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Additional Information

- Genetic parameters and direct, maternal and heterosis effects on litter size in a
- 2 diallel cross among three commercial varieties of Iberian pig
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ABSTRACT: To be written.... *Keywords:* Prolificacy, Diallel Cross, Heritability, Crossbreeding, Iberian pig

INTRODUCTION

The populations of Iberian pigs are recognized worldwide as some of the porcine populations with the best quality of meat. This is a key factor for their conservation, since they have lower growth and food efficiency (Barea et al., 2011) or prolificacy (Silió et al., 2001) than other commercial populations of pigs. However, the Iberian pork sector is affected by a profound transformation that has involved the disappearance of many traditional producers due to their replacement by intensive management farms. Under this new scenario, the improvement of the reproductive efficiency of the Iberian pig population has become a requirement for its future economic viability.

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The numerical productivity (number of piglets live born per year) can be considered as one of the most important parameters in the profitability of pig farms and its most important component is litter size (Legault, 1985). The genetic improvement of litter size can be achieved with two non-exclusive strategies: 1) within-line selection and 2) crossbreeding heterosis. appropriate between lines to exploit the Traditionally, genetic improvement for litter size in Iberian populations has been scarce (Fernández et al., 2008) and Iberian pig producers have not used crosses since they exclusively breed one of the varieties available in the Iberian breed (Martínez et al., 2000). This contrast markedly with other populations of commercial pigs, where the selection efforts in recent decades for prolificacy in maternal lines have been enormous and where crossbreeding is a standard procedure (Dekkers et al., 2011).

However, several studies have shown the existence of genetic variability for prolificacy within (Rodríguez et al., 1994, Fernández et al., 2008) and between (García-Casco et al., 2013) the Iberian varieties. Therefore, both strategies for improvement are

- 51 plausible, within line selection and implementation of a maternal crossbreeding scheme.
- However, to establish the most appropriate strategy, it is necessary to have reliable
- estimates of the heritability of prolificacy in each population and of the crossbreeding
- parameters defined by the Dickerson model (Dickerson, 1969).
- 55 Thus, the objective of this study was to carry out a full diallel experimental design
- 56 (Hayman, 1954) between 3 contemporary commercial varieties of Iberian pig
- 57 (Entrepelado, Retinto and Torbiscal), under intensive management conditions, as well as
- to estimate the additive genetic variation and the heritability and the crossbreeding effects
- 59 (direct, maternal and heterosis) in prolificacy. The final objective of the results is to
- 60 contribute to the definition of suitable genetic improvement strategies within a pyramidal
- breeding program in the Iberian pig population.

62 MATERIALS AND METHODS

63 Animals and experimental design.

- The research ethics committee of Institute of Agrifood Research and Technology (IRTA)
- 65 previously approved all the management and experimental procedures involving live
- animal and were carried out in accordance with the Spanish Policy of Animal Protection
- 67 RD1201/05, which complies with the European Union Directive 86/609 about the
- 68 protection of animals used in experimentation. The data sets used in this study consist of
- 69 18,193 records for TNB (Total Number Born) and NBA (Number Born Alive) from 3,800
- 70 sows that were obtained from a full diallelic experiment between three strains of Iberian
- 71 pig (Retinto [RR]- Torbiscal [TT] and Entrepelado [EE]) and their reciprocal crosses
- 72 (Retinto × Torbiscal [RT], Torbiscal × Retinto [TR], Retinto × Entrepelado [RE],
- 73 Entrepelado × Retinto [ER], Torbiscal × Entrepelado [TE] and Entrepelado × Torbiscal
- 74 [ET]). The three varieties are recognized in Spain's official Iberian herd-book (Spanish

Association of Iberian Purebred Pig Breeders [AECERIBER]). A more detailed description of its characteristics is presented by Ibáñez-Escriche et al. (2016). The purebred dams were located in two selection farms in intensive commercial management (FARM 1 and 2), while the boars were kept in a center of artificial insemination. The third group of sows is in a commercial production farm (FARM 3) were all sows, purebred and crossed, were mated with boars from a Duroc population according to usual commercial production system in Iberian pig. The distribution of data between farms and breed of sire of service are presented in Table II. In addition, the pedigree of each animal goes back up to 3 generations and it consist of 4,609 individuals. The number of founders per population was 47 (EE), 80 (TT) and 107 (RR).

Statistical analyses.

Both traits, TNB and NBA, were analyzed under a multiple population repeatability model following García-Cortés and Toro (2006), that it is able to divide the additive genetic values into separate fractions depending on their genetic origin. Thus, the general model of analysis was:

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$$y = Xb + \sum_{i=E,R,T}^{3} Z_i a_i + Wp + e$$

where **y** is the vector of phenotypic records (TNB and NBA), **b** is a vector of systematic effects: order of parity (6 levels -1st, 2nd, 3rd, 4rd, 5th and 6th and beyond-), farm-year-season (85 levels), genetic line of boar of mate (4 levels, EE, RR, TT and Duroc –DU-), and the direct line (L_E, L_R, L_T), maternal (M_E, M_R, M_T) and heterosis (H_{ER}, H_{ET}, H_{RT}) effects following Dickerson (1969) model. In addition, **a**_i is the vector of random additive genetic effects of *ith* line (4,609 levels), **p** is the vector of the permanent effect of the sow

(3,800 levels), and \mathbf{e} is the residuals vector; \mathbf{X} , \mathbf{Z}_i , and \mathbf{W} are known incidence matrices that links fixed and random effects with the vector \mathbf{y} . It should be noticed that this model allows for different additive variance component for each genetic origin and the permanent environmental coefficient \mathbf{p} was assumed homogeneous between lines. The model assumed uniform bounded prior for systematic effects and multivariate Gaussian distributions for the additive (\mathbf{a}_i), permanent environmental (\mathbf{p}) and residual effects (\mathbf{e}). The permanent environmental effects and the residuals were assumed identical and independently distributed, thus their prior distribution was:

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$$p \sim N(0, I\sigma_p^2)$$
 $e \sim N(0, I\sigma_e^2)$

On the contrary, the assumed prior distribution for the additive effects were:

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$$a_E \sim N(0, A_E \sigma_{a_E}^2)$$
 $a_R \sim N(0, A_R \sigma_{a_R}^2)$ $a_T \sim N(0, A_T \sigma_{a_T}^2)$

Being A_E , A_R and A_T , the partial kindship matrices generated by the founders of Entrepelado, Retinto and Torbiscal, respectively. The calculation of these matrices was performed following the algorithm proposed by García-Cortés and Toro (2006).

From this general model (LMH), several models model were defined by fixing to zero the direct (MH), maternal (LH), and heterosis effects (LM) and direct and maternal (H), direct and heterosis (M) and maternal and heterosis effects (L). The analyses for all models were carried out using a *Gibbs sampling* algorithm (Gelfand and Smith, 1990) with a single long chain of 1,250,000 iterations, after a period of "burn in" of 250,000 iterations.

Model Comparison

The above described models were compared using the deviance information criteria –

- DIC (Spiegelhalter et al., 2002) and the logarithm of the conditional predictive ordinate
- 121 LogCPO (Gelfand, 1996).
- 122 Deviance information criterion
- The DIC compares the global quality of two or more models accounting for model
- complexity. For a particular model M, the DIC is defined as:

$$DIC_{M} = 2\overline{D}_{M} - D(\overline{\theta}_{M}),$$

- where \overline{D}_M is the posterior expectation of the deviance $D(\theta_M)$, and $D(\overline{\theta}_M) = -2\log(p(y))$
- 127 $|\bar{\theta}_M)$ is the deviance evaluated at the posterior mean estimate of the parameter vector
- 128 (θ_M) . The computation of DIC is composed by two terms, i.e., \overline{D}_M is a measure of model
- fit and $\overline{D}_M D(\overline{\theta}_M)$ is related to the effective number of parameters. Models with smaller
- DIC exhibit a better global fit after accounting for model complexity.
- 131 *Log-marginal probability (logCPO)*
- If we consider the data vector $\mathbf{y} = (y_i, \mathbf{y}_{-i})$, where y_i is the *i*th datum and \mathbf{y}_{-i} is the vector
- of data with ith datum deleted, the conditional predictive distribution has a probability
- density equal to:

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$$p(y_i|\mathbf{y}_{-i}) = \int p(y_i, \theta|\mathbf{y}_{-i}) p(\theta|\mathbf{y}_{-i}) d\theta,$$

- where θ is the vector of unknown parameters in the model. Therefore, $p(y_i|\mathbf{y}_{-i})$ can be
- interpreted as the probability of each datum given the rest of the data, and it is known as
- the conditional predictive ordinate (CPO) for the *i*th datum. The pseudo log-marginal
- probability of the data is then:

 $\sum_{i} \ln p(y_i|\mathbf{y}_{-i}),$

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141 A Monte Carlo approximation of the CPO suggested by Gelfand (1996) is

- 142 $\sum_{i} \ln \hat{p}(y_i | \mathbf{y}_{-i})$, where $\hat{p}(y_i | \mathbf{y}_{-i}) = N \left[\sum_{j=1}^{N} \frac{1}{p(y_i | \theta^j)} \right]^{-1}$, and N is the number of Markov
- chain Monte Carlo (MCMC) draws, and θ^{j} is the *j*th draw from the posterior distribution
- of the corresponding parameter. The higher the value of the LogCPO, best fit of the model
- to data.

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RESULTS

- The results of the model comparison procedures are presented in Table II. For both NBA
- and TNB, the model with best fit with DIC and LogCPO was the complete LMH model,
- followed by a group of models (LH, MH and H) whose differences with the best model
- varied between 5.9 to 8.0 units for DIC and between 0.4 to 6.9 for LogCPO. Finally, the
- LM, D and M models had a worse adjustment and their differences with the LMH model
- ranged between 22.4 to 26.6 for DIC and between 19.8 to 23.3 for LogCPO.
- 153 The posterior mean (and posterior standard deviation) of the variance components,
- heritabilities and the percentages of permanent environmental variation for model LMH
- are presented in Table III for NBA and TNB, respectively. The posterior mean estimates
- 156 (\pm posterior standard deviations) of the additive variance ranged between 0.371 ± 0.106
- 157 (TT) to 0.665 ± 0.123 (EE) and between 0.418 ± 0.115 (TT) to 0.717 ± 0.126 (EE) for
- NBA and TNB, respectively, and the posterior mean estimates (± posterior standard
- deviations) of the permanent environmental and residual variances were 0.361 ± 0.051
- and 4.020 \pm 0.048 for NBA and 0.371 \pm 0.053 and 4.029 \pm 0.048 for TNB. As a
- 161 consequence of the figures described above, the posterior mean (± posterior standard
- deviations) estimates for the population specific heritabilities were 0.078 \pm 0.021 (TT),

- 163 0.084 ± 0.017 (RR) and 0.131 ± 0.022 (EE) for NBA and 0.086 ± 0.022 (TT), $0.090 \pm 0.086 \pm 0.086$
- 164 0.017 (RR) and 0.140 ± 0.022 (EE). The results under models MH, LH, LM, L, M and H
- were similar and they are presented in Supplementary Tables I and II for NBA and TNB,
- 166 respectively.
- The results of contrast between direct line and maternal effects and the heterosis effects
- under model LMH are presented in Table IV. The posterior mean \pm standard deviation
- estimates of the differences between EE and RR, EE and TT and RR and TT were -0.684
- \pm 0.317, 0.114 \pm 0.324 and 0.798 \pm 0.210 for NBA and -0.876 \pm 0.327, -0.024 \pm 0.334
- and 0.852 ± 0.217 piglets for TNB. Moreover, the posterior mean \pm standard deviation
- estimates of the differences of the maternal effects of EE with RR and TT were $0.443 \pm$
- 173 0.123 and 0.450 \pm 0.144 for NBA and 0.554 \pm 0.117 and 0.533 \pm 0.117 piglets for TNB,
- whereas the differences between RR and TT were only 0.007 ± 0.104 and -0.021 ± 0.107 .
- 175 These results of the maternal effects were confirmed with the comparison between
- reciprocals. Thus, the posterior mean \pm standard deviation estimates of differences
- between ER vs RE and between ET and TE were -0.443 ± 0.123 and -0.554 ± 0.125 and
- -0.450 ± 0.144 and -0.534 ± 0.147 for NBA and TNB, respectively, whereas the
- differences between TR and RT were negligible (-0.007 \pm 0.104 and 0.021 \pm 0.107).
- Moreover, the posterior mean (and standard deviation) estimates of the heterosis effects
- ranged between 0.600 ± 0.129 (EE and TT) to 0.690 ± 0.092 (RR and TT) for NBA and
- between 0.622 ± 0.131 (EE and TT) to 0.666 ± 0.093 (RR and TT) for TNB. The results
- under the remaining models followed the same pattern and they are presented in
- Supplementary Tables III and IV for NBA and TNB, respectively.
- Finally, the posterior mean (and standard deviation) of the difference between each cross
- with respect to the RE and the posterior probability of being the best cross with model
- LMH are presented in Table V. As observed, the differences ranged between -0.443 (with

ER) to -1.566 piglets (with TT) piglets for NBA and between -0.464 (with TE) to -1.607 (with TT) for TNB and the posterior probability of being the best cross given the LMH model were 0.985 and 0.991 for NBA and TNB, respectively.

DISCUSSION

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The main advantage of the proposed approach (García-Cortés and Toro, 2006) is that it provides specific estimates of the additive variance component (and heritabilities) for each population of origin and, therefore, it is able to detect differences in their genetic variation. The results of additive variance (and heritability) of the EE variety were markedly higher than the other two populations (RR and TT) under all model of analysis. In fact, in comparison with the estimates available in the literature, the results of the EE population were higher than those obtained in previous studies (Pérez-Enciso and Gianola, 1992; García-Casco et al., 2012) in Iberian pig populations. while the estimates for RR and TT are closer to agreeing with them. However, estimates of heritability over 0.10 are frequent in white pig populations (Noguera et al., 2002; Bidanel, 2011). The results of this study would have to be confirmed with a larger database, but in light of the results, it seems plausible that the response to selection in the EE population under an appropriate breeding scheme for improving litter size can be larger than in the other lines. Therefore, the EE variety is a good candidate to be included in a pyramidal scheme for the improvement of litter size. The results of the comparison of models indicated the relevance of the direct line, maternal and heterosis effects with both procedures (LogCPO and DIC) since the model with the best fit was LMH. However, the magnitude of this relevance was heterogeneous. For example, the best model was followed closely by a group of models that share in common the presence of the effects of heterosis (LH, MH and H). This result is reinforced

by the posterior probability of a heterosis effect greater than zero that was greater than 0.999 for all models and traits. Therefore, we can affirm that there is clear evidence of heterosis between any of the three potential crosses and without notable differences between them. The presence of a relevant degree of heterosis was expected, as is observed regularly in crosses between porcine populations (Haley et al., 1995, Cassady et al., 2002; García-Casco et al., 2012). However, the results were relatively surprising, because previous studies (Fabuel et al., 2004) had suggested that the genetic distance between EE and RR is lower than between TT. Therefore, we expected a lower degree of heterosis between EE and RR than in crosses involving TT. The superiority of model LMH with respect to MH and the posterior distribution of the contrast between direct line effects under the LMH model also confirm the differences between direct line effects in litter size. There are many references in the literature about differences in reproductive performance among swine populations (Bidanel, 2011). Therefore, it is not surprising that this variability also appears among the Iberian varieties since the genetic diversity of Iberian pig populations is as high as among white pig populations (Martinez et al., 2000; Fabuel et al., 2004). Moreover, the Iberian pig populations had evolved thanks to selection and adaptation processes that could be related with litter size. In general, the RR population had the best line effects and the EE and TT populations had a similar performance. In addition, the best fit of LMH with respect to the LH model and the posterior distribution of maternal effects also provides evidence of its relevance. In fact, the results showed a clear superiority (posterior probability over 0.99) of the maternal effect of the EE population with respect to the RR and TT populations, confirmed by the highly relevant differences between reciprocals (RE versus ER and TE versus to ET). In quantitative genetics, the importance of maternal effects is deeply recognized (Wolf and Wade, 2009)

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and they are typically found in growth traits in non-prolific species (Waldron et al., 1993) where they act as an environmental variation in the performance progeny and disappear after weaning. Therefore, the maternal effects are not expected in litter size as it is understood as a trait of the sow. However, some studies have detected the presence of significant maternal genetic variation (Southwood and Kennedy, 1990, Irgang et al., 1994) and, more recently, the influence of maternal care on the reproductive performance of adult females has been investigated in several species (Cameron, 2011). In addition, the recent developments on the genetic determinism of prolificacy in pigs are picturing a very complex panorama that can help to understand the presence of maternal effects in a diallelic cross. In fact, there is recent evidence of the effects of DNA methylation (Hwang et al., 2017) and the presence of imprinted genes (Coster et al., 2012) that affect the size of the litter of pigs. In addition, it has been claimed that epistasis is a very important source of genetic variation in litter size in mice (Peripato et al., 2004) and swine (Noguera et al., 2009). Finally, another possibility that can not be ruled out is the effects of mitochondrial DNA whose variation has recently been associated with the maturation of oocytes in cattle (Srirattana et al., 2017) and the reproductive performance in pigs (Tsai et al., 2016). Despite the uncertainty about the biological basis of the observed results, the conclusion of this study points out the clear advantage of crossbred individuals with respect to the purebred ones in litter size traits in the Iberian breed. Therefore, the implementation of pyramidal scheme to provide crossbred dams to the producers is strongly recommended. Among the analyzed varieties, we have detected that the RR population has a greater direct line effect, whereas the EE population has the greater maternal effect and provides a clear advantage when used as maternal line. As a consequence, in light of the results, the recommended cross for a practical implementation of a pyramidal scheme is RR as

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paternal and EE as maternal lines, reinforced by posterior probability of being the best cross for both NBA and TNB with the analyzed dataset.

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Table I. Distribution of phenotypic records by farm and breed of service sire and dam.

	FARM 1			FARM2			FARM3
Dam (Number of Sows)	Sire		Sire			Sire	
	EE	RR	TT	EE	RR	TT	DU
EE (707)	517	752	5	255	811	70	433
ER (527)	-	-	-	-	-	-	2336
ET (177)	-	-	-	-	-	-	942
RE (196)	-	-	_	-	-	-	806
RR (874)	1450	491	96	655	277	633	870
RT (488)	-	-	-	-	-	-	2450
TE (36)	-	-	-	-	-	-	192
TR (343)	-	-	-	-	-	-	1993
TT (452)	197	808	58	109	507	247	232
TOTAL	2164	2051	159	1019	1595	950	10255

Table II. Results of the Deviance Information Criterion (DIC) and the Logarithm of the conditional predictive ordinate (LogCPO) and the absolute value of the differences with the best model (between brackets) for number born alive (NBA) and total number born (TNB).

	NBA		TNB	
MODEL	DIC	LogCPO	DIC	LogCPO
LMH	112437.0 (0.0)	-56102.7 (0.0)	112417.2 (0.0)	-56062.4 (0.0)
МН	112443.3 (6.3)	-56103.6 (0.9)	112423.1 (5.9)	-56062.8 (0.4)
LH	112443.3 (6.3)	-56107.7 (5.0)	112422.9 (5.7)	-56065.3 (2.9)
LM	112460.8 (23.8)	-56122.5 (19.8)	112439.6 (22.4)	-56084.6 (20.2)
D	112463.3 (26.3)	-56125.0 (22.3)	112443.0 (25.8)	-56085.4 (23.0)
M	112463.6 (26.6)	-56123.6 (20.9)	112443.3 (26.1)	-56085.7 (23.3)
Н	112444.5 (7.5)	-56109.6 (6.9)	112425.2 (8.0)	-56067.2 (4.8)

Table III. Posterior Mean (and Standard Deviation) estimates of heritabilities (h²) and additive, permanent environmental and residual variance components for number born alive (NBA) and total number born (TNB) with model LMH.

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	NBA	TNB
σ_{aE}^2	0.665 (0.123)	0.717 (0.126)
σ_{aR}^2	0.401 (0.088)	0.439 (0.091)
σ_{aT}^2	0.371 (0.106)	0.418 (0.115)
σ_p^2	0.361 (0.051)	0.371 (0.053)
σ_e^2	4.020 (0.048)	4.029 (0.048)
h_E^2	0.131 (0.022)	0.140 (0.022)
h_R^2	0.084 (0.017)	0.090 (0.017)
h_T^2	0.078 (0.021)	0.086 (0.022)

Table IV.Posterior Mean (and Standard Deviation) estimates of the contrast between direct line and maternal effects and the heterosis effects for Number Born Alive (NBA) and Total Number Born (TNB) under model LMH.

		NBA	TNB
Direct Line	EE vs. RR	-0.684 (0.317)	-0.876 (0.327)
	EE vs. TT	0.114 (0.324)	-0.024 (0.334)
	RR vs. TT	0.798 (0.210)	0.852 (0.217)
Maternal	EE vs. RR	0.443 (0.123)	0.554 (0.147)
	EE vs. TT	0.450 (0.144)	0.533 (0.117)
	RR vs. TT	0.007 (0.104)	-0.021 (0.107)
Heterosis	EE with RR	0.653 (0.098)	0.661 (0.099)
	EE with TT	0.600 (0.129)	0.622 (0.131)
	RR with TT	0.690 (0.092)	0.666 (0.093)

Table V. Posterior mean (and standard deviation) of the differences of the crosses withthe RE population and posterior probability of being the best cross under model LMH.

	NBA		TNB	
	Dif.	Prob.	Dif.	Prob.
EE	-0.995 (0.196)	0.000	-1.09 (0.202)	0.000
ER	-0.443 (0.123)	0.000	-0.554 (0.125)	0.000
ET	-0.902 (0.175)	0.000	-1.000 (0.178)	0.000
RE	-	0.985	-	0.991
RR	-0.754 (0.189)	0.000	-0.777 (0.193)	0.000
RT	-0.471 (0.178)	0.003	-0.516 (0.183)	0.002
TE	-0.452 (0.164)	0.003	-0.464 (0.169)	0.003
TR	-0.463 (0.198)	0.009	-0.536 (0.203)	0.004
TT	-1.560 (0.226)	0.000	-1.607 (0.232)	0.000

Dif: Posterior mean estimate (and standard deviation) for the difference with the RE population.. Prob: Posterior probability of being the best cross.

Supplementary Table I. Posterior Mean (and Standard Deviation) estimates of heritabilities (h²) and additive, permanent environmental and residual variance components for Number Born Alive (NBA) with models MH, LH, LM, L, M, H.

	МН	LH	LM	L	M	Н
σ_{aE}^2	0.673 (0.124)	0.693 (0.126)	0.658 (0.123)	0.674 (0.125)	0.660 (0.123)	0.711 (0.127)
σ_{aR}^2	0.466 (0.094)	0.390 (0.087)	0.393 (0.090)	0.387 (0.089)	0.442 (0.094)	0.469 (0.094)
σ_{aT}^2	0.423 (0.110)	0.362 (0.104)	0.382 (0.109)	0.385 (0.109)	0.430 (0.112)	0.454 (0.111)
σ_p^2	0.339 (0.051)	0.367 (0.052)	0.385 (0.053)	0.387 (0.052)	0.368 (0.052)	0.332 (0.051)
σ_e^2	4.019 (0.048)	4.020 (0.047)	4.022 (0.047)	4.022 (0.047)	4.022 (0.048)	4.020 (0.047)
h_E^2	0.133 (0.022)	0.136 (0.022)	0.130 (0.022)	0.132 (0.022)	0.130 (0.022)	0.140 (0.022)
h_R^2	0.096 (0.018)	0.081 (0.017)	0.082 (0.017)	0.080 (0.017)	0.091 (0.018)	0.097 (0.018)
h_T^2	0.088 (0.021)	0.076 (0.021)	0.080 (0.021)	0.080 (0.021)	0.089 (0.021)	0.094 (0.021)

Supplementary Table II. Posterior Mean (and Standard Deviation) estimates of heritabilities (h²) and additive, permanent environmental and residual variance components for Total Number Born (TNB) with models MH, LH, LM, L, M, H.

	MH	LH	LM	L	M	Н
σ_{aE}^2	0.734 (0.128)	0.753 (0.130)	0.709 (0.126)	0.730 (0.128)	0.718 (0.127)	0.771 (0.132)
σ_{aR}^2	0.516 (0.097)	0.430 (0.090)	0.436 (0.093)	0.431 (0.092)	0.498 (0.098)	0.511 (0.097)
σ_{aT}^2	0.483 (0.119)	0.413 (0.115)	0.427 (0.117)	0.434 (0.118)	0.487 (0.121)	0.514 (0.121)
σ_p^2	0.343 (0.052)	0.376 (0.053)	0.395 (0.053)	0.396 (0.053)	0.372 (0.053)	0.339 (0.052)
σ_e^2	4.028 (0.047)	4.029 (0.047)	4.031 (0.047)	4.031 (0.048)	4.031 (0.048)	4.030 (0.048)
h_E^2	0.143 (0.022)	0.146 (0.022)	0.138 (0.022)	0.141 (0.022)	0.139 (0.022)	0.150 (0.022)
h_R^2	0.105 (0.018)	0.089 (0.017)	0.089 (0.018)	0.089 (0.018)	0.101 (0.019)	0.104 (0.018)
h_T^2	0.099 (0.022)	0.085 (0.022)	0.088 (0.022)	0.089 (0.022)	0.099 (0.023)	0.105 (0.022)

Supplementary Table III. Posterior Mean (and Standard Deviation) estimates of the contrast between direct line and maternal effects and the heterosis effects for Number Born Alive (NBA) with models MH, LH, LM, L, M, H.

		МН	LH	LM	L	M	Н
Direct Line	EE vs. RR	-	-0.241 (0.294)	-0.567 (0.311)	-0.250 (0.290)	-	-
	EE vs. TT	-	0.557 (0.295)	0.197 (0.318)	0.464 (0.290)	-	-
	RR vs. TT	-	0.798 (0.170)	0.761 (0.210)	0.715 (0.170)	-	-
Maternal	EE vs. RR	0.333 (0.115)	-	0.326 (0.119)	-	0.235 (0.111)	-
	EE vs. TT	0.564 (0.130)	-	0.294 (0.137)	-	0.409 (0.124)	-
	RR vs. TT	0.231 (0.088)	-	-0.032 (0.098)	-	0.174 (0.082)	-
Heterosis	EE with RR	0.629 (0.097)	0.579 (0.095)	-	-	-	0.598 (0.096)
	EE with TT	0.611 (0.129)	0.493 (0.123)	-	-	-	0.439 (0.123)
	RR with TT	0.693 (0.093)	0.674 (0.091)	-	-	-	0.629 (0.091)

Supplementary Table IV. Posterior Mean (and Standard Deviation) estimates of the contrast between direct line and maternal effects and the heterosis effects for Total Number Born (TNB) with models MH, LH, LM, L, M, H.

		МН	LH	LM	L	M	Н
Direct Line	EE vs. RR	-	-0.318 (0.305)	-0.733 (0.321)	-0.318 (0.300)	-	-
	EE vs. TT	-	0.499 (0.306)	0.089 (0.329)	0.418 (0.301)	-	-
	RR vs. TT	-	0.817 (0.178)	0.822 (0.217)	0.735 (0.176)	-	-
Maternal	EE vs. RR	0.417 (0.117)	-	0.427 (0.121)	-	0.314 (0.114)	-
	EE vs. TT	0.629 (0.133)	-	0.365 (0.140)	-	0.468 (0.127)	-
	RR vs. TT	0.212 (0.091)	-	-0.061 (0.100)	-	0.154 (0.084)	-
Heterosis	EE with RR	0.630 (0.099)	0.568 (0.097)	-	-	-	0.585 (0.097)
	EE with TT	0.628 (0.131)	0.501 (0.125)	-	-	-	0.448 (0.125)
	RR with TT	0.672 (0.094)	0.649 (0.092)	-	-	-	0.607 (0.092)