distant loci found in the A+B population occurs because, in the populations A and B, selection and genetic drift forced the frequencies of the loci to extremes. When data from these populations were mixed, the causative mutations and the markers that were fixed oppositely in the two populations appeared with intermediate frequencies that created a high LD over large distances. The A+B+C+D population presents an average LD within-chromosome just 6% (0.089) higher than in population A and between-chromosomes 58% (0.071) higher. The changes of the allelic frequencies in this population also cause the creation of LD between distant loci, but with lower intensity because when mixing 4 populations, it is less probable to encounter loci with extreme opposite frequencies.

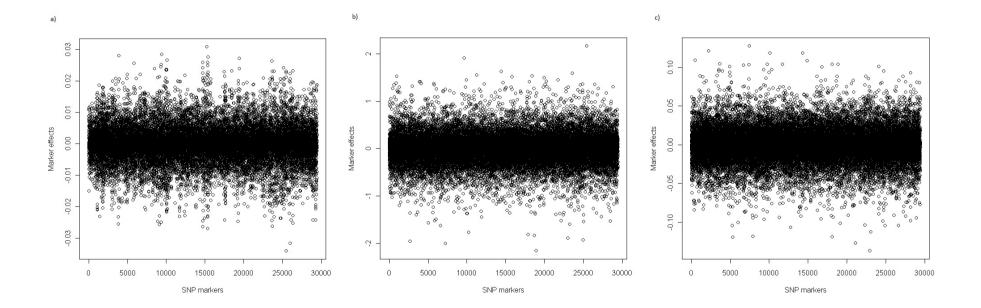
At the moment of estimating the effects of the markers, those markers that are in LD with a causative mutation can capture a part of its effect depending on the value of the LD between them. In the case of the A+B population, the fact that many markers are in strong LD with many distant causative mutations, results in these markers capturing effects from those distant loci and therefore, their effects are clearly overestimated, as it can be seen in the figure 8. The breeding values produced with these estimates are less accurate than the ones that are produced by the estimates from the A+B+C+D population because in this population fewer markers present high LD with distant loci.

Figure 9 shows the example of the marker 11437 which is located on the 8th chromosome. This marker was estimated to have a very small effect in the populations A and A+B+C+D but one of the largest in the population A+B. When considering the LD that this marker presents with all the causative mutations, it can be seen that it is very high even with distant loci in the A+B population. Instead, the same marker in the other populations has lower values of LD and with loci in shorter distances.

Table 9. The average LD (r) between markers and causative mutations when they are consecutive, located in the same (within) and different (between) chromosome, in the populations A, A+B and A+B+C+D.

	Α	A+B	A+B+C+D
Average r of consecutive loci	0.253	0.247	0.234
Average r within chromosome	0.084	0.105	0.089
Average r between chromosomes	0.045	0.089	0.071





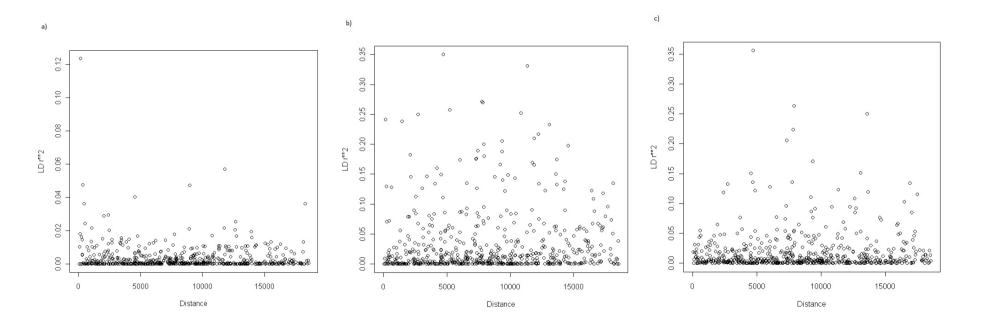


Figure 9. The LD between the marker 11437 and the causative mutations affecting the first trait represented over the distance between them in a) population A, b) population A+B and c) population A+B+C+D.

Marker density

In addition, a second scenario was simulated with a higher density of markers. For this purpose, 30,000 biallelic markers were distributed evenly along only 4 chromosomes of 1 Morgan each, resulting in a density of 75 markers per cM. The scenarios of the number of causative mutations that control the traits and the type and number of generations of divergence were the same as in the first case. The results of the accuracy of the predictions obtained from the purebred population A and the admixed populations A+B and A+B+C+D are presented in Table 10. For simplicity, the results from the other populations are not shown since they are equivalent to those presented.

The accuracies of the purebred populations, when training and validating in the same population, increased around 11% and 7% over the previous case for the traits 1 and 2, respectively. This increase is consistent under all scenarios of number of generations of divergence. The predictive ability of the purebred populations over the others was improved significantly (0.51 and 0.55 for traits 1 and 2, respectively) compared to the case of lower marker density (0.3 and 0.37 for traits 1 and 2, respectively) under the scenario of 5 generations of divergence. Nevertheless, these accuracies declined rapidly to 0.15 (trait 1) and 0.20 (trait 2) and close to 0 with 50 and 200 generations of divergence, respectively.

The higher marker density favored most the predictive ability of the admixed population A+B for the populations that were included in the mixture. The accuracies obtained were increased around 90% and 50% for traits 1 and 2, respectively. The accuracies obtained for the populations that are not included in the admixed population increased as well, being 0.49 for the first trait and 0.56 for the second, with 5 generations of divergence. These accuracies declined with the generations of divergence similarly to the ones obtained from the purebred when predicting another population.

Finally, the accuracies obtained from the A+B+C+D population, comparing with the previous scenario, were increased only 25% for both traits when the populations diverged for only 5 generations but showed a remarkable increase when the populations had diverged for more generations. In the case of 50 generations the increase was of 35% and 70% for traits 1 and 2, respectively and in the case of 200 generations up to 90% and 75%.

The results from the linkage disequilibrium study (Table 11) showed that the persistence of the LD across the populations is almost identical in both cases of marker density, with only a slight increase in the case of the higher marker density (i.e. The admixed population A+B maintains 65.6% and 66.7% of the gene-marker pairs in high LD found in population A in the case of 300 and 100 QTLs, respectively.)

As it can be noticed, the admixed population A+B performed better than the A+B+C+D population in this case, even though the LD patterns are the same as before. This happens due to the fact that the simulated genome consist of only 4 chromosomes and therefore, due to the effect of linkage, the frequencies of the loci in the populations A and B were not allowed to have extreme opposite frequencies by chance.

Overall, the results of the comparison between both scenarios agree with the studies of Toosi et al. (2009) and de Roos et al. (2009), and show that the increased marker density gives higher accuracies when the populations have diverged for only a few generations, benefits more the predictions made in the admixed populations than the ones made in the purebreds and finally, favors the accuracies for traits with low heritability.

Number of QTLs

Further, another scenario with only 100 causative mutations or QTL affecting each trait was also simulated and its predictive ability for all populations was compared with the scenario of 300 genes affecting the traits. The marker density was the same as in the first case. The results (not shown) of the accuracies show that the smaller number of genes affected only the predictive ability of the populations under selection and the admixed populations that contained data from these selected populations. We were not able to found any difference in the accuracy between the 100 and 300 QTL scenarios when the populations are separated for 5 generations, but the accuracy of the selected populations declines faster with the generations of divergence in the 100 genes scenario. Population A showed a decline in the accuracy for the trait 1 of 25% after 200 generations of selection whereas the same reduction in the case of 300 genes was only 15%. Likewise, the accuracy of population B for the trait 2 declined 23% in the case of 100 genes and only 10% in the case of 300 genes. The admixed populations that contain data from the purebred populations under selection showed as well a larger reduction in their accuracies with the generations of evolution in the case of 100 genes. This occurs

because the alleles of a small number of genes are fixed or lost easier during the process of selection and the loss of genetic variance is faster.

However, the study of Kizilkaya et al. (2010) showed that the accuracies of the predicted values depend on the number of genes that control the trait of interest. The accuracies he obtained from training in an admixed population started at 0.388 when there were 50 QTLs and declined to 0.2 when the number of QTLs was 500. It should be noted that this authors compare scenarios with 10 times more QTL whereas in this study the spectrum of scenarios was only a ratio of 3 in the number of genes. Thus, it is probable that the differences between them were too small to be detected with only 10 replicates of simulation.

The findings of this study suggest a promising alternative for the beef cattle industry. There are some important advantages in using an admixed population that consists of data from all available populations. Firstly, the size of the reference population can be easily increased just by adding a few individuals from each of the many existing populations. Secondly, the large number of populations of different genetic origin would lower further the intensity of the phenomenon of overestimation of the marker effects, and finally, it will not be necessary to have different procedures of prediction for every populations. This strategy, combined with the use of a high density marker map could result in accurate predictions suitable for the implementation of genomic selection in beef cattle.

Table 10. Accuracy (standard error) of the predicted breeding values of all validation sets obtained from training in the purebred population Aand in the admixed populations A+B and A+B+C+D.

			5 Genera	ations		50 Gener	ations	200 Generations		
	Tr* Val*	Α	A+B	A+B+C+D	Α	A+B	A+B+C+D	Α	A+B	A+B+C+D
(h ² =0.2)	Α	0.782 (0.007)	0.671 (0.011)	0.645 (0.015)	0.769 (0.013)	0.611 (0.017)	0.566 (0.018)	0.714 (0.012)	0.510 (0.036)	0.467 (0.025)
Trait 1 (l	В	0.517 (0.017)	0.688 (0.009)	0.674 (0.015)	0.156 (0.023)	0.660 (0.012)	0.604 (0.018)	0.024 (0.030)	0.619 (0.018)	0.542 (0.024)
Tra	С	0.517 (0.019)	0.489 (0.021)	0.645 (0.023)	0.171 (0.029)	0.163 (0.046)	0.570 (0.025)	-0.014 (0.023)	0.018 (0.028)	0.560 (0.014)
	D	0.498 (0.022)	0.471 (0.017)	0.640 (0.017)	0.140 (0.031)	0.137 (0.025)	0.541 (0.018)	0.004 (0.013)	0.017 (0.022)	0.585 (0.019)
	Tr Val	Α	A+B	A+B+C+D	Α	A+B	A+B+C+D	Α	A+B	A+B+C+D
(h ² =0.4)	Α	0.871 (0.005)	0.775 (0.008)	0.750 (0.011)	0.871 (0.006)	0.766 (0.008)	0.724 (0.009)	0.875 (0.005)	0.720 (0.026)	0.669 (0.014)
Trait 2 (l	В	0.554 (0.014)	0.746 (0.008)	0.746 (0.008)	0.142 (0.021)	0.695 (0.018)	0.628 (0.016)	0.025 (0.032)	0.535 (0.026)	0.536 (0.020)
Tra	С	0.556 (0.023)	0.560 (0.024)	0.763 (0.011)	0.266 (0.019)	0.226 (0.031)	0.704 (0.014)	0.023 (0.031)	0.047 (0.018)	0.687 (0.014)
	D	0.587 (0.015)	0.561 (0.013)	0.780 (0.007)	0.251 (0.027)	0.210 (0.032)	0.686 (0.015)	0.042 (0.024)	0.022 (0.030)	0.709 (0.016)

<u>**Table 11.** Percentage (standard error) of gene-marker pairs in high LD in the validation sets found in the training sets, and correlation (standard error) between the LD of these pairs in the validation sets and the purebred population A and the admixed populations A+B and A+B+C+D.</u>

	5	5 Generat	tions	5	0 Genera	tions	20	200 Generations			
Tr* Val*	Α	A+B	A+B+C+D	Α	A+B	A+B+C+D	Α	A+B	A+B+C+D		
A	100%	66.7% (0.746)	58.8% (1.317)	100%	32.5% (0.469)	17.3% (0.465)	100%	27.2% (0.526)	8.6% (0.248)		
	1	0.983 (0.001)	0.973 (0.001)	1	0.904 (0.002)	0.821 (0.003)	1	0.834 (0.003)	0.668 (0.006)		
В	49.9%	65.3%	57.8%	11.9%	31.6%	17.5%	2.8%	27%	8.6%		
	(0.952) 0.917	(0.902) 0.982	(0.920) 0.972	(0.346) 0.336	(0.510) 0.905	(0.628) 0.820	(0.228) 0.020	(0.514) 0.841	(0.213) 0.694		
	(0.004)	(0.001)	(0.001)	(0.009)	(0.003)	(0.005)	(0.007)	(0.004)	(0.008)		
С	55.8% (1.053)	57.4% (1.039)	64% (0.863)	13.9% (0.444)	16.3% (0.620)	20.7% (0.556)	3.2% (0.119)	10.8% (0.483)	9.5% (0.277)		
	0.935 (0.003)	0.952 (0.002)	0.977 (0.001)	0.396 (0.008)	0.485 (0.012)	0.823 (0.003)	0.024 (0.003)	0.054 (0.007)	0.659 (0.007)		
D	56.1% (0.738)	57.2% (0.845)	64.1% (0.986)	14.4% (0.628)	16% (0.456)	20.3% (0.585)	3.1% (0.164)	10.8%	9.9% (0.323)		
	0.936 (0.003)	0.953 (0.003)	0.978 (0.001)	0.400 (0.009)	0.481 (0.007)	0.811 (0.003)	0.015 (0.005)	0.049 (0.006)	0.667 (0.006)		

CONCLUSIONS

The conclusions of this study are:

- 1. Genomic prediction when training and validating in the same population gives high accuracies. However, these accuracies decline with the number of generations when the population is undergoing a process of selection for the trait of interest.
- 2. The predictive ability of one population over another is low and declines further with the number of generations of divergence.
- 3. The predicted ability of the admixed populations is better for the populations that are included in the mixture than for the ones that are not included. Nevertheless, these accuracies are remarkably lower than the ones obtained from the purebred populations.
- 4. When all populations available are represented in the admixed training set the accuracies increase for all populations compared to the accuracies obtained from the admixed x2 populations.
- 5. A high marker density genotyping benefits the accuracies of all populations but mostly the admixed populations and the predictions for traits with low heritability.
- 6. The number of QTLs that control the traits affects mainly the predictive ability of the populations that are under selection for these traits. The smaller the number of QTLs the faster the reduction in the accuracy with the number of generations of selection.
- 7. A genomic prediction procedure from an admixed population generated by several related genetic origins and with a high density map of markers may be useful for selection of all populations involved.

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APPENDIX: SIMULATION PROGRAM

program simulation

implicit none

integer *1
mar(100,20,1500,2), marn(100,20,1500,2), marnA(100,20,1500,2), marnB(100,20,1500,2), marnC(100,20,1500,2), marnD(100,20,1500,2), grupoA(100,20,1500,2), finalAC(2000,20,1500,2), finalAC(200,20,1500,2), finalAC(2000,20,1500,2), finalAC(200,20,1500,2), f

integer

irec(50), irecord(50), iefec1(300,2), iefec2(300,2), ef(600), n(30000), i, j, k, l, ijk, iale, ip, iv, im, ic, ik, ij, pois, neu, g, marca(30000), nanim, ncro, nmar, nale, nmac, nhem, ngen, nmut, nrec, xmin, imin, imut, nanim2, ngenes, ngen2, ai real *8

namin2, ngenes, ngenz, ait real *8 x1, u, lam, yfenA(100), yfenB(100), var1(300), var2(300), media1, media2, p1base(300,300), p2base(300,300), p1(300,300), p2(300,300), Vir1, Vr2, Va1, Va2, des1, des2, varA1(300), var2AB(300), varA2(300), var1B(300), var2B(300), var1Df(300), var1B(300,300), g2se(1300,300), efec1(300), efec2(300), var1B(300), var2B(530), var1Cf(300), var2Cf(300), var1AB(300), var2AB(300), varA72(300), var1AC(300), var2AB(300), var1AB(300), var2BG(300,300), g2se(1300, var1Bf(300), var2BB(300), var1Cf(300), var2AB(300), varA2(300,300), p1B(300,300), p1C(300,300), p1C(300,300), p2C(300,300), p2D(200,300), p1ABCD(300,300), p2AC(300,300), p2AC(300,300), p1A(300,300), p2D(300,300), p1ABCD(300,300), p2AC(300,300), p2AC(300,300), p1A(300,300), p2D(200,300), p1ABCD(300,300), p2ABCD(300,300), p2AC(300,300), p1AD(300,300), p2D(200,300), p1CD(300,300), p2C(300,300), p1ABCD(300,300), p2ABCD(300,300), p2AC(300,300), p1AD(300,300), p2AD(200,300), p1ABCD(300,300), p2ABCD(300,300), p2AC(300,300), p1AD(300,300), p2AD(200,300), p1ABCD(300,300), p2AAC(300,300), p2AD(200,300), p2AD(200,300), p1ABCD(300,300), p2AC(300,300), p1AD(300,300), p2AD(200,300), p1CD(300,300), p2AD(200,300), p1ABCD(300,300), p2AC(300,300), p1ABC(300,300), p1ABCD(300,300), p2AD(200,300), p2AD(200,300), p2AD(200,300), p1ABCD(300,300), p2AC(300,300), p1ABCD(300,300), p2AD(200,300), p1ABCD(300,300), p2ABCD(300,300), p1ABCD(300,300), p1ABCD(300,300), p1ABCD(300,300), p2AD(200,300), p1ABCD(300,300), p2ABCD(300), yfen1ABCD(2000), yfen1ABCD(2000), yfen2AAC(2000), yfen1ABCD(2000), yfen2AAC(2000), yfen2AAC(2000), yfen2AAC(2000), yfen1ABCD(2000), yfen2ABCD(2000), yfen2ABCD(2000), yfen2ABCD(2000), yfen2ABCD(2000), yfen2BCD(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2ABCD(2000), yfen2ABCD(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2BBC(2000), yfen2ABCD(2000), yfen2AB ngen=1000 nanim2=2000 ngenes=300 ngen2=200 <u>! simulation of the base population</u> do i=1,nanim do j=1,ncro do k=1,nmar do l=1 nale to l=1,nale call unif(x1,u) if (u.lt.0.5) then mar(i,j,k,l)=1 else mar(i,j,k,l)=2 end if enddo enddo enddo enddo ! simulation of 1000 generations of random mating do ijk=1,ngen do i=1,nanim call unif(x1,u) ip=int(u*nmac)+1 call unif(x1,u) im=int(u*nhem)+n <u>! father's chromosome</u> do j=1,ncro call unif(x1,u) c+1if (u.lt.0.5) th iale=1 else iale=2 endif marn(i,j,1,1)=mar(ip,j,1,iale)

 Iam-1.
 rec=pois(lam,x1)
 <u>! selection of the number of recombinations</u> do ik=1,rec call unif(x1,u)

 reckipicities
 <u>! selection of the location of the recombinations</u> irreckipi-int(u*(nmar-1))+1

 enddo irecord=0 do ik=1,nrec xmin=999999 do ij=1,nrec if (irec(ij).lt.xmin) then xmin=irec(ij) imin=ij endif enddo irecord(ik)=xmin irec(imin)=999999 enddo irecord(nrec+1)=nmar if (nrec.eq.0) then do k=2,nmar marn(i,jo,k,1)=mar(ip,j,k,iale) enddo else ic=2 do ik=1,nrec+1 do k=ic,irecord(ik) do k=ic,irecord(ik) marn(i,j,k,1)=mar(ip,j,k,iale) enddo ic=irecord(ik)+1 if (iale.eq.1) ther iale=2 else iale=1 endif enddo endif ! mutation lam=1. nmut=pois(lam,x1) do ik=1,nmut call unif(x1,u) imut=int(u*nmar)+1 if (marn(i,jinmt,1)-e,1) then marn(i,jinmt,1)=2 else marn(i,jinmt,1)=1 ! selection of the number of mutations ! selection of the location of the mutation marn(i,j,imut,1)=1 enddo enddo ! mother's chromosome do j=1,ncro call unif(x1,u) if (u.lt.0.5) then

iale=1 else iale=2 endif marn(i,i,1,2)=mar(im,i,1,iale) lam=1 lam=1. nrec=pois(lam,x1) do ik=1,nrec ! selection of the number of recombinations call unif(x1,u) irec(ik)=int(u*(nmar-1))+1 enddo ! selection of the location of the recombinations /ec(ik)=Ins_ nddo irecord=0 do ik=1.nrcc xmin=999999 do ij=1.nrcc if (irec(ij).lt.xmin) then xmin=irec(ij) imin=ij endif enddo irecord(ik)=xmin irec(imin)=999999 enddo irecord(nrec+1)=nmar if (nrec.eq.0) then do k=2,nmar marn(i,j,k,2)=mar(im,j,k,iale) enddo else ic=2 ic=2 do ik=1,nrec+1 do k=ic,irecord(ik) marn(i,j,k,2)=mar(im,j,k,iale) enddo ic=irecord(ik)+1 if (iale.eq.1) then iale=2 else iale=1 endif enddo endif ! mutation lam=1. nmut=pois(lam,x1) do ik=1,nmut call unif(x1,u) ! selection of the number of mutations ! selection of the location of the mutations can untr(x1,u) imut=int(u*nmar)+1 if (marn(i,j,imut,2).eq.1) then marn(i,j,imut,2)=2 else else marn(i,j,imut,2)=1 endif enddo enddo enddo mar=marn print *, ijk enddo open(11,file='info.txt') Trait 1, h²=0.2, mean=100 <u>! simulation of the causative mutations and their effect for trait 1</u> open(16,file='effects1.txt') do i=1,ngenes call unif(x1,u) ! selection of chromosome $\begin{array}{l} \mbox{initial}(x1,u) \\ \mbox{iefec1}(i,1) = \mbox{int}(u^*\mbox{ncro}) + 1 \\ \mbox{call unif}(x1,u) \\ \mbox{iefec1}(i,2) = \mbox{int}(u^*\mbox{nmar}) + 1 \end{array}$! selection of marker call normal(x1,u) efec1(i)=u*1 ! assignment of the effect $\begin{aligned} & \text{enert}(1) = a^{-1} \\ & \text{write}(16^{+}) i.((\text{icfect}(i,1)-1)^{*}\text{nmar+iefecl}(i,2)), \text{efecl}(i) \\ & \text{end } do \\ & \text{close}(16) \\ & \frac{lfrequencies}{16} \\ & \frac{lfrequencies}{16} \\ & \text{of } = 1, \text{ngenes} \\ & \text{plbase}(\text{iefecl}(i,1), \text{iefecl}(i,2)) = 0 \\ & \text{do } \mid = 1, \text{namin} \\ & \text{plbase}(\text{iefecl}(i,1), \text{iefecl}(i,2)) = 0 \\ & \text{do } \mid = 1, \text{namin} \\ & \text{plbase}(\text{iefecl}(i,1), \text{iefecl}(i,2)) = 0 \\ & \text{do } \mid = 1, \text{namin} \\ & \text{plbase}(\text{iefecl}(i,1), \text{iefecl}(i,2)) = 0 \\ & \text{base}(\text{iefecl}(i,1), \text{iefecl}(i,2)) = 0 \\ & \text{base}(\text{ie$ write(16,*) i,((iefec1(i,1)-1)*nmar+iefec1(i,2)),efec1(i) <u>! additive variance of trait 1</u> Va1=0 do i=1,ngenes $_{var1(i)=2^{v}plase(iefec1(i,1),iefec1(i,2))*(1-plbase(iefec1(i,1),iefec1(i,2)))*efec1(i)**2 Val=Val+var1(i)$ Val=Val+var1(t) end do print *, 'Val base =',Val write(11,*) 'Val base=',Val <u>! resistual variance of trait 1</u> Vr1=(Val/0.2)-Val print *, 'Vr1 =',Vr1 write(11,*) 'Vr1=',Vr1 iefec2(i,1)=int(u*ncro)+1 iefec2(i,2)=int(u*ncro)+1 iefec2(i,2)=int(u*ncro)+1 iefec2(i,2)=int(u*ncro)+1 ! assignment of the effect call normal(x1,u) efec2(i)=u*1 write(16,*) i,((iefec2(i,1)-1)*nmar+iefec2(i,2)),efec2(i) end do close(16) $\begin{aligned} & close(16) \\ & \underline{! frequencies of the causative mutations of trait 2} \\ & open(16.file='freq2-base.txt) \\ & do i=1, ngenes \\ & p2base(iefec2(i,1),iefec2(i,2))=0 \\ & do k=1,namin \\ & p2base(iefec2(i,1),iefec2(i,2))=p2base(iefec2(i,1),iefec2(i,2))+real(marn(k,iefec2(i,1),iefec2(i,2),1)+marn(k,iefec2(i,1),iefec2(i,2),2)-2) \\ & end do \\ & p2base(iefec2(i,1),iefec2(i,2))=p2base(iefec2(i,1),iefec2(i,2))/(2*real(nanim)) \\ & write(16,*) i,(iefec2(i,1)-i*nmar+iefec2(i,2)),p2base(iefec2(i,1),iefec2(i,2)) \\ & end do \end{aligned}$

```
<u>! additive variance of trait 2</u>
Va2=0
  \label{eq:va2} \begin{array}{l} Va2{=}0 \\ oi=1, ngenes \\ va2(i){=}2^{*}p2base(iefec2(i,1), iefec2(i,2))^{*}(1-p2base(iefec2(i,1), iefec2(i,2)))^{*}efec2(i)^{*}{+}2 \\ Va2{=}Va2{+}va2(i) \end{array}
  \label{eq:value} \begin{array}{l} value = value + var2(1) \\ end do \\ mitter + (Va2 base = :,Va2 \\ write(11, *) (Va2 base = :,Va2 \\ \underline{l \ resistal \ value and resistance \ of \ trait 2 \\ Vr2 = (Va2)(0, 4) Va2 \\ print *, (Vr2 = :,Vr2 \\ write(11, *) (Vr2 = :,Vr2 \\ \end{array}
\label{eq:constraint} \begin{array}{l} write(11, \cdot) \ V12=, V12\\ \hline w
end do
end do
ait=0
do i=1,2*ngenes
if (ef(i)==ef(i-1)) then
goto 40
else
ait=ait+1
write(16,4*, ait,ef(i)
end if
40 end do
close(16)
print *, N* of genes=',ait
open(16,file='neutral markers.txt')
neu=0
do i=1,.exo*nmar
do j=1,2*ngenes
if (i==ef(j)) then
end if
                                            end if
end if
end do
neu=neu+1
n(neu)=i
write(16,*) neu,n(neu)
9 end do
print *, 'N° of neutral markers=',neu
close(16)
write(11,*) 'N° of neutral markers=',neu
<u>I Population A under selection for trait 1</u>
marnA=marn
<u>I simulation of phenotypes (mean=100)</u>
do i=1,naim
yfenA(i)=100
do j=1,ngenes
yfenA(i)=yfenA(i)+(marnA(i,iefec1(j,1),iefec1(j,2),1)+marnA(i,iefec1(j,1),iefec1(j,2),2)-3)*efec1(j)
end do
call normal(x1,u)
yfen A(i)=yfenA(i)+(marnA(i,iefec1(j,1),iefec1(j,2),1)+marnA(i,iefec1(j,2),2)-3)*efec1(j)
wfen A(i)=yfenA(i)+(marnA(i,iefec1(j,1),iefec1(j,2),1)+marnA(i,iefec1(j,2),2)-3)*efec1(j)
     yfenA(i)=yfenA(i)+u*sqrt(Vr1)
end do
     ! phenotypic mean
  <u>l'phenotypic mean</u>
medial=0
do i=1,nanim
medial=medial+yfenA(i)
end do
medial=medial/nanim
end do

medial =medial /nanim

<u>l phenotypic standard deviation</u>

des1=0

do i=1,nanim

des1=des1+(yfenA(i)-media1)**2

end do

des1=aqrt(des1)

print *, 'mean & s. d. trait 1 =',media1,des1

write(11,*)'mean & s. d. =',media1,des1

open(20,file='pheno_mean_Atxt')

open(21,file='variance_1Atxt')

open(22,file='variance_1Atxt')

open(22,file='variance_1Atxt')

open(22,file='variance_1Atxt')

open(22,file='variance_1Atxt')

lselection for n generations

do ijk=1,ngen2

do ijk=1,ngen2

do ijk=1,ngen2

do ijk=1,ngen3

la lunif(x1,u)

ip=int(u*nnac)+1
                          ip=int(u*nmac)+1
if (yfenA(ip).lt.media1-des1) then
goto 1
end if
          n vy-
end if
2 call unif(x1,u)
im=in((u*nhem)+nmac+1
if (yfenA(im).lt.media1-des1) then
goto 2
     end if

<u>! father's chromosome</u>

do j=1,ncro

call unif(x1,u)

if (u.lt.0.5) then

iale=1

else

iale=2

z=#f
                                         endif
grupoA(i,j,1,1)=marnA(ip,j,1,iale)
lam=1.
                                    lam=1.

nrec=pois(lam,x1)

do ik=1,nrec

call unif(x1,u)

irec(ik)=int(u*(nmar-1))+1
                       lo
irecord=0
do ik=1,nrec
xmin=999999
do ij=1,nrec
if (irec(j),lt.xmin) then
xmin=irec(ij)
imin=ij
       enddo
```

close(16)

endif enddo irecord(ik)=xmin irec(imin)=999999 enddo irecord(nrec+1)=nmar if (nrec.eq.0) then do k=2,nmar $\begin{array}{l} \mbox{do k=2,mar} \\ \mbox{grupo}\ A(i,j,k,1) = mam\ A(ip,j,k,iale) \\ \mbox{enddo} \\ \mbox{else} \\ \mbox{ic=2} \\ \mbox{do k=1,nrec+1$} \\ \mbox{do k=1,nrec+1$} \\ \mbox{do k=i,irecord(ik)$} \\ \mbox{grupo}\ A(i,j,k,1) = mam\ A(ip,j,k,iale) \\ \mbox{enddo} \end{array}$ enddo ic=irecord(ik)+1 if (iale.eq.1) then iale=2 else iale=1 endif enddo endif ! mutation $lam=1, \\ mmt=pois(lam,x1) \\ do ik=1,mmt \\ call unif(x1,u) \\ immt=int(u^*nmar)+1 \\ if (grupoA(i,j,immt,1)=2) \\ else \\ grupoA(i,i,immt,1)=2 \\ else \\ rupoA(i,i,immt,1)=2 \\$ lam=1. grupoA(i,j,imut,1)=1 endif enddo enddo enddo <u>! mother's chromosome</u> do j= 1,ncro call unif(x1,u) if (u.lt.0.5) then iale=1 else iale=2 endif erupoA(i,j,1,2)=1 iale=2 endif grupoA(i,j,1,2)=marnA(im,j,1,iale) lam=1. nrec=pois(lam,x1) do ik=1,nrec call unif(x1,u) irecord=0 do ik=1,nrec xmin=99999 do ij=1,nrec if (irec(ij).lt.xmin) then xmin=irj endif enddo enddo irecord(ik)=xmin irec(imin)=999999 enddo irecord(nec+1)=nmar if (nrec.eq.()) then do k=2,nmar grupoA(i,j,k,2)=marnA(im,j,k,iale) enddo else enddo enddo else ic=2 do ik=1,nrec+1 do k=ic,irecord(ik) grupoA(i,j,k,2)=marnA(im,j,k,iale) enddo ic=irecord(ik)+1 if (iale.eq.1) then iale=2 '~=1 endif enddo endif ! mutation lam=1. mutt=pois(lam,x1) do ik=1.mmut call unif(x1,u) imut=int(u*nmar)+1 if (grupoA(i,j,imut,2),eq.1) then grupoA(i,j,imut,2)=2 else lam=1. grupoA(i,j,imut,2)=1 endif enddo $enddo\\enddo\\enddo\\i = indiadion of phenotypes\\do i = i, naemic$ $yfenA(i) = 100\\do j = i, ngenes$ $yfenA(i) = fonA(i) + (grupoA(i, iefec1(j, 1), iefec1(j, 2), 1) + grupoA(i, iefec1(j, 1), iefec1(j, 2), 2) - 3) * efec1(j)\\end do\\call normal(x1, u)$ yfenA(i) = yfenA(i) + u*sqrt(Vr1) $end do\\j mean$ <u>/ mean</u> media1=0 do i=1,nanim media1=media1+yfenA(i) end do media1=media1/nanim medial =medial /nanim write(20,*) jik.medial /<u>tstandard deviation</u> des1=0 des1=0 des1=(stanim-1) des1=agtr(des1) ges1=agtr(des1) print *, ijk.medial.des1

<u>l frequencies of the causative mutations of trait 1</u> do i=1,ngenes p1(iefec1(i,1),iefec1(i,2))=0 $\begin{array}{l} (\operatorname{ret}(i_1),\operatorname{ret}(i_2)) = 0 \\ obs(1,\operatorname{nini}) = 0 \\ obs(1,i_1), \operatorname{iefec1}(i_1,2) = 1 \\ (\operatorname{iefec1}(i_1,1),\operatorname{iefec1}(i_2,2)) = 1 \\ (\operatorname{iefec1}(i_1,1),\operatorname{iefec1}(i_1,2)) = 1 \\ (\operatorname{iefec1}(i_1,2),\operatorname{iefec1}(i_1,2)) = 1 \\ (\operatorname{iefec1}(i_1,2),$ $p1(iefec1(i,1),iefec1(i,2))=p1(iefec1(i,1),iefec1(i,2))+real(grupoA(k,iefec1(i,1),iefec1(i,2),1)+grupoA(k,iefec1(i,1),iefec1(i,2),2)-2) end do \\ p1(iefec1(i,1),iefec1(i,2))=p1(iefec1(i,1),iefec1(i,2))/(2*real(nanim)) end do \\ ladditive variance trait l \\ VarI=0 \\ varA1(i)=0 \\ do i=1,ngenes \\ varA1(i)=2*p1(iefec1(i,1),iefec1(i,2))*(1-p1(iefec1(i,1),iefec1(i,2)))*efec1(i)**2 \\ VaI=VaI=VaI=VaI=VaI=VaIA(i) \\ end do \\ lprint *, additive variance trait 1 popul A=',ijk,VaI \\ write(21,*) ijk,VaI_VrI \\ lfrequencies for causative mutations trait 2 \\ do i=1,ngenes \\ p2(iefec2(i,1),iefec2(i,2))=p2(iefec2(i,1),iefec2(i,2))+real(grupoA(k,iefec2(i,1),iefec2(i,2),1)+grupoA(k,iefec2(i,1),iefec2(i,2),2)-2) \\ end do \\ p2(iefec2(i,1),iefec2(i,2))=p2(iefec2(i,2))+real(grupoA(k,iefec2(i,1),iefec2(i,2),1)+grupoA(k,iefec2(i,1),iefec2(i,2),2)-2) \\ end do \\ p2(iefec2(i,1),iefec2(i,2))=p2(iefec2(i,2))+real(grupoA(k,iefec2(i,1),iefec2(i,2),1)+grupoA(k,iefec2(i,2),2)-2) \\ end do \\ p2(iefec2(i,1),iefec2(i,2))=p2(iefec2(i,2))+real(grupoA(k,iefec2(i,1),iefec2(i,2),1)+real(grupoA(k,iefec2(i,2),1)+real(grupoA(k,iefec2(i,2),1)+real(grupoA(k,iefec2(i,2),1)+real(grupoA(k,iefec2(i,2),1)+real(grupoA(k,iefec2(i,2),1)+real(grupoA(k,iefec2(i,2),1)+real(grupoA(k,iefec2(i,2),1)+real(grupoA(k,iefec2(i,2),1)+real(grupoA(k,iefec2(i,2$ end do p2(iefec2(i,1),iefec2(i,2))=p2(iefec2(i,1),iefec2(i,2))/(2*real(nanim)) end do <u>! additive variance trait 2</u> Va2=0 $\begin{array}{l} Va2{=}0 \\ varA2(i){=}0 \\ do i{=}1, genes \\ varA2(i){=}2^*p2(iefec2(i,1),iefec2(i,2))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*efec2(i)**2 \\ Va2{=}2^*p2(iefec2(i,1),iefec2(i,2))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*efec2(i)**2 \\ va2{=}2^*p2(iefec2(i,1),iefec2(i,2))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,1),iefec2(i,2))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,1),iefec2(i,2))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,1),iefec2(i,2)))*(1{-}p2(iefec2(i,2)))*(1$! Final population A do i=1,nanim2 3 call unif(x1,u) 3 call unif(x1,u) ip=in(u⁺mmac)+1 if (yfenA(ip)l.tmedia1-des1) then goto 3 end if 4 call unif(x1,u) im=in(u(*nhem)+nmac+1) if (yfenA(im).t.media1-des1) then goto 4 end if for here: g005 end if <u>l'father's chromosome</u> do j=1,ncro call unif(x1,u) if (u.lt.0.5) then iale=1 else iale=2 endif finalA(i,j,1,1)=grupoA(ip,j,1,iale) lam=1, mrcc=pois(lam,x1) do ik=1,mrc call unif(x1,u) irec(ik)=in(u*(nmar-1))+1 call unit(x1,u) irec(ik)=int(u*(nmar-1))+1 enddo irecord=0 !desde aqui ordeno do ik=1,nrec xmin=999999 do ij=1,nrec if (irec(ij).lt.xmin) then xmin=irec(ij) imin=ij endif enddo nddo irecord(ik)=xmin irec(imin)=999999 enddo irecord(nrec+1)=nmar if (nrec.eq.0) then do k=2_nmar finalA(i,j,k,1)=grupoA(ip,j,k,iale) enddo else else ic=2 tc=∠ do ik=1,nrec+1 do k=ic,irecord(ik) finalA(i,j,k,1)=grupoA(ip,j,k,iale) enddo ic=irecord(ik)+1 if (iale.eq.1) then iale=2 else iale=1 $\label{eq:alpha} \begin{array}{c} iale=1\\ endif\\ endif\\ !mutation\\ lam=1,\\ nnut=pois(lam,x1)\\ do ik=1,nmut\\ call unif(x1,u)\\ inut=in(u^{*}nnaz)+1\\ if(final A(i,j,imut,1).eq,1) then\\ final A(i,j,imut,1).=2\\ else \end{array}$ else finalA(i,j,imut,1)=1 endif enddo nate=2 endif finalA(i,j,1,2)=grupoA(im,j,1,iale) lam=1. nrec=pois(lam,x1)

do ik=1,nrec call unif(x1,u) irec(ik)=int(u*(nmar-1))+1 enddo irecord=0 do ik=1,nrec xmin=999999 do ij=1,nrec if (irec(ij).lt.xmin) then xmin=irec(ij) imin=ii imin=ii endif enddo do irecord(ik)=xmin irecord(imin)=999999 enddo irecord(irec+1)=nmar if (nrec.eq,0) then do k=2,nmar finalA(i,j,k,2)=grupoA(im,j,k,iale) enddo enddo else ic=2 do ik=1,nrec+1 do k=ic,irecord(ik) finalA(i,j,k,2)=grupoA(im,j,k,iale) enddo cindu0 ic=irecord(ik)+1 if (iale.eq.1) then iale=2 else else iale=1 endif enddo endif endur <u>! mutation</u> lam=1. nmut=pois(lam,x1) do ik=1,nmut call unif(x1,u) imut=in imut=int(u*nmar)+1 if (finalA(i,j,imut,2).eq.1) then finalA(i,j,imut,2)=2 else finalA(i,j,imut,2)=1 endif enddo enddo enddo enddo <u>!frequencies of the causative mutations of trait]</u> open (16,file="freq1-A.txl") do i=1,ngenes p1A(iefec1(i,1),iefec1(i,2))=0 do k=1,nanim2 p1A(iefec1(i,1),iefec1(i,2))=p1A(iefec1(i,1),iefec1(i,2))+real(finalA(k,iefec1(i,1),iefec1(i,2),1)+finalA(k,iefec1(i,1),iefec1(i,2),2)-2) end do p1A(iefec1(i,1),iefec1(i,2))=p1A(iefec1(i,1),iefec1(i,2))/(2#eod(nonin2)) end do pl A(iefecl (i,1),iefecl (i,2))=pl A(iefecl (i,1),iefecl (i,2))/(2*real(nanim2)) write(16,*) i,((iefecl (i,1)-1)*nmar+iefecl (i,2)),pl A(iefecl (i,1),iefecl (i,2)) end do close(16) <u>l additive variance of trait 1</u> Val=0 varAf1=0 do i=1,ngenes wreAf1(i)=2*ntA(iefecl (i,1),iefecl (i,2))*(1,ntA(iefecl (i,1),iefecl (i,2))* $\label{eq:last} \begin{array}{l} varAf1(i)=2*p1A(iefec1(i,1),iefec1(i,2))*(1-p1A(iefec1(i,1),iefec1(i,2)))*efec1(i)**2\\ Val=Va1+varAf1(i) \end{array}$ end do print *, 'Variance trait 1 population A=',Val write(11,*) 'Variance trait 1 population A=',Val write(11,*) Variance trait 1 population A=;Va1 [/frequencies of the causative mutation of trait 2 open(16,file=freq2-Atxt') do i=1,ngenes p2A(iefec2(i,1),iefec2(i,2))=0 do k=1,nanim2 $p2A(icfec2(i,1),icfec2(i,2))=0 \\ do k=1,nanim2 \\ p2A(icfec2(i,1),icfec2(i,2))=p2A(icfec2(i,1),icfec2(i,2))+real(finalA(k,icfec2(i,1),icfec2(i,2),1)+finalA(k,icfec2(i,1),icfec2(i,2),2)-2) \\ end do \\ p2A(icfec2(i,1),icfec2(i,2))=p2A(icfec2(i,1),icfec2(i,2))/(2*real(nanim2)) \\ write(16,*),ic(icfec2(i,1)-1)*nmar+icfec2(i,2)),p2A(icfec2(i,1),icfec2(i,2)) \\ end do \\ close(16) \\ \underline{Ladditive variance trait 2 \\ Va2=0 \\ VarAD2=0 \\ do i=1,ngenes \\ varAD2(D=2*p2A(icfec2(i,1),icfec2(i,2))*(1-p2A(icfec2(i,1),icfec2(i,2)))*efec2(i)**2 \\ va2=Va2+varA12(i) \\ end do \\ \hline end end \\ \hline end do \\ \hline end end \\ \hline end end \\ \hline end end \\ \hline end \\ end \\ \hline en$ varA2(i)=2*p2A(iefce2(i,1),iefce2(i,2))*(1-p2A(iefce2(i,1))iefce2(i,2)))*efce2(i)**2
va2=Va2+varA2(i)
end do
pint *, Variance trait 2 population A=',Va2
write(11,*) Variance trait 2 population A=',Va2
write(11,*) Variance trait 2 population A=',Va2
vrite(11,*) Variance trait 2 population A=',Va2
vrite(11,*) Variance trait 2 population A=',Va2
vrite(14,file=T2Atxt)
do i=1,namin2
yfen1A(i)=yfen1A(i)+(finalA(i,iefce1(j,1),iefce1(j,2),1)+finalA(i,iefce1(j,1),iefce1(j,2),2)-3)*efce1(j)
end do
call normal(x1,u)
yfen1A(i)=yfen1A(i)+u*sqtt(Vr1)
write(13)
do i=1,namin2
yfen2A(i)=1000
do
j=1,ngenes
yfen2A(i)=1000 do j=1,ngenes yfen2A(i)=yfen2A(i)+(finalA(i,iefec2(j,1),iefec2(j,2),1)+finalA(i,iefec2(j,1),iefec2(j,2),2)-3)*efec2(j) end do call normal(x1,u) yfen2A(i)=yfen2A(i)+u*sqrt(Vr2) write(14,*) yfen2A(i),i,1 end do close(14) close(14) <u>! write genotypes</u> open(13,file='gA.txt') do i=1,nanim2 do j=1,ncro do k=1,nmar write(13,*) i,((j-1)*nmar+k),finalA(i,j,k,1),finalA(i,j,k,2) end do end do end do

close(13) <u>! marker and gene frequencies</u> open(13,file='pA.txt') $\label{eq:constraints} \begin{array}{l} \hline constraints} \\ constraints \\ constraints$ ar do g=1,ngenes if (j==iefec1(g,1),and,k==iefec1(g,2)) then goto 25 end if if (j==iefec2(g,1),and,k==iefec2(g,2)) then goto 25 end if end do write(13*) i, (j=1)*nwar4k (final A(i ik 1)k) ewrite(13,*) i,(j-1)*nmar+k,(final/ 25 end do end do end do close(13) *Livrite neutral marker genotypes in format* open(20,file='nA.txt') open(13,file='genotA.txt') do i=1,anim2 marca(j)=0 do j=1.neu read(20,*) ipim,iv marca(j)=iv+1 end do vrite(13,'(4,1x,30000i1)') i,(marca(j),j=1,neu) end do close(20) *Litue breeding values for traits 1 & 2* open(14,file='tbv1A.txt') open(15,file='tbv2A.txt') do i=1,anim2 bv1A(j)=0 do j=1,ngenes bt1A(j)=bv1A(j)=bv1A(j) iofere1(j1) iofere1/ write(13,*) i,(j-1)*nmar+k,(finalA(i,j,k,1)+finalA(i,j,k,2)-3) $\label{eq:interm} \begin{array}{l} \text{intermal} \mbox{intermal} \mbox{int$ $\label{eq:constraint} \begin{array}{l} {}^{b}_{i}(Ai)=bv1A(i)+(finalA(i,iefec1(j,1),iefec1(j,2),2)-3)^*efec1(j)\\ end do\\ write([4,*),i,bv1A(i)\\ end do\\ close(14)\\ do i=i,namin2\\ bv2A(i)=0\\ do j=1,nagenes\\ bv2A(i)=bv2A(i)+(finalA(i,iefec2(j,1),iefec2(j,2),1)+finalA(i,iefec2(j,1),iefec2(j,2),2)-3)^*efec2(j)\\ end do\\ write(15,*),i,bv2A(i)\\ end do\\ close(15) \end{array}$! Population B under selection for trait 2 marnB=marn <u>! simulation of phenotypes (media=1000)</u> do i=1,nanim yfenB(i)=1000 end if 6 call unif(x1,u) if (yfenB(im).lt.media2-des2) then goto 6 end if ! father's chromosome do j=1,ncro call unif(x1,u) if (u.lt.0.5) then iale=1 else iale=2 endif endif grupoB(i,j,1,1)=marnB(ip,j,1,iale) lam=1. nrec=pois(lam,x1) do ik=1,nrec call unif(x1,u) irec(ik)=int(u*(nmar-1))+1 do

enddo

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irecord=0 do ik=1,nrec xmin=9999999 999 do ij=1,nec if (irec(ij).lt.xmin) then xmin=irec(ij) imin=ij endif enddo irecord(ik)=xmin irec(imin)=999999 ntecord(nrec+1)=nmar if (nrec.eq.0) then do k=2,nmar grupoB(i,j,k,1)=marnB(ip,j,k,iale) enddo else ic=2 do ik=1,nrec+1 do k=ic,irecord(ik) grupoB(i,j,k,1)=mamB(ip,j,k,iale) enddo ic=irecord(ik)+1 if (iale.eq.1) then iale=2 else iale=1 endif enddo enddo endif mutation lam=1 ! nmut=pois(lam,x1) do ik=1,nmut call unif(x1,u) imut=int(u*nmar)+1 if (grupoB(i,j,imut,1)=q,1) then grupoB(i,j,imut,1)=2 else erunoB(i,i,imut,1)=1 grupoB(i,j,imut,1)=1 enddo enddo if (irec(ij).lt.xmir xmin=irec(ij) imin=ij endif enddo irecord(ik)=xmin irec(imin)=999999 $\label{eq:cond} \begin{array}{l} \mbox{recond}(arce+1) = mar & \mbox{if (arce+1)} & \mbox{do } k = 2, \mbox{mar } \\ \mbox{grupoB}(i,j,k,2) = mar B(im,j,k,iale) & \mbox{end} & \mbox{else} & \mbox{else} & \mbox{ic=2} & \mbox{do } k = 1, \mbox{arcevil} k + 1 & \mbox{if (alce+2)} & \mbox{end} & \mbox{end} & \mbox{else} & \mbox{end} & \mbox{end} & \mbox{end} & \mbox{ic=1} & \mbox{end} & \mbox{end} & \mbox{ic=1} & \mbox{end} & \mbox{$ enddo endif enddo endif endif <u>Imutation</u> lam=1, nmut=pois(lam,x1) do ik=1,mmut call unif(x1,u) imut=int(u⁴nmat)+1 if (grupoB(i,jimut,2)=2) -~~nB(i,j,imut,2)=1 else grupoB(i,j,imut,2)=1 endif enddo enddo (*simulation of phenotypes* do i=1,nanim yfenB(i)=yfenB(i)+(grupoB(i,iefec2(j,2),1)+grupoB(i,iefec2(j,2),2)-3)*efec2(j) end do call normal(x1,u) yfenB(i)=yfenB(i)+u*sqrt(Vr2) end do [*phenotypic mean* media2=media2+yfenB(i) end do *i_phenotypic mean* media2=media2+yfenB(i) end do *i_phenotypic mean* media2=media2/nanim write(20,*) ijk.media2 *i_phenotypic standard deviation*

des2=0 do i=1,nanim des2=des2+(yfenB(i)-media2)**2 do i=1,naini des2=des2+(yfenB(i)-media2)**2 end do des2=des2/(naini-1) des2=des2/(naini-1) des2=des2/(naini-1) des2=des2/(naini-1) des2=des2/(naini-1) des2=des2/(naini-1) des2=des2/(naini-1) des2=des2/(naini-1) des2=des2/(naini-1) for end interval (i,1); end do fadditive variance of trait I Val=0 var1B(i)=0 do i=1,ngenes var1B(i)=1 (i,1); des2(iefec1(i,1); des2(iefec1(i,1); des2(iefec1(i,2))=0 (iefec1(i,2))=0 (ief Val=Val+var1B(i) end do write(21,*) ijk,Val <u>/frequencies of the causative mutations of trait 2</u> do i=1.ngenes p2sel(iefec2(i,1),iefec2(i,2))=0 D2set(letec2(1,1),letec2(1,2))=0 do k=1,naim p2set(letec2(1,1),letec2(i,2))=p2set(letec2(i,1),letec2(i,2))+real(grupoB(k,letec2(i,1),letec2(i,2),1)+grupoB(k,letec2(i,1),letec2(i,2),2)-2) end do end do p2sel(iefec2(i,1),iefec2(i,2))=p2sel(iefec2(i,1),iefec2(i,2))/(2*real(nanim))) end do $\frac{l additive variance of trait2}{var2B(i)=0}$ var2B(i)=0 do i=1.ngenes var2B(i)=2*p2sel(iefec2(i,1),iefec2(i,2))*(1-p2sel(iefec2(i,1),iefec2(i,2)))*efec2(i)**2 Va2=Va2+var2B(i) end do write(22,*) ijk,Va2 mamB=grupoB end do end do close(20) close(21) close(22) <u>! Final population B</u> do i=1,nanim2 7 call unif(x1,u) ip=int(u*nmac)+1 if (yfenB(ip).lt.media2-des2) then if (yfenB(ip).lt.media2-des2) then goto 7 end if 8 call unif(x1,u) im=int(u*nhem)+nmac+1 if (yfenB(im).lt.media2-des2) then goto 8 end if <u>/father's chromosome</u> do j=1,ncro call unif(x1,u) if (u.lt.0.5) then iale=1 else else iale=2 endif finalB(i,j,1,1)=grupoB(ip,j,1,iale) nrec=pois(lam,x1) do ik=1,nrec do ik=1,mec call unif(x1,u) irec(ik)=in(u*(nmar-1))+1 enddo irecord=0 do ik=1,mec xmin=999999 do ij=1,arec if (irec(ij).lt.xmin) then xmin=irec(ij) imin=ij endif endif enddo irecord(ik)=xmin irec(imin)=999999 recci(mn)=999999 enddo irecord(nrec+1)=nmar if (nrec.eq.0) then do k=2,nmar finalB(i,j,k,1)=grupoB(ip,j,k,iale) finalB(i,j,k,1)-0-enddo else ic=2 do ik=1,nrc+1 do k=ic,irecord(ik) finalB(i,j,k,1)=grupoB(ip,j,k,iale) enddo ic=irecord(ik)+1 if (iale.eq.1) then iale=2 else iale=1 iale=1 endif endif ! mutar este cromosoma lam=1. mmut=pois(lam,x1) do ik=1,nmut call unif(x1,u) imut=int(u*mmar)+1 if (finalBG i innut. 1br if (finalB(i,j,imut,1).eq.1) then finalB(i,j,imut,1)=2 else finalB(i,j,imut,1)=1 f enddo enddo ! mother's chromoo do j=1,ncro call unif(x1,u) if (u.lt.0.5) then iale=1

else iale=2 endif finalB(i,j,1,2)=grupoB(im,j,1,iale) lam=1. lam=1. nrec=pois(lam,x1) do ik=1,nrec call unif(x1,u) irec(ik)=int(u*(nmar-1))+1 enddo irecord=0 irecord=0 do ik=1,nrec xmin=999999 do ij=1,nrec if (irec(ij).lt.xmin) then xmin=irec(ij) endif enddo imin=ij ddo irecord(ik)=xmin irecc(imin)=999999 enddo irecord(nrec+1)=nmar if (nrec.eq.0) then do k=2,mmar finalB(i,j,k,2)=grupoB(im,j,k,iale) enddo enddo else ic=2 do ik=1.nrec+1 do ik=i,urec+1
 do k=ic.irecord(ik)
 frnalB(i,j,k,2)=grupoB(im,j,k,iale)
 enddo
 ic=irecord(ik)+1
 if (iale.eq.1) then
 iale=2
 else
 iale=1
 endif
enddo endif <u>! mutation</u> lam=1. nmut=pois(lam,x1) do ik=1,nmut call unif(x1,u) imut=int(u*nmar)+1 if (finalB(i,j,imut,2).eq.1) then finalB(i,j,imut,2)=2 else else finalB(i,j.imut,2)=1 endif enddo enddo enddo enddo enddo enden popen(16.file=Treq2-B.txt') do i=1.ngenes p2B(iefec2(i,1).iefec2(i,2))=0 do k=1.nanim2 p2B(iefec2(i,1).iefec2(i,2))=p2B(iefec2(i,1).iefec2(i,2))+ real(finalB(k,iefec2(i,1).iefec2(i,2),1)+finalB(k,iefec2(i,1),iefec2(i,2),2)-2) end do end do p2B(iefec2(i,1),iefec2(i,2))=p2B(iefec2(i,1),iefec2(i,2))/(2*real(nanim2)) write(16,*) i,((iefec2(i,1))=p2B(iefec2(i,1),iefec2(i,2))/(2*real(nanim2)) end do close(16) <u>l additive variance of trait 2</u> Va2=0 var2Bf(i)=0 Val=N(1)=0 do i=1.genes var2Bf(i)=2*p2B(iefec2(i,1),iefec2(i,2))*(1-p2B(iefec2(i,1),iefec2(i,2)))*efec2(i)**2 Va2=Va2+var2Bf(i) $\label{eq:statistical_statis$ $\begin{array}{l} v_{a1} = 0 \\ w_{a1}B(i) = 0 \\ do i = 1, ngenes \\ warl Bf(i) = 2^* p1 B(iefec1(i,1), iefec1(i,2))^* (1-p1 B(iefec1(i,1), iefec1(i,2)))^* efec1(i)^{*+} 2 \\ v_{a1} = v_{a1} + var1 Bf(i) \\ v_{a1} = v_{a1} + var1 Bf(i) \end{array}$ $\label{eq:alpha} \begin{array}{l} Val=Val+varlBf(i) \\ end do \\ print *, Varaince trait 1 population B=',Val \\ write(11,*) Variance trait 1 population B=',Val \\ \underline{j:simulation of phenotype:} \\ open(13,file='12B,txt') \\ open(15,file='1B,txt') \\ do := 1,nanim2 \\ yfen2B(i)=1000 \\ do := 1,negenes \\ yfen2B(i)=yfen2B(i)+(finalB(i,iefec2(j,1),iefec2(j,2),1)+finalB(i,iefec2(j,1),iefec2(j,2),2)-3)*efec2(j) \\ end do \end{array}$ yfen2B(i)=yfen2B(i)+(finalB(i,iefe end do call normal(x1,u) yfen2B(i)=yfen2B(i)+u*sqrt(Vr2) write(13,*) yfen2B(i),i,1 end do close(13) do i=1,nanim2 $\begin{array}{l} yfen 1 B(i){=}100 \\ do \; j{=}1, ngenes \\ yfen 1 B(i){=}yfen 1 B(i){+}(final B(i, jefec 1(j, 1), jefec 1(j, 2), 1){+}final B(i, jefec 1(j, 1), jefec 1(j, 2), 2){-}3)^*efec 1(j) \\ end \; do \\ call normal(x1,u) \\ yfen 1 B(i){=}yfen 1 B(i){+}u^*sqnt(Vr1) \\ write(15,*)\; yfen 1 B(i), i, 1 \\ end \; do \\ close(15) \\ \underline{/write\; genotypes} \\ open(13, file{=}gB, txt') \\ do \; i{=}1, nanim2 \\ \end{array}$ vfen1B(i)=100

do j=1,ncro do k=1,nmar write(13,*) i,((j-1)*nmar+k),finalB(i,j,k,1),finalB(i,j,k,2) end do end do end do close(13)
$$\label{eq:close} \begin{split} &close(13) \\ &\underline{-lmarker} \ and \ ener \ frequencies} \\ &open(13, file="pB.txt") \\ &do \ i=1, ncro \\ &do \ j=1, nmar \\ &pB=0 \\ &do \ k=1, nanim2 \\ &pB=pB+real(finalB(k,i,j,1)+finalB(k,i,j,2)-2) \\ &end \ do \end{split}$$
pB=pB+real(finalB(k,i,j,1)+finalI end do write(13,*) (i-1)*nmar+j,pB/(2*real(nanim2)) end do close(13) <u>! write neutral marker genotypes</u> open(13,file='nB.txt') do i=1,nanim2 do i=1,=== do j=1,ncro do k=1,nmar do g=1,ngenes if (j==iefec1(g,1).and.k==iefec1(g,2)) then goto 26 end if if (j==iefec2(g,1).and.k==iefec2(g,2)) then goto 26 goto 26 end if end do write(13,*) i,(j-1)*nmar+k,(finalB(i,j,k,1)+finalB(i,j,k,2)-3) 26 end do end do end do close(13) <u>!write neutral marker genotypes in format</u> open(20,file='nB.txt') open(13,file='genotB.txt') do i=1,nanim2 marca(j=0 marca(j)=0 do j=1,neu read(20,*) ip,im,iv marca(j)=iv+1 end do write(13,'(i4,1x,30000i1)') i,(marca(j),j=1,neu) end do end do close(13) close(20) $\begin{array}{l} close(20) \\ \underline{1\, true \, breeding \, values \, for \, traits \, 1 \, \& \, 2}{open(14, file=1-tbv28, txt)} \\ open(15, file=1-tbv18, txt) \\ do \, i=1, namine \\ bv28(i)=0 \\ do \, j=1, ngenes \\ bv28(i)=bv28(i)=vk28(i)+(final8(i, iefec2(j,1), iefec2(j,2), 1)+final8(i, iefec2(j,1), iefec2(j,2), 2)-3)*efec2(j) \\ end \, do \\ write(14, *), i, bv2B(i) \\ end \, do \\ close(14) \\ do \, i=1, namin2 \end{array}$ do i=1,nanim2 bv1B(i)=0
$$\label{eq:linear} \begin{split} &h(i,j) = h(i,j) = h(i,j)$$
end do write(15,*) i,bv1B(i) end do close(15) <u>! population C under random mating</u> martG=marn <u>t random mating for n generations</u> do ijk=1.ngen2 do i=1.nanim call unif(x1,u) ip=int(u*nmac)+1 call unif(x1,u) im=int(u*nhem)+nmac+1 <u>! father's chromosome</u> do j=1.ncro call unif(x1,u) if (u.1t.0.5) then iale=1 else iale=2 endif grupoC(i,j,1,1)=marrC(ip ! population C under random mating endif grupoC(i,j,1,1)=marnC(ip,j,1,iale) lam=1. nrec=pois(lam,x1) do ik=1,nrec call unif(x1,u) inddo irecord=0 do ik=1.mec xmin=999999 do ij=1.mec if (irec(ij).lt.xmin) then xmin=irec(ij) imin=ij endif irec(ik)=int(u*(nmar-1))+1 nddo irecord(ik)=xmin irec(imin)=9999999 enddo irecord(nrec+1)=nmar if (nrec.eq.0) then do k=2,nmar areas (Ci ik h) merer grupoC(i,j,k,1)=marnC(ip,j,k,iale) enddo else ic=2 do ik=1,nrec+1 io ik=1,nrec+1 do k=ic,irecord(ik) grupoC(i,j,k,1)=marnC(ip,j,k,iale) grup enddo ic=irecord(ik)+1 if (iale.eq.1) then iale=2 else

iale=1 endif enddo endif vm=' ! mutation nmut=pois(lam,x1) do ik=1,nmut call unif(x1,u) imut=int(u*nmar)+1 if (grupoC(i,j,imut,1).eq.1) then grupoC(i,j,imut,1)=2 else grupoC(i,j,imut,1)=1 endif endo <u>i mother's chronosome</u> do j= 1,ncro call unif(x1,u) if (u.1t.05.) then iale=1 else iale=2 endif "poC(i,j.1 etse iale=2 endif grupoC(i,j,1,2)=marnC(im,j,1,iale) lam=1. nrec=pois(lam,x1) do ik=1,nec call unif(x1,u) irrec(ik)=int(u*(nmar-1))+1 enddo irrecord=0 do ik=1,nec xmin=999999 do j=1,nec if (irrec(ij).lt.xmin)then xmin=irrec(ij) imin=ij endif endif enddo irecord(ik)=xmin irec(imin)=999999 irec(imin)=999999 enddo irecord(ince+1)=nmar if (ince:eq.0) then do k==2,nmar grupoC(i,j,k,2)=marnC(im,j,k,iale) enddo else ic=2 do ik=1,nrec+1 do k==i.riccord(ik) grupoC(i,j,k,2)=marnC(im,j,k,iale) enddo ic=irecord(ik)+1 if (iale.eq.1) then id=2 else else iale=1 endif enddo endif ! mutation lam=1. nmut=pois(lam,x1) do ik=1,nmut call unif(x1,u) imut=im((u*nmar)+1 if (grupoC(i,jimut,2)=c,1) then grupoC(i,jimut,2)=2 else grupoC(i,jimut,2)=2 lam=1. erse grupoC(i,j,imut,2)=1 endif enddo enddo enddo marnC=grupoC print*, ijk end do Lina do <u>L'Final population C</u> do i=1,nanim2 call unif(x1,u) ip=int(u*nmac)+1 call unif(x1,u) im=int(u*nhem)+nmac+1 <u>L'father's chromosoma</u> do j=1,ncro call unif(x1,u) if (u.lt.0.5) then iale=1 else else else iale=2 endif finalC(i,1,1)=grupoC(ip,1,1,iale) lam=1. mrcc=pois(lam,x1) do ik=1.mrec call unif(x1,u) irec(ik)=int(u*(nmar-1))+1 enddo irecord=0 do ik=1.mrec xmin=999999 do ij=1.mrcc if (irec(ij).lt.xmin) then xmin=irec(ij) imin=ij endif enddo irecord(k)=xmin irec(imin)=999999 enddo iale=2 IFEC(IIIIII)-------enddo irecord(nrec+1)=nmar if (nrec.eq.0) then do k=2,nmar finalC(i,j,k,1)=grupoC(ip,j,k,iale) enddo else ic=2

do ik=1,nrec+1 do k=ic,irecord(ik) finalC(i,j,k,1)=grupoC(ip,j,k,iale) enddo ic=irecord(ik)+1 if (iale.eq.1) then iale=2 else iale=1 endif 1 mutation lam=1. mut=pois(lam,x1) do ik=1,mut imut=in((u*mmar)+1 if (finalC(i,jimut,1)e_1) then finalC(i,jimut,1)=1 endif endif call unif(x1,u) if (u.lt.0.5) then iale=1 else iale=2 $\late-2 \\ endif \\ finalC(i,j,1,2)=grupoC(im,j,1,iale) \\ lam=1. \\ nrec=pois(lam,x1) \\ do ik=1, nrec \\ call unif(x1,u) \\ irrec(ik)=int(u*(nmar-1))+1 \\ .. \\ \label{eq:late-2}$ enddo irecord=0 do ik=1,nrec xmin=999999 do ij=1,nrec if (irec(ij).lt.xmin) then xmin=irec(ij) imin=ij endif enddo irecord(ik)=xmin irec(imin)=999999 enddo irecord cord(nrec+1)=nmar if (nrec.eq.0) then do k=2,nmar finalC(i,j,k,2)=grupoC(im,j,k,iale) enddo else ic=2 do ik=1,nrec+1 do k=ic,irecord(ik) finalC(i,j,k,2)=grupoC(im,j,k,iale) enddo ic=irecord(ik)+1 if (iale.eq.1) then iale=2 else iale=1 endif enddo endif <u>Imutation</u> lam=1. nnut=pois(lam,x1) do ik=1,nmut call unif(x1,u) imut=int(u*nmar)+1 if (finalC(i,jimut,2)=2 finalC(i,j, else finalC(i,j,imut,2)=1 enddo enddo enddo enddo [*frequencies of the causative mutations of trait 1* open(16,file='freq1-Ctxt') do i=1,ngenes p1C(ifec2(i,1),iefec2(i,2))=0 do k=1,nanim2 $\begin{array}{l} \text{oo} \ \kappa=1, namm2 \\ p | C(iefec1 (i,1), iefec1 (i,2)) = p | C(iefec1 (i,1), iefec1 (i,2)) + real(finalC(k, iefec1 (i,1), iefec1 (i,2), 1) + finalC(k, iefec1 (i,1), iefec1 (i,2), 2) - 2) \\ end \ do \end{array}$ $p1C(iefec1(i,1),iefec1(i,2))=p1C(iefec1(i,1),iefec1(i,2))+real(finalC(k,iefec1(i,1),iefec1(i,2),1)+finalC(k,iefec1(i,1),iefec1(i,2),2)-2; end do \\ p1C(iefec1(i,1),iefec1(i,2))=p1C(iefec1(i,1),iefec1(i,2))/(2*real(namim2)) \\ write(116*); i,(iefec1(i,1)+i)*mar+iefec1(i,2),p1C(iefec1(i,1),iefec1(i,2)) \\ end do \\ close(16) \\ \frac{ladditive variance trail 1}{Val=0} \\ varl C(fi)=2^p1C(iefec1(i,1),iefec1(i,2))*(1-p1C(iefec1(i,1),iefec1(i,2)))*efec1(i)**2 \\ varl C(fi)=2^p1C(iefec1(i,1),iefec1(i,2))*(1-p1C(iefec1(i,1),iefec1(i,2)))*efec1(i)**2 \\ varl C(fi)=2^p1C(iefec1(i,1),iefec1(i,2))*(1-p1C(iefec1(i,1),iefec1(i,2)))*efec1(i)**2 \\ varl C(fi)=2^p1C(iefec1(i,1),iefec1(i,2))*(1-p1C(iefec1(i,1),iefec1(i,2)))*efec1(i)**2 \\ varl C(fi)=2^p1C(iefec2(i,1),iefec1(i,2))*(1-p1C(iefec1(i,1),iefec1(i,2)))*efec1(i)**2 \\ varl C(fi)=2^p1C(iefec2(i,1),iefec1(i,2))*(1-p1C(iefec1(i,1),iefec2(i,2)))*efec1(i)**2 \\ varl C(fi)=2^p1C(iefec2(i,1),iefec2(i,2))=0 \\ do =1 , angens \\ p2C(iefec2(i,1),iefec2(i,2))=p2C(iefec2(i,1),iefec2(i,2))+real(finalC(k,iefec2(i,1),iefec2(i,2),1)+finalC(k,iefec2(i,1),iefec2(i,2),2)-2) \\ end do \\ p2C(iefec2(i,1),iefec2(i,2))=p2C(iefec2(i,1),iefec2(i,2))+real(finalC(k,iefec2(i,1),iefec2(i,2),1)+finalC(k,iefec2(i,1),iefec2(i,2),2)-2) \\ end do \\ do =1 , angin 2 \\ p2C(iefec2(i,1),iefec2(i,2))=p2C(iefec2(i,1),iefec2(i,2))+real(finalC(k,iefec2(i,1),iefec2(i,2),1)+finalC(k,iefec2(i,1),iefec2(i,2),2)-2) \\ end do \\ do =1 , angin 2 \\ p2C(iefec2(i,1),iefec2(i,2))=p2C(iefec2(i,1),iefec2(i,2))+real(finalC(k,iefec2(i,1),iefec2(i,2),1)+finalC(k,iefec2(i,1),iefec2(i,2),2)-2) \\ end do \\ do =1 , angin 2 \\ p2C(iefec2(i,1),iefec2(i,2))=p2C(iefec2(i,1),iefec2(i,2))+real(finalC(k,iefec2(i,1),iefec2(i,2),1)+finalC(k,iefec2(i,1),iefec2(i,2),2)-2) \\ end do \\ do =1 , angin 2 \\ p2C(iefec2(i,1),iefec2(i,2))=p2C(iefec2(i,1),iefec2(i,2))+real(finalC(k,iefec2(i,1),iefec2(i,2),2)) \\ end do \\ do =1 , angin 2 \\ do =1 , angi$ end do p2C(iefec2(i,1),iefec2(i,2))=p2C(iefec2(i,1),iefec2(i,2))/(2*eal(maim2)) write(16,*) i,(iefec2(i,1)-1)*nmar+iefec2(i,2)),p2C(iefec2(i,1),iefec2(i,2)) end do close(16) *Ladditive variance trait* 2 Va2=0 var2Cf(i)=0 do i=1,ngenes var2Cf(i)=2*p2C(iefec2(i,1),iefec2(i,2))*(1-p2C(iefec2(i,1),iefec2(i,2)))*efec2(i)**2

 $\begin{array}{l} Va2=Va2+var2Cf(i)\\ end do\\ print *, Variance trait 2 population C=',Va2\\ write(11,*)^{Variance trait 2 population C=',Va2\\ j:imulation of phenotypes\\ open(15,file='IDC.txt)\\ do := 1,namin2\\ yfen1C(i)=100\\ do := 1,ngenes\\ yfen1C(i)=yfen1C(i)+(finalC(i,iefec1(j,1),iefec1(j,2),1)+finalC(i,iefec1(j,1),iefec1(j,2),2)-3)*efec1(j)\\ end do\\ call normal(x1a)\\ yfen1C(i)=yfen1C(i)+u^*sqrt(Vr1)\\ write(15,*)yfen1C(i)+u^*sqrt(Vr1)\\ write(15,*)yfen1C(i)+(inalC(i,iefec2(j,1),iefec2(j,2),1)+finalC(i,iefec2(j,1),iefec2(j,2),2)-3)*efec1(j)\\ do = 1,namin2\\ yfen2C(i)=1000\\ do = 1,namin2\\ yfen2C(i)=1000 \end{array}$ Va2=Va2+var2Cf(i)
$$\label{eq:constraint} \begin{split} & \text{rel}(2,j) = \text{tot}(0) \\ & \text{d}_j = 1, \text{genes} \\ & \text{yfen}2(2(j) = \text{yfen}2C(i) + (\text{finalC}(i, \text{iefec}2(j, 1), \text{iefec}2(j, 2), 1) + \text{finalC}(i, \text{iefec}2(j, 2), 2) - 3)^* \text{efec}2(j) \\ & \text{end do} \\ & \text{call normal}(x1, u) \\ & \text{yfen}2(1) = \text{yfen}2C(i) + \text{iefen}2C(i) + \text{iefen}2C(i$$
end do close(16)
$$\label{eq:closed} \begin{split} & \text{close(16)} \\ \underline{l_write \ eenotypes} \\ & \text{open(14,file='gC.txt')} \\ & \text{do } j=1, namin2 \\ & \text{do } j=1, ncro \\ & \text{do } j=1, ncro \\ & \text{do } k=1, ncro \\ & \text{write(14,*)} \ i, ((j-1)*nmar+k), finalC(i,j,k,1), finalC(i,j,k,2) \\ & \text{end } \text{do } \\ & \text{end } \text{do } \\ & \text{end } \text{do } \\ & \text{close(14)} \\ & \underline{l_marker \ and \ gene \ frequencies} \\ & \overline{open(13,file='pC.txt')} \\ & \text{do } i=1, ncm \\ & \text{do } j=1, ncm \\ & \text{do } j=1, ncm \\ & \text{do } j=1, ncm \\ \end{split}$$
do j=1,ncro do j=1,nmar pC=0 do k=1,nanim2 r~−pc+real(finalC(k,i,j,1)+final-write(13,*) (i-1)*nmar+j.pC/(2*real(nanim2)) end do end do close(13) / write – pC=pC+real(finalC(k,i,j,1)+finalC(k,i,j,2)-2) close(13) <u>! write neutral marker genotypes</u> open(13,file='nC.txt') do i=1,nanim2 do j=1,ncro do k=1,nmar ar do g=1,ngenes if (j==iefec1(g,1),and.k==iefec1(g,2)) then goto 27 end if if (j==iefec2(g,1),and.k==iefec2(g,2)) then goto 27 end if end io end do write(13,*) i,(j-1)*nmar+k,(finalC(i,j,k,1)+finalC(i,j,k,2)-3) 27 end do end do end do end do close(13) *l_write neutral marker genotypes in format* open(20,file='nC.txt') open(13,file='genotC.txt') do i=1, nanim2 marca(j)=0 do j=1, neu read(20,*) ip.im.iv marca(j)=iy+1 end do write(13,'(i4,1x,30000i1)') i,(marca(j),j=1,neu) end do close(13) close(20) *l_true breeding values for traits 1 & 2* open(14,file='bv2C.txt') open(15,file='bv1C.txt') do j=1, nanim2 bv2C(j)=0 do j=1, ngenes close(13)
$$\label{eq:constraint} \begin{split} bv2C(j)=0 & \\ do & j=1, agenes \\ bv2C(i)=tv2C(i)+(finalC(i,iefec2(j,1),iefec2(j,2),1)+finalC(i,iefec2(j,1),iefec2(j,2),2)-3)*efec2(j) \\ end do & \\ vrite(14,*), ibv2C(i) & \\ end do & \\ close(14) & \\ do & i=1, namin2 & \\ bv1C(j)=0 & \\ do & j=1, agenes & \\ bv1C(j)=bv1C(j)+(finalC(i,iefec1(j,1),iefec1(j,2),1)+finalC(i,iefec1(j,1),iefec1(j,2),2)-3)*efec1(j) \\ end do & \\ vrite(15,*), ibv1C(i) & \\ end do & \\ close(15) & \\ \end{split}$$
<u>! Population D under random mating</u> marnD=marn <u>! random mating for n generations</u> do ijk=1,ngen2 do i=1,nanim call unif(x1,u) ip=int(u*nmac)+1 call unif(x1,u) im=int(u*nhem)+nmac+1 <u>father's chromosome</u> do j=1,ncro call unif(x1,u) if (u.lt.0.5) then iale=1 else iale=2 endif grupoD(i,j,1,1)=marnD(ip,j,1,iale) lam=1. nrec=pois(lam,x1) do ik=1,nrec

call unif(x1,u) irec(ik)=int(u*(mmar-1))+1 enddo irecord=0 do ik=1,mrc xmin=999999 do ij=1,mrc if (irec(ij).1t.xmin) then xmin=irec(ij) endif enddo irecord(ik)=xmin irecord(ik)=xmin irecord(mrc+1)=nmar if (mrc.eq.0) then do k=2,mmar grupoD(i,j,k,1)=marnD(ip,j,k,iale) enddo else irec-2 ене ic=2 do ik=1,nrec+1 do k=ic,irecord(ik) grupoГ grupoD(i,j,k,1)=marnD(ip,j,k,iale) enddo ic=irecord(ik)+1 if (iale.eq.1) then iale=2 else .se endif enddo endif iale=1 ! mutation lam=1. nmut=pois(lam,x1) do ik=1,nmut call unif(x1,u) imut=int(u*nmar)+1 if (grupoD(i,jimut,1)eq.1) then grupoD(i,jimut,1)=2 else grupoD(i,jimut,1)=1 eise grupoD(i,j,imut,1)=1 endif enddo enddo do j=1,ncro call unif(x1,u) if (u.lt.0.5) then iale=1 else iale=2 endif grupDD(i,j.1,2)=marnD(im,j.1,iale) lam=1. nrcc=pois(lam,x1) do ik=1,nrec call unif(x1,u) irec(ik)=int(u*(nmar-1))+1 enddo irecord=0 do ik=1,nrec xmin=999999 do ij=1,nrec xmin=999999 do ij=1,nrec if (irec(ij).1t.xmin) then xmin=ije(ij) imin=ij endif enddo ! mother's chromosome endif endif endif irecord(ik)=xmin irecord(arec+1)=mmar if (arec.eq.0) then do k=2,mmar grupoD(i,j,k,2)=marnD(im,j,k,iale) enddo else ic=2 do ik=1,arec+1 do k=ic,irecord(ik) grupoD(i,j,k,2)=marnD(im,j,k,iale) enddo ic=irecord(ik)+1 if (ale.eq.1) then iale=1 endif enddo endif <u>Immtaion</u> Iam=1. nmut=pois(lam,x1) do ik=1,nmut call unif(x1,u) if(grupoD(i,jimut2),eq.1) then grupoD(i,jimut2)=2 else grupoD(i,jimut2)=1 endif enddo $\label{eq:constraints} \begin{array}{l} \underline{l:Final population D} \\ doi=1, nanim2 \\ call unif(xl,u) \\ ip=int(u^mmac)+1 \\ call unif(xl,u) \\ im=int(u^mhem)+nmac+1 \\ \underline{l:father's chromosome} \\ doi=1, ncro \\ call unif(xl,u) \\ if(u,lt.0.5) then \end{array}$

iale=1 else iale=2 endif finalD(i,j,1,1)=grupoD(ip,j,1,iale) nnaiD(1,1,1,1)=grupoL lam=1. nrec=pois(lam,x1) do ik=1,nrec call unif(x1,u) irec(ik)=int(u*(nmar-1))+1 enddo lo irecord=0 do ik=1,nrec xmin=999999 999 do ij=1,nrec if (irec(ij).lt.xmin) then xmin=irec(ij) imin=ij endif enddo lo irecord(ik)=xmin irec(imin)=999999 enddo irecord(nrec+1)=nmar if (nrec.eq.0) then do k=2,nmar finalD(i,j,k,1)=grupoD(ip,j,k,iale) enddo else $\label{eq:ansatz} \begin{array}{c} \mbox{endif} \\ \mbox{immatched} \\ \mbox{lam=1}, \\ \mbox{nmut=pois(lam,x1)} \\ \mbox{do ik=1,nmut} \\ \mbox{call unif(x1,u)} \\ \mbox{immut=in(u^nnmr)+1} \\ \mbox{if(nn1D(i,j,imut,1),eq.1)} \mbox{then} \\ \mbox{if(nn1D(i,j,imut,1)=2)} \\ \mbox{else} \\ \mbox{else} \\ \mbox{$$^{-r}$na1D(i,j,imut,1)=1$} \end{array}$ enddo <u>enddo</u> <u>i mother's chromosome</u> do j=1,ncro call unif(x1,u) if (u.lt.0.5) then iale=1 else iale=2 endif finalD(i,j,1,2)=grupoD(im,j,1,iale) is(lam.x1) do ik=1,mec call unif(x1,u) irec(ik)=int(u*(mmar-1))+1 enddo irecord=0 do ik=1,mec xmin=999999 do ij=1,mec if (irec(ij)1,txmin) then xmin=irec(ij) imin=ij endif enddo irecord(ik)=xmin irec(imin)=999999 enddo irecord(ik)=xmin irec(imin)=mmar if (inec e.q.0) then do k=2,mmar finatD(i,j,k,2)=grupoD(im,j,k,iale) enddo else i=2 lam=1. nrec=pois(lam,x1) nute enddo else ic=2 do ik=1,nrec+1 do k=ic,irecord(ik) finalD(i,j,k,2)=grupoD(im,j,k,iale) enddo ic=irecord(ik)+1 if (iale.eq.1) then iale=2 else iale=1 endif endif ! mutation lam=1. lam=1. nmut=pois(lam,x1) do ik=1,nmut call unif(x1,u) imut=int(u*nmar)+1 if (finalD(i,j,imut,2)=eq.1) then finalDo(i,j,imut,2)=2 else else finalD(i,j,imut,2)=1 endif enddo $\label{eq:enddo} enddo \\ enddo \\ lfrequencies of the causative mutations of trait 1 \\ open(16,file="freq1-D.txt") \\ do := 1, agenes \\ p1D(iefec1(i,1),iefec1(i,2))=0 \\ do k=1, namin2 \\ p1D(iefec1(i,1),iefec1(i,2))=p1D(iefec1(i,1),iefec1(i,2))+real(finalD(k,iefec1(i,1),iefec1(i,2),1)+finalD(k,iefec1(i,1),iefec1(i,2),2)-2) \\$

end do $p1D(iefec1(i,1),iefec1(i,2)) = p1D(iefec1(i,1),iefec1(i,2))/(2*real(nanim2)) \\ write(16,*)i,((iefec1(i,1)-1)*nmar+iefec1(i,2)), p1D(iefec1(i,1),iefec1(i,2)) \\ (1,2), (1$ end do close(16) <u>! additive variance trait 1</u> Val=0 var1Df(i)=0 $\label{eq:starting} \begin{array}{l} Vari Df(i)=0 \\ do & = 1, agenes \\ vari Df(i)=2^p D(iefec1(i,1),iefec1(i,2))^*(1-p1D(iefec1(i,1),iefec1(i,2)))^*efec1(i)^{**2} \\ Val = Val + vari Df(i) \\ end do \\ print ^*, Variance trait 1 population D=',Val \\ \underline{frequencies} \ fcausative mutations for trait 2 \\ open(16,file= {\rm freq2-D}\,xt) \\ do & := 1, agenes \\ p2D(iefec2(i,1),iefec2(i,2))=0 \\ do & := 1, anim2 \\ p2D(iefec2(i,1),iefec2(i,2))=p2D(iefec2(i,1),iefec2(i,2))+real(finalD(kiefec2(i,1),iefec2(i,2),1)+finalD(kiefec2(i,1),iefec2(i,2),2)-2) \\ end do \\ \hline vari = 0 \\ context \\ do & := 1, anim2 \\ p2D(iefec2(i,1),iefec2(i,2))=p2D(iefec2(i,1),iefec2(i,2))/(2^*real(nanim2)) \end{array}$ $\label{eq:product} \begin{array}{l} p_{2D}(\text{recc}_{(i, 2)}) = p_{2D}(\text{recc}_{(i, 1)}) + (\text{real}(\text{name}), \text{set}) \\ p_{2D}(\text{refc}_{2}(i, 1), \text{iefec}_{2}(i, 2)) = p_{2D}(\text{refec}_{2}(i, 1), \text{iefec}_{2}(i, 2)) \\ \text{write}(16, *) \ i, ((\text{refec}_{(i, 1)-1})*\text{nmar} + \text{refec}_{2}(i, 2)), p_{2D}(\text{refec}_{2}(i, 1), \text{iefec}_{2}(i, 2)) \\ \text{end do} \end{array}$ close(16) <u>! additive variance trait 2</u> Va2=0 $\begin{array}{l} v_{42=0} \\ w_{a2}D(i)=0 \\ do i=1, ngenes \\ w_{a2}D(i)=2^*p_2D(iefec2(i,1),iefec2(i,2))^*(1-p_2D(iefec2(i,1),iefec2(i,2)))^*efec2(i)^{**2} \\ v_{a2}=v_{a2}+v_{a2}2D(i) \\ w_{a2}=v_{a2}+v_{a2}2D(i) \end{array}$ Va2=Va2+Var2Df(i) end do print *, 'Variance trait 2 population D=',Va2 write(11,*) 'Variance trait 2 population D=',Va2 /<u>simulation of phenorypes</u> open(15,file=f1D.txt') open(16,file=f2D.txt) do i=1,nanim2 yfen1D(i)=100 do i=1 neenes yfen1D(i)=yfen1D(i)+u*sqrt(Vr1) write(15,*) yfen1D(i),i,1 end do close(15) do i=1,nanim2 vfen2D(i)=1000
$$\label{eq:linear} \begin{split} & = 1000 \\ & d_j = 1, \text{genes} \\ & y fen 2D(i) = y fen 2D(i) + (finalD(i, iefec2(j, 1), iefec2(j, 2), 1) + finalD(i, iefec2(j, 1), iefec2(j, 2), 2) \cdot 3)^* efec2(j) \\ & end \ do \\ & call \ normal(x1, u) \end{split}$$
call normal(x1,u) yfen2D(i)=yfen2D(i)+yfen2D(i),i,1 end do close(16) *_write* (genotypes open(14,ifle=gB).txt) do i=1,nanim2 do i=1,nanim2 do j=1,ncro do k=1,nmar write(14,*) i,((j-1)*nmar+k),finalD(i,j,k,1),finalD(i,j,k,2) end do end do end do close(14) <u>*! marker and gene frequencies*</u> open(13,file='pD.txt') do i=1,ncro do j=1,nmar pD=0 do k=1,nanim2 pD=pD+real(finalD(k,i,j,1)+finalD(k,i,j,2)-2) pD=pD+real(finalD(k,i,j,1)+finalD end do write(13,*) (i-1)*nmar+j,pD/(2*real(nanim2)) end do close(13) <u>! write neutral marker genotypes</u> open(13,file='nD.txt') do i=1,nanim2 do i=1,naro do j=1,ncro do k=1,nmar do g=1,ngenes if (j==iefec1(g,1).and.k==iefec1(g,2)) then goto 28 end if if (j==iefec2(g,1).and.k==iefec2(g,2)) then goto 28 end if end do write(13,*) i,(j-1)*nmar+k,(finalD(i,j,k,1)+finalD(i,j,k,2)-3) end do write(13,'(i4,1x,30000i1)') i,(marca(j),j=1,neu) end do close(13) end do close(14)

 $\begin{array}{l} do \ i=1, anim2 \\ bv 1D(j)=0 \\ do \ j=1, ngenes \\ bv 1D(i)=bv 1D(i)+(finalD(i, iefec 1(j,1), iefec 1(j,2), 1)+finalD(i, iefec 1(j,1), iefec 1(j,2), 2)-3)^{*}efec 1(j) \end{array}$ end do write(15,*) i,bv1D(i) end do close(15)
$$\label{eq:constraint} \begin{split} \underline{l \ Population \ A+B} \\ open(12, file='f1 \ AB, txt') \\ open(13, file='f2 \ AB, txt') \\ do i= 1, anim2 \\ call unif(x1, u) \\ ip=int(nanim2^*u)+1 \\ call unif(x1, u) \\ if (u, lt. 0.5) \\ then \\ yfen 1 \ AB(i)=yfen 1 \ A(ip) \\ yfen 2 \ AB(i)=yfen 2 \ A(ip) \\ do j=1, ncro \end{split}$$
do j=1,ncro do k=1,nmar finalAB(i,j,k,1)=finalA(ip,j,k,1) finalAB(i,j,k,2)=finalA(ip,j,k,2) end do end do else yfen1AB(i)=yfen1B(ip) yfen17AB(i)=yfen1B(ip) yfen2AB(i)=yfen2B(ip) do j=1,ncro do k=1,nmar do k=1,nmar do k=1,nmar finalAB(ij,j,k,1)=finalB(ip,j,k,1) finalAB(ij,j,k,2)=finalB(ip,j,k,2) end do end if write(1,2*) yfen1AB(i),j,1 write(1,3*) yfen2AB(i),i,1 end do close(12) close(13) <u>L</u>frequencies of the causative mutati $\label{eq:close} close (13) \\ \frac{\int frequencies of the causative mutations of trait I}{\int requencies of the causative mutations of trait I} \\ open (16,file="freq1-AB.txt") \\ do i=1, ageneric distribution of trait I \\ p1AB(iefec1(i,1),iefec1(i,2))=0 \\ do k=1, namin2 \\ p1AB(iefec1(i,1),iefec1(i,2))=p1AB(iefec1(i,1),iefec1(i,2))+real(finalAB(k,iefec1(i,1),iefec1(i,2),1)+finalAB(k,iefec1(i,1),iefec1(i,2),2)-2) \\ end do \\ p1AB(iefec1(i,1),iefec1(i,2))=p1AB(iefec1(i,1),iefec1(i,2))/(2*real(namin2)) \\ write(16,*) i,(iefec1(i,1)-1)*nmar+iefec1(i,2),p1AB(iefec1(i,1),iefec1(i,2)) \\ end do \\ end end \\ end \\$ ena do close(16) <u>! additive variance of trait 1</u> Va1=0 Val=0 var1ABf(i)=0 do i=1,ngenes var1ABf(i)=2*p1AB(iefec1(i,1),iefec1(i,2))*(1- p1AB(iefec1(i,1),iefec1(i,2)))*efec1(i)**2 Val=Val+var1ABf(i) end do primt *, Variance trait 1 population A+B=:,Val write(11,*) Variance trait 1 population A+B=:,Val //requencies of the causative mutations of trait 2 open(16,file=freq2-AB_txt') do i=1,ngenes p2AB(iefec2(i,1),iefec2(i,2))=0 do k=1,namin2 V2AB(iefec2(i,1),iefec2(i,2))=D do k=1,nanim2 p2AB(iefec2(i,1),iefec2(i,2))=p2AB(iefec2(i,1),iefec2(i,2))+real(finalAB(k,iefec2(i,1),iefec2(i,2),1)+finalAB(k,iefec2(i,1),iefec2(i,2),2)-2) end do p2AB(iefec2(i,1),iefec2(i,2))=p2AB(iefec2(i,1),iefec2(i,2))/(2*real(nanim2)) write(16,*) i,((iefec2(i,1)-1)*nmar+iefec2(i,2)),p2AB(iefec2(i,1),iefec2(i,2)) = -tend do close(16) close(16) <u>Ladditive variance trait 2</u> Va2=0 do i=1,ngenes var2ABf(i)=2*p2AB(iefec2(i,1),iefec2(i,2))*(1-p2AB(iefec2(i,1),iefec2(i,2)))*efec2(i)**2 Va2=Va2+var2ABf(i) end do print *, Variance trait 2 population A+B=',Va2 write(11,*) Variance trait 2 population A+B=',Va2 write(11,*) Variance trait 2 population A+B=',Va2 <u>verite genotypes</u> open(14,file='gAB.txt') do i=1,narin do j=1,ncro 0 i=1,mam... do j=1,ncro do k=1,nmar write(14,*) i,((j-1)*nmar+k),finalAB(i,j,k,1),finalAB(i,j,k,2) end do end do end do end do close(14) <u>! marker and gene frequencies</u> open(13,file='pAB.txt') do i=1,ncro do i=1,nero do j=1,nmar pAB=0 do k=1,namin2 pAB=pAB+real(finalAB(k,i,j,1)+finalAB(k,i,j,2)-2) end do write(13,*) (i-1)*nmar+j,pAB/(2*real(nanim2)) end do close(13) *y write neutral marker genotypes* <u>*! write neutral marker genotypes*</u> open(13,file='nAB.txt') do i=1,nanim2 do j=1,ncro do k=1,nmar do g=1,ngenesif (j==iefec1(g,1).and.k==iefec1(g,2)) then if (j===tecl(g,1).and.k==tefcl(g,2)) then
goto 29
end if
if (j===tefc2(g,1).and.k==tefcc2(g,2)) then
goto 29
end if end if end do write(13,*) i,(j-1)*nmar+k,(finalAB(i,j,k,1)+finalAB(i,j,k,2)-3) 29 end do end do close(13) *(write neutral marker genotypes in format* open(20,file='nAB.txt')

open(13,file='genotAB.txt') do i=1,nanim2 marca(j)=0 do j=1,neu read(20,*) ip,im,iv marca(j)=iv+1 end do write(13,'(i4,1x,30000i1)') i,(marca(j),j=1,neu) end do close(13) close(20) <u>IPopulation A+C</u> open(12,file='f1AC.txt') open(13,file='f2AC.txt') do i=1,nanim2 call unif(x1,u) ip=int(nanim2*u)+1 call unif(x1,u) if (u,lt0,5) then if (u.lt.0.5) then yfen1AC(i)=yfen1A(ip) yfen2AC(i)=yfen2A(ip) do j=1,ncro do k=1,nmar finalAC(i,j,k,1)=finalA(ip,j,k,1) finalAC(i,j,k,2)=finalA(ip,j,k,2) end do end do else else final action of the second secon end do close(12) close(13) <u>! frequencies of the causative mutations for trait 1</u> open(16,file='freq1-AC.txt') open (16,111e=treq1-AC.txt) doi=1,ngenes p1AC(iefec1(i,1),iefec1(i,2))=0 do k=1,nanim2 p1AC(iefec1(i,1),iefec1(i,2))=p1AC(iefec1(i,1),iefec1(i,2))+real(finalAC(k,iefec1(i,1),iefec1(i,2),1)+finalAC(k,iefec1(i,1),iefec1(i,2),2)-2) end do end do pl AC(iefec1(i,1),iefec1(i,2))=pl AC(iefec1(i,1),iefec1(i,2))/(2*real(nanim2)) write(16,*) i,((iefec1(i,1)-1)*nmar+iefec1(i,2)),pl AC(iefec1(i,1),iefec1(i,2)) end do $\begin{array}{l} ao \; k=1,nanim2 \\ p2AC(iefec2(i,1),iefec2(i,2))=p2AC(iefec2(i,1),iefec2(i,2))+real(finalAC(k,iefec2(i,1),iefec2(i,2),1)+finalAC(k,iefec2(i,1),iefec2(i,2),2)-2) \\ end \; do \end{array}$ end do p2AC(iefec2(i,1),iefec2(i,2))=p2AC(iefec2(i,1),iefec2(i,2))/(2*real(nanim2)) write(16,*) i,((iefec2(i,1)-1)*nmar+iefec2(i,2)),p2AC(iefec2(i,1),iefec2(i,2)) $\label{eq:constraint} \begin{array}{l} \mbox{write}(16,*)\ i_{i}(iefec2(i,1).i)^{9}nmar+iefec2(i,2))\ p2AC(iefec2(i,1).iefec2(i,2))\ end do \\ \ close(16)\ label{eq:close}(16)\ label{eq:close}(16)\ label{eq:close}(16)\ va2AC(iefec2(i,1).iefec2(i,2))\ va2AC(iefec2(i,1).iefec2(i,2))\ va2ACf(i)=0\ do \\ \ va2ACf(i)=0\ va2ACf(i)=2\ va2AC(iefec2(i,1).iefec2(i,2))^{*}(1-p2AC(iefec2(i,1).iefec2(i,2)))^{*}efec2(i)^{**}2\ va2ACf(i)=2\ va2$ <u>! write genotypes</u> open(14,file='gAC.txt') do i=1,nanim2 do j=1,ncro do k=1,nmar write(14,*) i,((j-1)*nmar+k),finalAC(i,j,k,1),finalAC(i,j,k,2) end do end do end do ento u. end do close(14) <u>!marker and gene frequencies</u> open(13.file=pAC.txt) do i=1.ncno do j=1.nmar pAC=0 do k=1.nanim2 pAC=pAC+real(finalAC(k.i,j.1)+finalAC(k.i,j.2)-2) _*An -~~i nAC/(2*real(nanim2)) pAC=pAC+real(finalAC(k,i,j,1)+fin end do write(13,*) (i-1)*nmar+j,pAC/(2*real(nanim2)) end do close(13) *! write neutral marker genotypes* open(13,file='nAC.txt') do i=1,nanim2 do j=1,ncro do k=1,nmar ar do g=1,ngenes if (j=-iefec1(g,1),and.k==iefec1(g,2)) then goto 30 end if if (j==iefec2(g,1),and.k==iefec2(g,2)) then goto 30 end if end do write(13,*) i,(j-1)*nmar+k,(finalAC(i,j,k,1)+finalAC(i,j,k,2)-3) 30 end do end do

end do close(13) <u>! write neutral marker genotypes in format</u> open(20,file='nAC.txt') open(13,file='genotAC.txt') do i=1,namin2 $\label{eq:constraint} \begin{array}{l} do i=1,nanim2\\ marca(j)=0\\ do j=1,neu\\ marca(j)=iv+1\\ end \ do\\ write(13,(c4,1x,30000i1))\ i,(marca(j),j=1,neu)\\ end \ do\\ elose(13)\\ elose(20) \end{array}$ <u>Population A+D</u> open(12,file='fiAD.txt') open(13,file='f2AD.txt') do i=1,nanim2 call unif(x1,u) ip=int(nanim2*u)+1 call unif(x1,u) if (u.1t.0.5) then of an ID()=vfen1A() yfen1 AD(i)=yfen1 A(ip) yfen2 AD(i)=yfen2 A(ip) do j=1,ncro do k=1,nmar final AD(i,j,k,1)=final A(ip,j,k,1) final AD(i,j,k,2)=final A(ip,j,k,2) end do end do end do else yfen1AD(i)=yfen1D(ip) yfen2AD(i)=yfen2D(ip) do j=1,ncro do k=1,nmar final AD(i,j,k,1)=finalD(ip,j,k,1) final AD(i,j,k,2)=finalD(ip,j,k,2) end do end do end if write(12,*) yfen1AD(i),i,1 write(13,*) yfen2AD(i),i,1 end do close(12) close(13) $\label{eq:interm} \begin{array}{l} close(13) \\ l frequencies of the causative mutations for trait 1 \\ open(16,file=freq1-AD,txt) \\ do i=1, ngenes \\ plAD(iefec1(i,1),iefec1(i,2))=0 \\ db k=1,namin2 \\ plAD(iefec1(i,1),iefec1(i,2))=plAD(iefec1(i,1),iefec1(i,2))+real(finalAD(k,iefec1(i,1),iefec1(i,2),1)+finalAD(k,iefec1(i,1),iefec1(i,2),2)-2) \\ end do \\ pl AD(iefec1(i,1),iefec1(i,2))=pl AD(iefec1(i,1),iefec1(i,2))/(2*real(namim2)) \\ write(16,*),i((iefec1(i,1)-1)*nmar+iefec1(i,2)),pl AD(iefec1(i,1),iefec1(i,2)) \\ end do \\ close(16) \\ l additive variance of trait 1 \\ Val=0 \\ Val=0 \\ Varl AD(i)=0 \\ Varl AD(i)=0 \\ \end{array}$ $\label{eq:constraint} \begin{array}{l} \mbox{var}(i,j) = 0 \\ \mbox{o} = 1, \mbox{gens} \\ \mbox{var} | AD(i) = 2^* p | AD(iefec1(i,1), iefec1(i,2))^* (1-p | AD(iefec1(i,1), iefec1(i,2)))^* efec1(i)^{**} \\ \mbox{Val} = Val + var | ADf(i) \end{array}$ Val=Val+Var1AD(t) end do print *, Variance trait 1 population A+D=',Val write(11,*) 'Variance trait 1 population A+D=',Val *[frequencies of the causative mutations for trair_*2 open(16.file='freq2-AD.txt') do i=1,ngenes p2AD(refece2(i,1));efec2(i,2))=0 do te-1 write? Provide c2(i,1),retec2(i,2),-0-do k=1,namin2 p2AD(refec2(i,1),iefec2(i,2))=p2AD(refec2(i,1),iefec2(i,2))+real(finalAD(k,iefec2(i,1),iefec2(i,2),1)+finalAD(k,iefec2(i,1),iefec2(i,2),2)-2) end do end do $\label{eq:production} p2 AD(iefec2(i,1),iefec2(i,2))=p2 AD(iefec2(i,1),iefec2(i,2))/(2*real(nanim2))) write(116,*) i.((iefec2(i,1))*nmar+iefec2(i,2)),p2 AD(iefec2(i,1),iefec2(i,2))) end do <math display="block">\label{eq:production} eof trait 2 Va2=0 V$ Val=val+val2AD(1) end do print *, 'Variance trait 2 population A+D=',Va2 write(11.*) 'Variance trait 2 population A+D=',Va2 <u>/write genorpes</u> open(14,file='gAD.txt') do i=1, nanim2 do j=1,ncro do k=1,nmar do k= :,.. write(:-, end do end do close(14) <u>Imarker and gene frequencies</u> open(13,file=pAD.txt) do i=1,ncn do k=1,nmar pAD=0 do k=1,nanim2 pAD=pAD+real(finalAD(k,i,j,1)+finalAD(k,i,j,2)-2) ~~4 do write(14,*) i,((j-1)*nmar+k),finalAD(i,j,k,1),finalAD(i,j,k,2) close(13) close(13) <u>/_write neutral marker genotypes</u> open(13,file='nAD.txt') do i=1.namin2 do j=1.ncro do k=1.nmar do g=1.ngenes if (j==iefc2(g,1).and.k==iefcc1(g,2)) then goto 31 end if if (j==iefcc2(g,1).and.k==iefcc2(g,2)) then goto 31 end if

end do write(13,*) i,(j-1)*nmar+k,(finalAD(i,j,k,1)+finalAD(i,j,k,2)-3) 31 end do end do end do close(13) close(13) *lwrite neutral marker genotypes in format* open(20,file='nAD.txt') open(13,file='genotAD.txt') do i=1,nami marca(j)=-0 do j=1,neu read(20,°) ipi,m,iv marca(j)=iv+1 end do write(13,(i4,1x,30000i1)) i,(marca(j),j=1,neu) end do close(13) close(20) <u>! Population B+C</u> open(12,file='f1BC.txt') open(13,file='f2BC.txt') do i=1,nanim2 call unif(x1,u) ip=int(nanim2*u)+1 call unif(x1,u) call unif(x1,u) if (u.lt.0.5) then yfen1BC(i)=yfen1B(ip) yfen2BC(i)=yfen2B(ip) do j=1,ncro do k=1,nmar finalBC(i,j,k,1)=finalB(ip,j,k,1) finalBC(i,j,k,2)=finalB(ip,j,k,2) end do end do else yfen1BC(i)=yfen1C(ip) yfen2BC(i)=yfen2C(ip) do j=1,ncro do k=1,nmar finalBC(i,j,k,1)=finalC(ip,j,k,1) finalBC(i,j,k,2)=finalC(ip,j,k,2) end if write(12,*) yfen1BC(i),i,1 write(13,*) yfen2BC(i),i,1 end do close(12) close(13) $close(13) \\ \frac{frequencies of the causative mutations for trait 1}{open(16,file=Treq1-BC.txt)} \\ do i=1,ngenes \\ plBC(iefec1(i,1),iefec1(i,2))=0 \\ do k=1,nanim2 \\ plBC(iefec1(i,1),iefec1(i,2))=plBC(iefec1(i,1),iefec1(i,2))+real(finalBC(k,iefec1(i,1),iefec1(i,2),1)+finalBC(k,iefec1(i,1),iefec1(i,2),2)-2) \\ end do \\ plBC(iefec1(i,1),iefec1(i,2))=plBC(iefec1(i,1),iefec1(i,2))/(2*real(nanim2)) \\ write(16*) i,((iefec1(i,1)-1)*nmar+iefec1(i,2)),plBC(iefec1(i,1),iefec1(i,2)) \\ eldset (16) \\ close(16) \\ ladditive variance of trait 1 \\ \end{tabular}$ $\begin{array}{l} close(16) \\ \hline Val=Vallity evriance of trait 1 \\ \hline Val=0 \\ varl BCf(i)=0 \\ do i=1, ngenes \\ varl BCf(i)=2^*p1BC(iefec1(i,1),iefec1(i,2))^*(1-p1BC(iefec1(i,1),iefec1(i,2)))^*efec1(i)^{**}2 \\ Val=Val+varl BCf(i) \end{array}$ Vac2bc1(r)=2 do i=1,ngenes var2BCf(i)=2*p2BC(iefec2(i,1),iefec2(i,2))*(1-p2BC(iefec2(i,1),iefec2(i,2)))*efec2(i)**2 Va2=Va2+var2BCf(i) Va2=Va2=Va2+Va2EV_1(r) end do print *, Variance trait 2 population B+C=',Va2 write(11.*) Variance trait 2 population B+C=',Va2 <u>write(acotype</u> open(14.file='gBC.txt) do i=1,namin2 do j=1,ncro do k=1,nmar write(14.*) i,((j-1)*nmar+k),finalF write(14,*) i,((j-1)*nmar+k),finalBC(i,j,k,1),finalBC(i,j,k,2) end do end do end do close(14) <u>! marker and gene frequencies</u> open(13,file='pBC.txt') do i=1,ncro do j=1,nmar pBC=0 do k=1,nanim2 pBC=pBC+real(finalBC(k,i,j,1)+finalBC(k,i,j,2)-2) write(13,*) (i-1)*nmar+j,pBC/(2*real(nanim2)) end do end do end do close(13) <u>! write neutral marker genotypes</u> open(13,file='nBC.txt') do i=1,naim2 do j=1,ncro do k=1,nmar do g=1,ngenes rf do g=1,ngenes if (j==iefec1(g,1).and.k==iefec1(g,2)) then goto 32

end if if (j==iefec2(g,1).and.k==iefec2(g,2)) then goto 32 end if end do write(13,*) i,(j-1)*nmar+k,(fina 32 end do end do close(13) *lwrite neutral markers genotypes in format* open(20,file='nBC.txt') open(13,file='genotBC.txt') do i=1,nami marca(j)=0 do j=1,neu read(20,*) ipi,mi,v marca(j)=iv+1 end do write(13,i(i,1x,30000i1)) i,(marca(j),j=1,neu) end do write(13,*) i,(j-1)*nmar+k,(finalBC(i,j,k,1)+finalBC(i,j,k,2)-3) end do close(13) close(20) <u>! Population B+D</u> open(12,file='f1BD.txt') open(13,file='f2BD.txt') do i=1,nanim2 call unif(x1,u) call unif(x1,u) ip=int(nanim2*u)+1 call unif(x1,u) if (u.l.t.0.5) then yfen1BD(i)=yfen1B(ip) yfen2BD(i)=yfen2B(ip) do j=1,ncro do k=1,nmar finalBD(i) finalBD(i,j,k,1)=finalB(ip,j,k,1) finalBD(i,j,k,2)=finalB(ip,j,k,2) finalBD(i) end do else yfen1BD(i)=yfen1D(ip) yfen2BD(i)=yfen2D(ip) do j=1,ncro do k=1,nmar finalBD(i,j,k,1)=finalD(ip,j,k,1) finalBD(i,j,k,2)=finalD(ip,j,k,2) end do end do end if end if write(12,*) yfen1BD(i),i,1 write(13,*) yfen2BD(i),i,1 end do elose(12) elose(13) <u>/frequencies of the causative mutations of trait 1</u> open(16,file=Treq1-BD.txt) do i=1,ngenes p1BD(iefec1(i,1),iefec1(i,2))=0 do k=1,naim.2 p1BD(iefec1(i,1),iefec1(i,2))=p1BD(iefec1(i,1),iefec1(i,2))+real(finalBD(k,iefec1(i,1),iefec1(i,2),1)+finalBD(k,iefec1(i,1),iefec1(i,2),2)-2) end do end do close(16) <u>l additive variance of trait 1</u> Val=0 var1BDf(i)=0 var1BDf(i)=2*p1BD(iefec1(i,1),iefec1(i,2))*(1-p1BD(iefec1(i,1),iefec1(i,2)))*efec1(i)**2 Val=Val+var1BDf(i) $\label{eq:second} \begin{array}{l} \mbox{Val} = Val = Val = Dottee C(1,1), lefter (1,2))^{n}(-p) = Dottee C(1,1), lefter (1,2))^{n} effect (1)^{n+2} \\ \mbox{Val} = Val = Val = Val \\ \mbox{Val} = Val = Val \\ \mbox{Var} = Val \\ \mbox{Var} = Causative mutations of trait 2 \\ \mbox{open} (16, file = freq2 = B) X(1) \\ \mbox{do is } = I, agenes \\ \mbox{open} (16, file = freq2 = B) X(1) \\ \mbox{do is } = I, agenes \\ \mbox{open} (16, file = freq2 = B) X(1) \\ \mbox{do is } = I, agenes \\ \mbox{p2BD}(iefec2(i, 1), iefec2(i, 2)) = p 2BD(iefec2(i, 1), iefec2(i, 2)) + real(finalBD(k, iefec2(i, 1), iefec2(i, 2), 1) + finalBD(k, iefec2(i, 1), iefec2(i, 2), 2) \\ \mbox{p2BD}(iefec2(i, 1), iefec2(i, 2)) = p 2BD(iefec2(i, 1), iefec2(i, 2)) / (2^{n}real(nanim2)) \\ \mbox{write} (16, file) \\ \mbox{if} (16, file) \\ \mbo$ Va2=0 var2BDf(i)=0 do i=1.ngenes var2BDf(i)=2*p2BD(iefec2(i,1),iefec2(i,2))*(1-p2BD(iefec2(i,1),iefec2(i,2)))*efec2(i)**2 Va2=Va2+var2BDf(i) Va2=Va2+var2BDf(i) end do print *, 'Variance trait 2 population B+D=',Va2 *write* (11.*)' Variance trait 2 population B+D=',Va2 *jwrite* genotypes open(14.file='gBD1xt') do i=1,nanim2 do j=1,ncro do k=1,nmar write(14.*) i.((i-1)*nmar+k).finalB write(14,*) i,((j-1)*nmar+k),finalBD(i,j,k,1),finalBD(i,j,k,2) end do end do end do close(14) l marker and gene frequencies open(13,file='pBD.txt') do i=1,ncro do j=1,nmar pBD=0 do k=1,nanim2 pBD=pBD+real(finalBD(k,i,j,1)+finalBD(k,i,j,2)-2) end do end do write(13,*) (i-1)*nmar+j.pBD/(2*real(nanim2)) end do end do close(13) <u>/write neutral marker genotypes</u> open(13,file='nBD.txt') do i=1,nanim2 do j=1,ncro

do k=1,nmar do g=1,ngenes if (j==iefec1(g,1).and.k==iefec1(g,2)) then goto 33 end if if (j==iefec2(g,1).and.k==iefec2(g,2)) then goto 33 end if end do write(13,*) i,(j-1)*nmar+k,(finalBD(i,j,k,1)+finalBD(i,j,k,2)-3) 33 end do end do end do close(13) close(13) <u>!write neutral marker genotypes in format</u> open(20,file='nBD.txt') open(13,file='genotBD.txt') do i=1,namin2 marca(j)=0 do j=1,neu read(20,*) ip,im,iv marca(j)=iv+1 end do end do write(13,'(i4,1x,30000i1)') i,(marca(j),j=1,neu) end do close(13) close(20) <u>LPopulation C+D</u> open(12,file='f1CD.txt') open(13,file='f2CD.txt') do i=1,namin2 call unif(x1,u) ip=int(namin2*u)+1 call unif(x1,u) if (u.lt.0.5) then yfen1CD(i)=yfen2C(ip) do j=1,ncro do k=1,nmar finalCD(i) finalCD(i,j,k,1)=finalC(ip,j,k,1) finalCD(i,j,k,2)=finalC(ip,j,k,2) end do end do else stse yfen1CD(i)=yfen1D(ip) yfen2CD(i)=yfen2D(ip) do j=1,ncro do k=1,nmar r finalCD(i,j,k,1)=finalD(ip,j,k,1) finalCD(i,j,k,2)=finalD(ip,j,k,2) finalCD(i end do end if write(12,*) yfen1CD(i),i,1 write(13,*) yfen2CD(i),i,1 end do close(12) close(13) ! frequencies of the causativ <u>! frequencies of the causative mutations of trait 1</u> open(16,file='freq1-CD.txt') do i=1,ngenes p1CD(iefec1(i,1),iefec1(i,2))=0 $\label{eq:uc} \begin{array}{l} u_{0} \; \kappa = 1, nanim2 \\ p | CD(iefec1(i,1), iefec1(i,2)) = p | CD(iefec1(i,1), iefec1(i,2)) + real(finalCD(k, iefec1(i,1), iefec1(i,2), 1) + finalCD(k, iefec1(i,1), iefec1(i,2), 2) - 2) \\ end \ do \end{array}$ do k=1,nanim2 end do p1CD(iefec1(i,1),iefec1(i,2))=p1CD(iefec1(i,1),iefec1(i,2))/(2*real(nanim2)) write(16,*) i,((iefec1(i,1)-1)*nmar+iefec1(i,2)),p1CD(iefec1(i,1),iefec1(i,2)) end do close(16) <u>! additive variance of trait 1</u> Val=0 $\label{eq:var_constraint} \begin{array}{l} \hline Val=0 \\ var(DD(i)=0 \\ doi=1, ngenes \\ var(DD(i)=2^p)(D(iefec1(i,1),iefec1(i,2))^*(1-p1CD(iefec1(i,1),iefec1(i,2)))^*efec1(i)^{**2} \\ Val=Val+var(1CDf(i) \\ end do \\ print *, Variance trait 1 population C+D=',Val \\ wrie(11,*)^V variance trait 1 population C+D=',Val \\ \underline{!frequencies of the causative mutations of trait 2} \\ open(16,file=freq2-CD.txt') \\ doi=1, ngenes \\ p2CD(iefec2(i,1),iefec2(i,2))=0 \\ do k=1, namin2 \end{array}$ do k=1,nanim2 p2CD(iefec2(i,1),iefec2(i,2))=p2CD(iefec2(i,1),iefec2(i,2))+real(finalCD(k,iefec2(i,1),iefec2(i,2),1)+finalCD(k,iefec2(i,1),iefec2(i,2),2)-2) end do end do p2CD(iefec2(i,1),iefec2(i,2))=p2CD(iefec2(i,1),iefec2(i,2))/(2*real(nanim2))) write(16,*) i,((iefec2(i,1)-1)*nmar+iefec2(i,2)),p2CD(iefec2(i,1),iefec2(i,2)))end do close(16) <u>I additive variance of trait 2</u> Va2=0 va2=0 var2CDf(i)=0 do i=1.ngenes var2CDf(i)=2*p2CD(iefec2(i,1),iefec2(i,2))*(1-p2CD(iefec2(i,1),iefec2(i,2)))*efec2(i)**2 Va2=Va2+var2CDf(i) Va2=Va2+Va2CDI(1) end do print *, 'Variance trait 2 population C+D=',Va2 write(11,*) 'Variance trait 2 population C+D=',Va2 *<u>l write genotypes</u>* open(14,file='gCD.txt') do i=1,nanim2 do j=1,ncro do k=1,nmar write(14,*) i,((j-1)*nmar+k),finalCD(i,j,k,1),finalCD(i,j,k,2) end do end do end do close(14) close(14) <u>umarkers and genes frequencies</u> open(13,file='pCD.txt') do i=1,ncro do j=1,nmar pCD=0 do k=1,nanim2 pCD=real pCD=pCD+real(finalCD(k,i,j,1)+finalCD(k,i,j,2)-2) end do write(13,*) (i-1)*nmar+j,pCD/(2*real(nanim2)) end do end do close(13)

<u>! write neutral marker genotypes</u> open(13,file='nCD.txt') do i=1,nanim2 do j=1,ncro do k=1,nmar do g=1,ngenes if (j==iefec1(g,1).and.k==iefec1(g,2)) then goto 34 end if end if if (j==iefec2(g,1),and,k==iefec2(g,2)) then goto 34 end if end do write(13,*) i,(j-1)*nmar+k,(finalCD(i,j,k,1)+finalCD(i,j,k,2)-3) 34 end do end do end do close(13) ! write neautral marker genotypes in format open(20,file='nCD.txt') open(13,file='genotCD.txt') do i=1,nanim2 marca(j)=0 do j=1,neu read(20,*) ip,im,iv marca(j)=iv+1 end do write(13,'(i4,1x,30000i1)') i,(marca(j),j=1,neu) end do close(13) close(20)
$$\begin{split} & \frac{Population A+B+C+D}{Population A+B+C+D} \\ & open(12,file=f1ABCD.txt') \\ & open(13,file=12ABCD.txt') \\ & do i=1, namin2 \\ & call unif(x1,u) \\ & ip=in(tnamin2^{-}u)+1 \\ & call unif(x1,u) \\ & if(u.l.0.25) then \\ & yfen1ABCD(i)=yfen1A(ip) \\ & yfen2ABCD(i)=yfen2A(ip) \\ & do k=1, mar \\ & finalABCD(i,j,k,1)=finalA(ip,j,k,1) \\ & finalABCD(i,j,k,2)=finalA(ip,j,k,2) \\ & finalABCD(i,j,k,2) \\ & finalABCD(i,j,k$$
finalABCD(i,j,k,2)=finalA(ip,j,k,2) end do end do end ao end if if (u.gt.0.25.and.u.lt.0.5) then yfen1ABCD(i)=yfen1B(ip) yfen2ABCD(i)=yfen2B(ip) do j=1,ncro do k=1,nmar finalABCD(finalABCD(i,j,k,1)=finalB(ip,j,k,1) finalABCD(i,j,k,2)=finalB(ip,j,k,2) final ABCD(end do end do end if if (u.gt.0.5.and.u.lt.0.75) then yfen1 ABCD(i)=yfen1 C(ip) yfen2ABCD(i)=yfen2C(ip) do ia_1 wz do j=1,ncro do k=1,nmar r finalABCD(i,j,k,1)=finalC(ip,j,k,1) finalABCD(i,j,k,2)=finalC(ip,j,k,2) end do end do end if end if if (u.gt.0.75) then yfen1ABCD(i)=yfen1D(ip) yfen2ABCD(i)=yfen2D(ip) do j=1,ncro do k=1,nmar final ABCD(i,j,k,1)=finalD(ip,j,k,1) final ABCD(i,j,k,2)=finalD(ip,j,k,2) end do end do end id end if write(12,*) yfen1ABCD(i),i,1 end do close(12) close(13) / frequencies of the causative n <u>Lfrequencies of the causative mutations of trait 1</u> open(16,file='freq1-ABCD.txt') do i=1,ngenes p1ABCD(iefec1(i,1),iefec1(i,2))=0 http://www.international.com/international/internatio write(16,*) i,((iefec1(i,1)-1)*nmar+iefec1(i,2)),p1ABCD(iefec1(i,1),iefec1(i,2))
end do
close(16)
Ladditive variance of trait I
Val=0
var1ABCDf(i)=0
do i=1,ngenes
var1ABCDf(i)=2*p1ABCD(iefec1(i,1),iefec1(i,2))*(1-p1ABCD(iefec1(i,1),iefec1(i,2)))*efec1(i)**2
Val=Val+var1ABCDf(i)
end do
print *, Variance trait 1 population A+B+C+D=',Va1
write(11,*) Variance trait 1 population A+B+C+D=',Va1
Lfrequencies of the curvative mutations of trait 2 <u>I frequencies of the causative mutations of trait 2</u> open(16,file='freq2-ABCD.txt') do i=1,ngenes end do end do p2ABCD(iefec2(i,1),iefec2(i,2))=p2ABCD(iefec2(i,1),iefec2(i,2))/(2*real(nanim2)) write(16,*) i,((iefec2(i,1),i*nmar+iefec2(i,2)),p2ABCD(iefec2(i,1),iefec2(i,2)))write(16,*)1;((ie end do close(16) <u>! additive variance</u> Va2=0 var2ABCDf(i)=0 vat2ABCDf(i)=0 do i=1,gencu vat2ABCDf(i)=2*p2ABCD(iefec2(i,1),iefec2(i,2))*(1-p2ABCD(iefec2(i,1),iefec2(i,2)))*efec2(i)**2 Va2=Va2+wa2ABCDf(i) end do print *, 'Variance trait 2 population A+B+C+D=',Va2

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write(11.*) 'Variance trait 2 population A+B+C+D=',Va2
<u>!write genotypes</u>
open(14,file='gABCD.txt')
do i=1,maim2
             do j=1,ncro
do k=1,nmar
end do
end do
close(14)
<u>/ markers and genes frequencies</u>
open(13,fil=pABCD.txt)
do i=1,nero
do j=1,nero
do k=1,nanim2
pABCD=0
do k=1,nanim2
pABCD=pABCD+real(finalABCD(k,i,j,1)+finalABCD(k,i,j,2)-2)
~~d do
                                                      write(14,*) i,((j-1)*nmar+k),finalABCD(i,j,k,1),finalABCD(i,j,k,2)
      pABCD=pABCD+real(finalABCD(k,i,
end do
write(13,*) (i-1)*nmar+j,pABCD/(2*real(nanim2))
end do
close(13)
<u>1 write neutral marker genotypes</u>
open(13,file="nABCD.txt")
do i=1,nanim2
do i=1,naro
              do j=1,ncro
do k=1,nmar
                                                    ar do g=1,ngenes if (j=i-iefec1(g,2)) then goto 35 end if (=i-iefec2(g,2)) then goto 35 end if if (j=i-iefec2(g,1),and,k==iefec2(g,2)) then goto 35 end if end do write(13,*) i,(j-1)*nmar+k,(finalABCD(i,j,k,1)+finalABCD(i,j,k,2)-3)
      35 end do
end do
end do
close(13)
        <u>l write neutral marker genotypes in format</u>
open(20,file='nABCD.txt')
open(13,file='genotABCD.txt')
do i=1,nanim2
          marca(j)=0
do j=1,neu
read(20,*) ip,im,iv
      read(20,*) ip,im,iv
marca(j)=iv+1
!print *, i,marca(j)
end do
write(13,'(i4,1x,30000i1)') i,(marca(j),j=1,neu)
end do
close(13)
close(20)
close(11)
end
        !---
                subroutine normal(x1,z)
generacion de un numero normal z -> n(0,1)
x1 es la semilla
implicit double precision(a-h,o-z)
real*8 x1,z,u1,u2
         !
                \begin{array}{l} \mbox{call unif}(x1,u1) \\ \mbox{call unif}(x1,u2) \\ \mbox{z=}((-2.*\log(u1))**0.5)*\cos(2.*3.1416*u2) \end{array}
        return
        end
        1 ---
     subroutine unif(x1,u)
generacion de un numero uniforme u[0,1]
st a les a semilla
implicit double precision(a-h,o-z)
divis=2=**63-1.
trans=7**5
divid=trans*x1
lsol=int(divid/divis)
x1=divid=lsol*divis
u=x1/divis
lu=rand()
return
end
function pois(lam,x1)
                 subroutine unif(x1,u)
         function pois(lam,x1)
        implicit none
integer i,pois,n
logical o
     Integer 1,pois.n

logical o

real.parameter:: e=2.71828

real*8.x1sum,Jam

real*8 pro(0:100),prosum(0:100)

pros=0

prosum=0

do n=0,100

pro(n)=1

do i=1,10

do i=1,10

do i=1,10

sum=sum=pro(n)*i

end do

!pro(n)=sum

!print*, prosum(n)
       end do
call unif (x1,u)
!print *,u
        o=.true.
i=0
        do while (o)
       if (u>=prosum(i)) then
i=i+1
else
      else
pois=i
O=.false.
endif
end do
return
end
```